Barium in seawater 1

- Dissolved distribution, relationship to silicon, and barite saturation state 2
- determined using machine learning 3
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13 Abstract

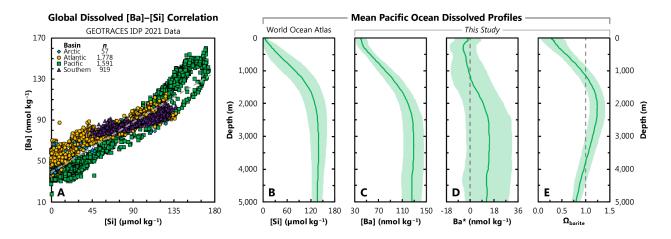
Barium is widely used as a proxy for dissolved silicon and particulate organic carbon fluxes in 14 seawater. However, these proxy applications are limited by insufficient knowledge of the dissolved 15 distribution of Ba ([Ba]). For example, there is significant spatial variability in the barium-silicon 16 17 relationship, and ocean chemistry may influence sedimentary Ba preservation. To help address these issues, we developed 4.095 models for predicting [Ba] using Gaussian Progress Regression 18 19 Machine Learning. These models were trained to predict [Ba] from standard oceanographic 20 observations using GEOTRACES data from the Arctic, Atlantic, Pacific, and Southern Oceans. 21 Trained models were then validated by comparing predictions against withheld [Ba] data from the 22 Indian Ocean. We find that a model trained using depth, temperature, salinity, as well as dissolved 23 dioxygen, phosphate, nitrate, and silicate can accurately predict [Ba] in the Indian Ocean with a 24 mean absolute percentage deviation of 6.0 %. We use this model to simulate [Ba] on a global basis 25 using these same seven predictors in the World Ocean Atlas. The resulting [Ba] distribution constrains the Ba budget of the ocean to $122(\pm 7) \times 10^{12}$ mol and reveals systematic variability in 26 27 the barium-silicon relationship. We also calculate the saturation state of seawater with respect to 28 barite. In addition to revealing systematic spatial and vertical variations, our results show that the 29 ocean below 1,000 m is at equilibrium with respect to barite. We describe a number of possible 30 applications for our model output, ranging from use in biogeochemical models to paleoproxy calibration. Our approach demonstrates the utility of machine learning to accurately simulate the 31 32 distributions of tracers in the sea and provides a framework that could be extended to other trace

33 elements.

34 **1. Introduction**

35 Barium (Ba) is a Group II trace metal that is widely applied in studies of modern and ancient marine biogeochemistry, despite lacking a recognized biochemical function (e.g., Horner & 36 37 Crockford, 2021). These applications of Ba are based on two empirical correlations relating to its 38 dissolved and particulate cycles. The first correlation relates to the dissolved concentration of Ba, 39 hereafter [Ba], which is strongly correlated with that of the algal nutrient silicon (Si; as dissolved 40 silicic acid; Fig. 1; Chan et al., 1977). Unlike [Si], ambient [Ba] concentrations are faithfully 41 recorded by a number of marine carbonates, such as planktonic (e.g., Hönisch et al., 2011) and 42 benthic foraminifera (e.g., Lea & Boyle, 1990), surface- (e.g., Gonneea et al., 2017) and deep-sea 43 corals (e.g., Anagnostou et al., 2011; LaVigne et al., 2011), and mollusks (e.g., Komagoe et al., 44 2018). Preservation of these signals means that the Ba content of carbonates can be related to the 45 Ba content of seawater and, by extension, that of Si. Accordingly, the Ba-Si proxy has been applied 46 to understand ocean nutrient dynamics on decadal (e.g., Lea et al., 1989) to millennial timescales 47 (e.g., Stewart et al., 2021).

48 The nutrient-like distribution of dissolved Ba in seawater is thought to be sustained by the second 49 empirical correlation, relating to cycling of particulate Ba. Particulate Ba in seawater occurs mostly 50 in the form of discrete, micron-sized crystals of the mineral barite (BaSO₄(s), barium sulfate; e.g., 51 Dehairs et al., 1980; Stroobants et al., 1991). Pelagic BaSO₄ is an ubiquitous component of marine 52 particulate matter (e.g., Light & Norris, 2021) and constitutes the principal removal flux of 53 dissolved Ba from seawater (Paytan & Kastner, 1996). Pelagic BaSO₄ is thought to precipitate 54 within ephemeral particle-associated microenvironments that develop during the microbial 55 oxidation of sinking organic matter (e.g., Chow & Goldberg, 1960; Bishop, 1988). The flux of particulate BaSO₄ to the seafloor is correlated with the flux of exported organic matter (e.g., 56 Dymond et al., 1992; Eagle et al., 2003; Serno et al., 2014; Hayes et al., 2021). This correlation 57 58 means that the accumulation rate of sedimentary BaSO₄—or its main constituent, Ba—can be used 59 to trace patterns of past organic matter export on timescales ranging from millenia to millions of 60 years (e.g., Bains et al., 2000; Paytan & Griffith, 2007; Schmitz, 1987; Schroeder et al., 1997).



61 **Figure 1. Distribution of barium in seawater. A.** Property–property plot showing the 4,345 co-located, 62 core-feature complete dissolved data used in ML model training (Sect. 2). Sample locations shown in Figure 63 2. Dashed line shows best-fit linear regression through these data, whereby [Ba] = $0.54 \cdot [Si] + 39.3$. Panels 64 **B., C., D.,** and **E.** show average Pacific Ocean dissolved depth profiles of [Si], [Ba], Ba*, and Ω_{barite}, 65 respectively. Solid line denotes the arithmetic mean and the shaded region encompasses one standard 66 deviation either side of the mean. Dashed line indicates Ba* = 0 (**D**) and Ω_{barite} = 1 (**E**).

While the Ba-based proxies are valuable, their applications are potentially limited by insufficient
knowledge of the distribution of [Ba]. For example, there is significant vertical and spatial
variability in the Ba–Si relationship (Sect. 3.3.; Fig. 1), which we quantify using Ba* (barium-star;
e.g., Horner et al., 2015):

71
$$Ba^* = [Ba]_{in \ situ} - [Ba]_{predicted}$$
[Eq. 1]

72 where [Ba]_{predicted} is based on the Ba–Si linear regression (Fig. 1):

73
$$[Ba]_{predicted} = 0.54 \cdot [Si]_{in \, situ} + 39.3$$
 [Eq. 2]

Here, $[Si]_{in \ situ}$ has units of μ mol kg⁻¹ and $[Ba]_{predicted}$ nmol kg⁻¹; therefore, Ba* also has units of nmol kg⁻¹. The vertical profile of Ba* is rarely conservative (Fig. 1D) and these variations could introduce uncertainty in the reconstruction of [Si] using Ba.

The relationship between sedimentary BaSO₄ accumulation rates and productivity also contains a significant degree of scatter (e.g., Serno et al., 2014; Hayes et al., 2021). Some of this scatter may relate to variability in BaSO₄ preservation, which is at least partially sensitive to ambient saturation

state, Ω_{barite} (e.g., Schenau et al., 2001; Singh et al., 2020; Fig. 1). The saturation state of a parcel of water with respect to BaSO₄ is defined as:

82
$$\Omega_{\text{barite}} = Q / K_{\text{sp}}$$
 [Eq. 3]

where Q is the Ba and sulfate ion product and K_{sp} is the *in situ* BaSO₄ solubility product. Discerning the importance of Ω_{barite} on BaSO₄ preservation has hitherto been challenging owing to the sparsity of *in situ* [Ba] measurements. Accurately determining the global distribution of [Ba] would be valuable for geochemists and oceanographers, and would enable a more thorough investigation of the effects of preservation on BaSO₄ fluxes and refinement of the Ba–Si nutrient proxy.

88 A powerful way of interrogating oceanic element distributions is through modeling. Broadly, there 89 are two modeling approaches relevant for simulating [Ba]: mechanistic (i.e., theory driven) and 90 statistical modeling (i.e., data driven; e.g., Glover et al., 2011). In mechanistic or process-based 91 modeling, model outputs are derived from sets of underlying equations that are based on 92 fundamental theory. As such, mechanistic model outputs can be interrogated to obtain 93 understanding of processes and their sensitivities. However, creating a mechanistic model of the 94 marine Ba cycle requires embedding a biogeochemical model of BaSO₄ cycling within a 95 computationally expensive global circulation model. Although the computational cost associated 96 with building mechanistic models has been reduced by the development of ocean circulation 97 inverse models (e.g., DeVries, 2014; John et al., 2020), this approach still requires detailed 98 parametrizations of the marine Ba cycle, which do not currently exist. In contrast, statistical models 99 are based on extracting patterns from existing data and using those relationships to make 100 predictions. Statistical models encompass a wide variety of approaches ranging from regression 101 analysis to machine learning (ML). Of particular interest to our study are ML models, which can make predictions without any explicit parameterizations of causal relationships. Machine learning 102 103 models are computationally efficient and can be highly accurate, though they offer limited 104 interpretability. Machine learning is increasingly being used to solve problems in Earth and 105 environmental sciences, including simulating the dissolved distribution of tracers in the sea (e.g., 106 for cadmium, Roshan & DeVries, 2021; copper, Roshan et al., 2020; iodine, Sherwen et al. 2019; 107 nitrogen isotopes of nitrate, Rafter et al., 2019; and zinc, Roshan et al., 2018).

108 The goal of this study is to obtain an accurate simulation of [Ba], which ML makes possible even 109 in the absence of a process-level understanding of the marine Ba cycle. We tested thousands of 110 ML models that were trained using quality-controlled GEOTRACES data from the Arctic, 111 Atlantic, Pacific, and Southern Oceans, supplemented by Argo, satellite chlorophyll, and 112 bathymetry data products (Sect. 2.). Models were tested for their accuracy by simulating [Ba] in 113 the Indian Ocean and comparing predictions against observations made between 1977–2013. Since 114 no Indian Ocean data were seen by any of the models during training, we are able to identify 115 models with high generalization performance (Sect. 2.). We then identify an optimal set of 116 predictor variables, calculate model uncertainties, and simulate [Ba], Ba^{*}, and Ω_{barite} on a global 117 basis (Sect. 5.). This result will be valuable for researchers interested in marine Ba cycling, and 118 demonstrates the utility of ML to tackle problems in marine biogeochemistry.

119 **2. Training and testing data**

120 Machine learning algorithms are adept at making accurate predictions of a target variable by 121 identifying relationships between variables within large data sets. However, making accurate 122 predictions first requires that a ML algorithm is trained on existing observations of that variable 123 alongside a number of other parameters. These other parameters, hereafter termed features, are an 124 important part of model training; features should encode information that may help the ML 125 algorithm predict [Ba], otherwise their inclusion may diminish model performance. Features 126 should also be well characterized in the global ocean, which allows ML models to make predictions 127 in regions beyond the initial training dataset. We selected 12 model features by considering the 128 tradeoff between feature availability and presumed predictive power (Table 1). While testing more 129 features may have resulted in a more accurate final model, we found that many observations of 130 [Ba] did not have corresponding data for multiple features; thus, including more features would 131 have meant fewer training data. Moreover, we find that including more than nine features can 132 actually diminish model performance. As such, we did not evaluate the predictive power of other 133 features beyond the 12 initially selected.

134 **Table 1. List of oceanographic parameters selected as model features.** The features tested were

#	Parameter Name	Abbreviation	Units	Coverage*	
1	Latitude	Lat.	degrees north (°N)	_	
2	Longitude	Long.	degrees east (°E)	_	
3	Sample collection depth	Z	meters (m)	_	
4	Temperature	Т	degrees Celsius (°C)	97.44%	
5	Salinity	S	unitless, but often written in 'units' of PSU or PSS	97.44%	
6	Dissolved oxygen	[O ₂]	µmol kg ⁻¹	97.44%	
7	Dissolved nitrate	[NO ₃	µmol kg ⁻¹	97.44%	
8	Dissolved phosphate	[PO4]	µmol kg ⁻¹	97.44%	
9	Dissolved silicon (as silicic acid)	[Si]	µmol kg ⁻¹	97.44%	
10	Maximum monthly mean mixed-layer depth	MLD	meters (m)	88.20%	
11	Mean average annual surface chlorophyll	Chl. a	mg m ⁻³	93.95%	
12	Bathymetry	Bathy.	meters (m)	100%	

135 selected based on their presumed predictive power and geospatial coverage.

*Coverage values represent the percentage of data points within the World Ocean Atlas 2018 grid that have available data for a given parameter. Latitude, longitude, and depth have 100 % coverage as these features define the grid itself.

136 The 12 features used to predict [Ba] and their associated data sources are summarized in Table 1 137 and described below. The first three features (latitude, longitude, depth) record geospatial information that defines the location of an observation in three-dimensional space. To avoid 138 139 numerical discontinuities, latitude and longitude were introduced into the model as a 140 hyperparameter consisting of the cosine and sine of their respective values (in radians). Data for 141 features 1-3 were included in the sample metadata. Features 4-9 encode physical (temperature, 142 salinity) and chemical (oxygen, nutrients) information that is routinely measured alongside [Ba]. 143 These data were generally available for the same bottle as the [Ba] measurements; however, when 144 that was not the case, nutrient data were taken from the corresponding location during a separate

145 cast, or, in the case of oxygen, from linearly interpolated sensor data. The final three features are independent of depth, meaning that all samples within a given vertical profile exhibit the same 146 147 value for MLD (mixed-layer depth), sea-surface chlorophyll a, and bathymetry. Features 10–12 148 were drawn from several data sources. A climatology of MLD (feature 10) was compiled using 149 the Argo database (Holte et al., 2017). We selected maximum monthly mean MLD as the feature 150 of interest, as this appears to be the spatiotemporal scale most relevant for influencing [Ba] 151 distributions (Bates et al., 2017). Feature 11 represents a blended SeaWiFS and MODIS 152 climatology of chlorophyll a that was obtained from the Copernicus Marine Environment 153 Monitoring Service (CMEMS, 2021). We calculated the mean annual chlorophyll *a* for each grid 154 cell in the data product and log transformed the data to reduce parameter weighting (e.g., Rafter et 155 al., 2019). Data for MLD and chlorophyll a were extracted at the location of [Ba] observations 156 using nearest-neighbor interpolation and their values logged in the master record. Bathymetric 157 information (feature 12) was extracted from one of two sources. Our preferred source was the 158 sample metadata, which generally included a value for bathymetry. For samples lacking 159 bathymetric information, we used nearest-neighbor interpolation to extract a value from the 160 ETOPO5 Global Relief Model (National Geophysical Data Center, 1993). Occasionally, the 161 ETOPO5-extracted bathymetry was shallower than the deepest observation of [Ba] in a given 162 vertical profile. In such cases, the bathymetry logged in the master record was set to 1.01 times the 163 depth of the deepest observation in that profile.

164 The [Ba] data from the Indian Ocean were collected from several, primarily pre-GEOTRACES 165 sources (Table 2). As such, these data were generally incomplete for the 12 features used to train 166 the ML models. Rather than using a mixture of *in situ* and interpolated data, we decided to 167 interpolate all Indian Ocean data for parameters 4–12. Data for parameters 4–9 were linearly 168 interpolated from the nearest vertical profile in the World Ocean Atlas 2018 (WOA; Boyer et al., 169 2018; García et al., 2018a; 2018b; Locarnini et al., 2018; Zweng et al., 2018) and values for MLD 170 and chlorophyll *a* were extracted from the aforementioned data products using nearest-neighbor 171 interpolation. Bathymetric information was obtained from either the WOA or ETOPO5. For the 172 vast majority of most samples, bathymetry was taken as the arithmetic mean of the maximum 173 depth of the nearest vertical profile in the WOA and the depth at the standard level below. For 174 example, if the maximum depth at a station was 950 m, the bathymetry was recorded as 975 m, 175 which is the mean of levels 46 (950 m) and 47 (1,000 m). For profiles with a maximum depth of

- 176 5,500 m—level 102, the lowest in the WOA—bathymetry was recorded as either 5,550 m or the
- 177 nearest-neighbor interpolated value from *ETOPO5*, whichever was deeper.

178	Table 2. Data sources. Information regarding the source of [Ba] incorporated into the master record.	
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Purpose	Region	Expedition ID	Data source	Data Originators (if unpublished)			
	South Atlantic (Meridional)	GA02	GEOTRACES IDP 2017 (Schlitzer et al., 2018)	Jose M. Godoy			
	North Atlantic (Zonal)	GA03	Rahman et al., 2022				
	South Atlantic (Zonal)	GA10	Horner et al., 2015; Bates et al., 2017; Hsieh & Henderson, 2017; Bridgestock et al., 2018				
Model training	Southern Ocean (Meridional)	GIPY04	GEOTRACES IDP 2017 (Schlitzer et al., 2018)	Frank Dehairs			
	Southern Ocean (Zonal)	GIPY05	Hoppema et al., 2010				
	Arctic	GIPY11	Roeske et al., 2012				
		GN01	Whitmore et al., 2022				
	Pacific (Meridional)	GP15	GEOTRACES IDP 2021 (GEOTRACES IDP Group, 2021)	Laura Whitmore, Melissa Gilbert, Emilie Le Roy, Tristan Horner, Alan Shiller			
	Subtropical South Pacific (Zonal)	GP16	Rahman	et al., 2022			
		GEOSECS	Craig & Turekian (1980)				
		INDIGO 1					
Model testing	Indian Ocean	INDIGO 2	Jeandel et al. (1996)				
		INDIGO 3					
		SR3	Jacquet et al. (2004)				
		SS259	Singh et al. (2013)				

179

180 This data ingestion process resulted in a master record containing 5,502 observations of [Ba] that 181 also contained a corresponding value for all 12 core features (Table 1). The record was then split 182 into a Pareto partition: the first partition was used for ML model training (4,345 observations, 79 183 % of data; Fig. 1A) and the second for model testing (1,157 data; 21 %). This partitioning was 184 determined based on the basin from which the sample was collected; data from the Arctic, Atlantic, 185 Pacific, and Southern Oceans were used in model training, whereas the 1,157 [Ba] data from the 186 Indian Ocean were reserved for model testing (Table 2; Fig. 2). This location-based separation of 187 training and testing data was chosen to minimize overfitting, which can occur when the training-188 testing separation is randomly assigned (see Sect. 3.2.).

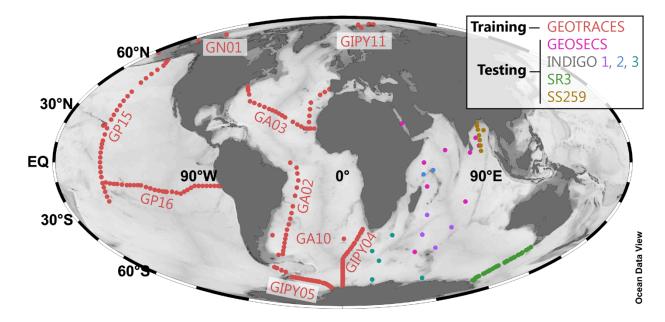


Figure 2. Geographical distribution of the training and testing data. The 4,345 core-feature complete training data (red; Fig. 1) are from the GEOTRACES 2021 Intermediate Data Product (GEOTRACES IDP Group, 2021); GEOTRACES expedition identifiers are noted next to each section. The n = 1,157 testing data from the Indian Ocean are color-coded by expedition. Data sources listed in Table 2.

193 **3. Methods**

194 In the following subsections we discuss details of the specific ML algorithm that was used for

- 195 model development (Sect. 3.1.), explain the model training and testing process (Sect. 3.2.), and
- 196 describe how a global prediction of [Ba] was obtained and interrogated (Sect. 3.3.).

3.1. Algorithm selection and training

198 We opted for supervised ML using a Gaussian Process Regression learner, implemented in 199 MATLAB. This particular ML algorithm is non-parametric, kernel-based, and probabilistic, which 200 means that it does not make strong assumptions about the mapping function, can handle 201 nonlinearities, and takes into account the effect of random occurrences when making predictions. 202 Gaussian Process Regression algorithms are widely used in geostatistics, where it is often referred 203 to as 'kriging' (e.g., Cressie, 1993; Rasmussen & Williams, 2006; Glover et al., 2011). This type 204 of algorithm is ideal when working with continuous data that also contains a certain level of noise, 205 such as from measurement uncertainty or oceanographic variation. The MATLAB function, 206 fitrgp, was used for model training. A full list of the parameter selections used in fitrgp is 207 provided in Table S1. All predictors were normalized and standardized to have a mean of zero and 208 a standard deviation of unity. This process places all parameters on the same relative range and 209 reduces scale dependencies.

210 A selection of the training data were used to train 4,095 different machine learning models with 211 the goal of finding a model that could accurately simulate the global distribution of [Ba]. The 212 number of models derives from the number of features investigated; each model uses a unique 213 combination of the 12 features in Table 1 and our testing followed a factorial design whereby each feature was either enabled or disabled. This design yields a total of 2^{12} unique feature combinations 214 (i.e., levels^{features}); however, since it is not possible to train a model with no features enabled, the 215 216 final number of unique, trainable, ML models with ≥ 1 features is $2^{12}-1=4,095$. The full experiment 217 list is provided in Section 6. Each of the 4,095 models was trained using the same data and with 218 the same function parameters described in Table S1.

219 **3.2.** Assessing model performance

Model performance—accuracy and generalizability—was assessed during two phases: training and testing. During model training, the 4,345 observations of [Ba] from the Arctic, Atlantic, Pacific, and Southern Oceans were randomly split into two folds: a training fold containing 80 % of the observations, and a holdout fold containing the other 20 %. Model accuracy was assessed by comparing model-predicted [Ba] against observed [Ba] for the 20 % of the data in the holdout

225 fold. We then performed additional testing to establish model generalizability. A significant 226 problem in supervised ML, and particularly Gaussian Process Regression learning, is overfitting: 227 models may fit the noise in the training data, leading to poor generalization performance 228 (Rasmussen & Williams, 2006). Since our goal was to develop a global model of [Ba] using 229 regional training data, we deemed it especially important to identify generalizable models. 230 Generalizable models were identified through a testing process involving regional cross-231 validation; each trained model was used to predict [Ba] for the 1,157 samples from the Indian 232 Ocean and model predictions were again compared against observations. Importantly, no [Ba] data 233 from the Indian Ocean were seen by any of the models during training. This process helped to 234 identify models that may have been overfit to the training data and can further be used to calculate 235 generalization errors (Sect. 4.1).

The accuracy of trained models was determined by comparing ML model predictions against withheld data and calculating the mean absolute error (MAE) and mean absolute percentage error (MAPE), defined as:

239
$$MAE = \frac{\sum_{i=1}^{n} |[Ba]_{predicted} - [Ba]_{observed}|}{n}$$
[Eq. 4]

240 and:

241
$$MAPE = \frac{100 \%}{n} \sum_{i=1}^{n} \left| \frac{[Ba]_{predicted} - [Ba]_{observed}}{[Ba]_{observed}} \right|$$
[Eq. 5]

respectively, where *n* is the sample size.

243 Models with lower accuracy exhibit higher errors, whereas models with high accuracy have lower 244 errors. We calculated MAE and MAPE for every possible feature combination, which enables 245 quantification of how specific features affect model performance. Likewise, we calculated errors 246 for each model on predictions made during training (i.e., for the holdout fold) and during model 247 testing (i.e., during regional cross-validation; Fig. 3). This information is used to quantify generalization performance; low errors for both training and testing indicate models that are both 248 249 accurate and generalizable, whereas models with low training errors and high testing errors might 250 indicate models that are overfit to the training data.

3.3. Global predictions

252 A select number of models with low MAE and MAPE were used to simulate [Ba] on a global 253 basis. The process by which we selected these models is described in Section 5.1. Global 254 simulations were performed on the same grid as the WOA, which was also used as the data source 255 for features 1–9 (Boyer et al., 2018). The WOA is a $1^{\circ} \times 1^{\circ}$ resolution data product with around 256 41,000 stations that contain up to 102 depth levels spanning 0-5,500 m in 5, 25, 50, or 100 m 257 increments. Data for features 10–12 (MLD, chlorophyll a, and bathymetry) were also resampled 258 to the WOA grid using the same sources and interpolation methods as described for the Indian 259 Ocean testing data in Section 2. Model outputs were visualized using Ocean Data View software 260 (ODV; Figs. 5-8; Schlitzer, 2023).

261 A selection of the most accurate models of [Ba] were then used to simulate Ba^{*} and $\Omega_{\text{barite.}}$ Star tracers, such as Ba*, are valuable for illustrating processes that influence the cycling of elements 262 263 in the ocean. First defined for N-P decoupling (N*; Gruber & Sarmiento, 1997) star tracers show variations whenever there are differences in the sources and sinks of the two elements being 264 compared. If there are no differences in sources and sinks, the tracer will show conservative 265 266 behavior because both elements share the same circulation. Barium-star is based on Ba-Si decoupling and was first defined by Horner et al. (2015). The definition of Ba* is shown in 267 268 Equations 1 and 2. The coefficients in Equation 2 are based on data from the GEOTRACES 2021 269 Intermediate Data Product and specifically the subset of these data shown in Figure 1. These 270 coefficients differ from previous formulations of Ba* that were based primarily on [Ba] and [Si] 271 data from the Southern and Atlantic Oceans (e.g., Horner et al., 2015; Bates et al., 2017). The 272 global distribution of Ba* was determined by calculating [Ba]_{predicted} (Eq. 2) from [Si]_{in situ} in the 273 WOA 2018 (García et al., 2018b). Values of [Ba]in situ were taken from ML model output and 274 [Ba]_{predicted} were subtracted from this to yield Ba* (Eq. 1).

Values of Ω_{barite} were computed using the method described by Rushdi et al. (2000), summarized in Equation 3. The numerator, Q, represents the *in situ* Ba and sulfate ion product and, in this formulation, depends only on [Ba] and [SO₄^{2–}] molality. The denominator, K_{sp} , depends on T, S, and z (i.e., pressure) and is calculated in two steps: *in situ* T and S are used to calculate the stoichiometric solubility product and then this value is modified by calculating the effect of pressure on partial molal volume and compressibility, which are functions of T and z. As with the 281 calculation of Ba*, values of [Ba]_{in situ} were obtained from ML models and co-located data for T, 282 S, and z were extracted from the WOA (Locarnini et al., 2018; Zweng et al., 2018). Sulfate 283 concentrations were assumed to be conservative with respect to S using $[SO_4^{2-}] = 29.26$ mmol kg⁻ 284 ¹ when salinity = 35 PSU. This latter assumption likely breaks down in certain environments (e.g., 285 where $[SO_4^{2-}]$ reduction occurs); as such, our model is not used to predict Ω_{barite} in restricted basins, 286 such as the Black Sea or Caspian Sea. Given that our estimates of Ω_{barite} exhibit a MAE of 0.08 287 (Appendix), we believe that values of Ω_{barite} between 0.92 and 1.08 are indicative of equilibrium 288 between BaSO₄ and seawater.

289 Output from the most accurate ML models was then used to calculate mean [Ba] and Ω_{barite} for 290 each basin, for a series of prescribed depth bins, and for the global ocean. This calculation was 291 performed by weighting each cell in the model output by its volume, which ensures a fair 292 comparison between any two points in the model output. We then subdivided the global ocean into 293 five sub-basins: Arctic, Atlantic, Indian, Pacific, and Southern. Basin boundaries were defined as 294 per Eakins & Sharman (2010), though we merged the Mediterranean and Baltic Seas into the 295 Atlantic and considered the South China Sea as part of the Pacific Ocean. Neither [Ba] nor Ω_{barite} 296 were simulated in the Black or Caspian Seas and thus these regions are not included in the global 297 mean calculations.

4. Results

4.1. Factors affecting model accuracy

300 Here we examine how model performance is influenced by the number and nature of features 301 included during training. We consider model performance in terms of accuracy and 302 generalizability, which we quantify using MAE (Eq. 4). We first explore how the number of 303 features influences model performance (Fig. 3). Here we see that increasing the number of features 304 generally improves the accuracy of trained models; however, the response differs depending on 305 whether accuracy is calculated based on comparison to the holdout fold (i.e., during model 306 training) or to the withheld Indian Ocean data (i.e., during model testing). When considering only 307 the holdout fold, trained models predict [Ba] with a high level of accuracy—the mean, median, 308 and most-accurate trained models achieve a MAE of 2.4, 1.7, and 1.3 nmol kg⁻¹, respectively. 309 Similarly, increasing the number of features almost always improves model accuracy; the MAE of

- 310 the most accurate model for a given number of features decreases from 6.5 to 1.3 nmol kg^{-1} as the
- 311 number of features is increased from one to nine, at which point MAE plateaus between 1.4–1.5
- 312 nmol kg⁻¹ for models with 10–12 features (Fig. 3A).

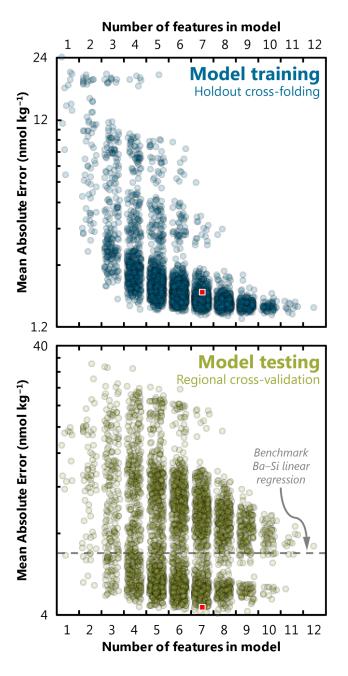


Figure 3. Effect of feature addition on ML model accuracy. Accuracy was quantified for each of the 4,095 trained models and quantified here using MAE (note log scale, which differs between panels). The accuracy of trained models is shown for random holdout cross-validation during training (top) and for regional cross-validation during testing (bottom). Square indicates the performance of our favored predictor model, #3080 (see Fig. 4, Sect. 5.1). The accuracy of the Ba–Si linear regression benchmark is shown as a dashed line in the lower panel (MAE = 6.8 nmol kg⁻¹). To illustrate data density, points have been randomly positioned within their respective bin and plotted with 80 % transparency.

320 Moving to the regional cross-validation, the overall performance of models is lower; the same 321 4,095 trained models achieve a mean, median, and most-accurate MAE for the Indian Ocean 322 dataset of 8.8, 7.9, and 4.0 nmol kg⁻¹, respectively. For comparison, if [Ba] was estimated for these same 1,157 Indian Ocean samples using the linear [Ba]–[Si] relationship (Fig. 1) and ambient [Si] 323 324 as the only predictor, this linear model would achieve a MAE of 6.8 nmol kg⁻¹. Thus, there are 1,687 ML models that achieve a superior accuracy to existing methods for estimating [Ba], 325 326 offering an improvement of as much as 41 % (Fig. 4). However, regional cross-validation also 327 shows that the addition of more features may, in fact, degrade model performance. The MAE of 328 the most accurate model for a given number of features decreases from 6.6 to 4.0 nmol kg⁻¹ as the number of features is increased from one to eight. As the number of features is increased from 9-329 330 12, the MAE of the most-accurate models increases monotonically from 4.1 to 7.1 nmol kg^{-1} . The overall lower performance of trained models during regional cross validation-and the observation 331 332 that many of the feature-rich models perform worse than models with fewer features—is indicative 333 of certain models being over-fit to the training data. Together, these observations suggest that the 334 optimum number of features needed to accurately predict [Ba] is between six and nine.

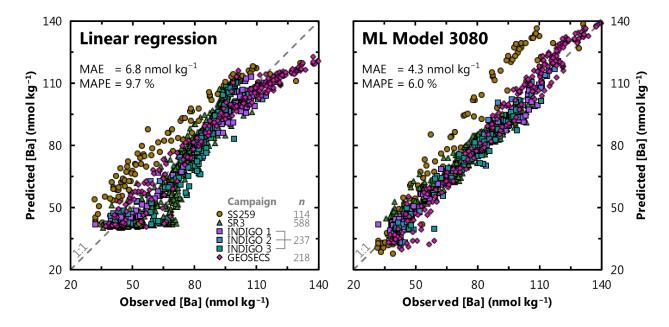


Figure 4. Comparison of existing and ML methods to estimate [Ba] in seawater. Left panel shows the performance benchmark: predicted [Ba] for the Indian Ocean testing data using the [Ba]–[Si] linear regression and ambient [Si] as the sole predictor. Right panel shows predicted [Ba] using ML model 3080, which improves on existing methods by more than 37 %. Perfect correspondence between predictions and observations is indicated b the dashed line marked '1:1.'Data locations and sources are shown in Fig. 2 and Table 2, respectively; *n* refers to the number of testing data for each campaign. Mean Absolute Error (MAE; Eq. 4) and Mean Absolute Percentage Error (MAPE; Eq. 5) are noted for both models.

342 We also evaluated the nature of the predictors used to estimate [Ba]. The full factorial experiment 343 design enables us to perform comparisons between all models that contained a certain feature and all of those that did not (Sect. 3.1). We quantified the effect of adding a feature by comparing the 344 345 absolute and percentage change in MAE relative to the mean MAE of the two sets of models. This 346 comparison was performed three times: for all 4,095 models based on the holdout cross-folded 347 training data, for all models using the regionally cross-validated testing data, and again for the 348 testing data, but only considering those 1,687 models that achieved a superior accuracy compared 349 to the [Ba]–[Si] linear regression model (Table 3).

Table 3. Feature addition analysis. Effect of each feature on model performance for Training and Testing datasets. Model performance is quantified using MAE, thus all columns have units of nmol kg⁻¹ unless otherwise shown. The Testing analysis is further subdivided into a comparison of all models and 'good' models, meaning those that achieved superior accuracy than the Ba–Si linear regression (Fig. 1).

	Training All models (<i>n</i> = 4,095)			Testing						
				All models (<i>n</i> = 4,095)			Good models (<i>n</i> = 1,687)			
Feature	Mean MAE of models with feature	Mean MAE of models without feature	Relative change in MAE	Mean MAE of models with feature	Mean MAE of models without feature	Relative change in MAE	Mean MAE of models with feature	Mean MAE of models without feature	Relative change in MAE	Share of models with feature
[Si]	1.71	3.03	-56%	7.08	10.6	-39%	5.06	5.50	-8.3%	63%
z	1.83	2.90	-45%	7.94	9.70	-20%	5.05	5.44	-7.4%	55%
[O ₂]	2.03	2.71	-29%	8.25	9.39	-13%	5.14	5.33	-3.8%	54%
Τ	1.78	2.96	-50%	7.61	10.0	-27%	5.17	5.31	-2.8%	59%
[NO ₃]	2.09	2.65	-24%	8.27	9.36	-12%	5.16	5.30	-2.7%	53%
[PO ₄]	2.11	2.63	-22%	8.24	9.40	-13%	5.17	5.30	-2.4%	53%
S	2.02	2.72	-29%	8.67	8.97	-3.5%	5.23	5.23	0.0%	53%
Bathy.	2.30	2.44	-6.1%	8.55	9.08	-6.0%	5.23	5.22	0.2%	51%
Chl.	2.25	2.48	-10%	8.67	8.97	-3.5%	5.24	5.22	0.4%	50%
MLD	2.31	2.43	-4.8%	8.69	8.95	-3.0%	5.24	5.21	0.5%	50%
Lat.	2.16	2.58	-18%	8.13	9.51	-16%	5.32	5.11	4.0%	54%
Long.	2.17	2.57	-17%	11.4	6.24	58%	6.45	5.19	22%	3%

This analysis yields three main results. When considering only the holdout cross-folded training data, the addition of any of the 12 features improves model performance by between -4.8 and -56%. Excepting longitude, similar across-the-board improvements were observed when considering only the testing data, though the improvements for most features were more modest (between -3.0and -39 %). If considering only the 'good' models, six features improved model performance by -2.4 and -8.3 % ([PO₄], [NO₃], *T*, [O₂], *z*, and [Si]), five degraded model performance by +0.2 to +22 % (bathy., Chl. a, MLD, lat., and long.), and salinity had no significant effect (Table 3). Overall, our results indicate that between six and nine features will result in an accurate and generalizable ML model of [Ba], and that $[PO_4]$, $[NO_3]$, *T*, $[O_2]$, *z*, [Si], and possibly *S*, are likely to be included as predictors in such a model.

364 **4.2. Model outputs**

Almost 1,700 models achieved superior accuracy compared to the Ba–Si linear regression benchmark of 6.8 nmol kg⁻¹. We winnow this list to a single model, #3080, in the next section. We henceforth refer to model #3080 as our favored predictor model, which achieves a MAE of 4.3 nmol kg⁻¹ using *z*, *T*, *S*, [O₂], [PO₄], [NO₃], and [Si] as predictors (Fig. 4). Model #3080 is used to simulate [Ba], Ba*, and Ω_{barite} on a global basis and to calculate whole-ocean averages. Surface plots showing the model outputs for the sea surface, 1,000 m, 2,000 m, and 4,000 m are shown in Figures 5, 6, 7, and 8, respectively.

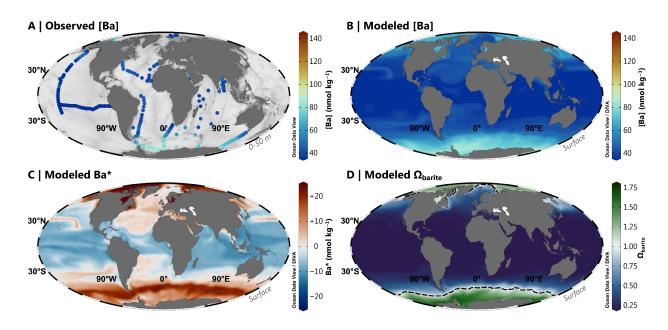


Figure 5. Barium at the sea surface. Observed [Ba] between 0–50 m (A); Model 3080 [Ba] (B), Ba* (C), and Ω_{barite} (D). The dashed line in Panel D indicates the BaSO₄ saturation horizon (i.e., Ω_{barite} = 1.0). Panels A and B use the *roma* color map, whereas Panels C and D use *vik* and *cork*, respectively (Crameri, 2018). Color palettes and parameter ranges are the same for the respective panels in Figure 6–8.

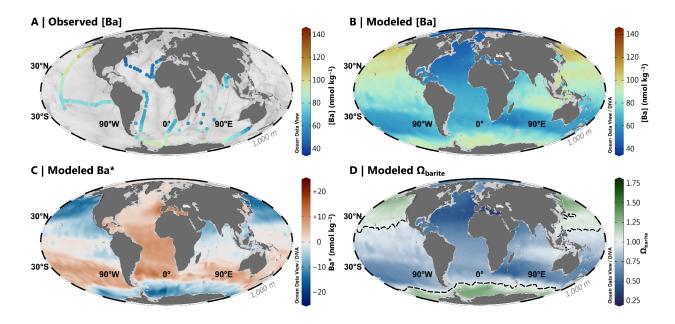
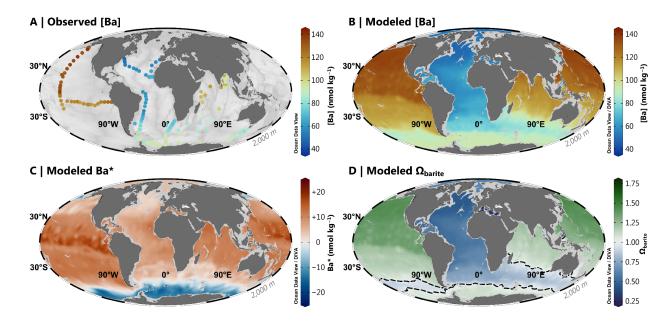


Figure 6. Barium at 1,000 m. Observed [Ba] (**A**); Model 3080 [Ba] (**B**), Ba* (**C**), and Ω_{barite} (**D**). The dashed

 $377 \qquad \mbox{line in Panel D indicates the BaSO_4 saturation horizon.}$



378

Figure 7. Barium at 2,000 m. Observed [Ba] (**A**); Model 3080 [Ba] (**B**), Ba* (**C**), and Ω_{barite} (**D**). The dashed line in Panel D indicates the BaSO₄ saturation horizon.

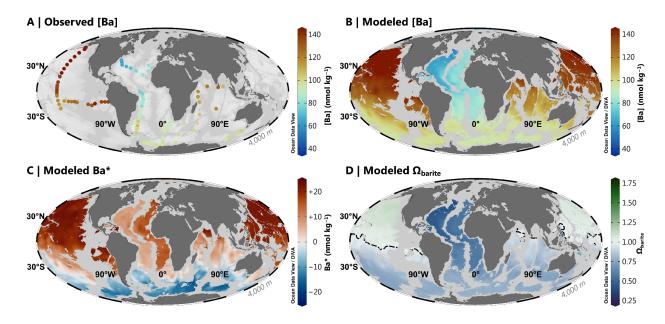
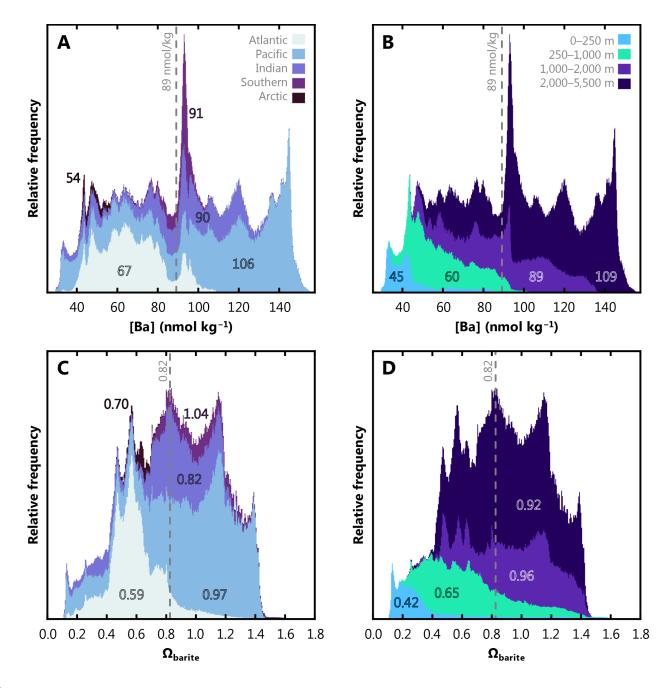


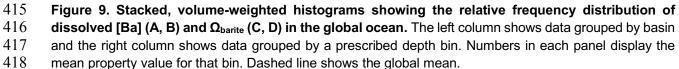
Figure 7. Barium at 4,000 m. Observed [Ba] (A); Model 3080 [Ba] (B), Ba^{*} (C), and Ω_{barite} (D). The dashed line in Panel D indicates the BaSO₄ saturation horizon.

Model #3080 contains 3,302,570 predictions for each of [Ba], Ba*, and Ω_{barite} (Sect. 6). Assuming that the MAPE and MAE are good estimates of the prediction error, we estimate that modeled [Ba] and Ba* have uncertainties of 6.0 % and 4.3 nmol kg⁻¹, respectively. Uncertainties on Ω_{barite} were estimated by comparison to literature data, which yields a MAE of 0.08. These estimates are discussed in more detail in Section 5.2 and the Appendix.

Modeled [Ba] ranges from 26.2–156.8 nmol kg⁻¹ and the data exhibit an unweighted mean of 72.0 388 389 nmol kg⁻¹. The range of model #3080 predictions is within the range of [Ba] encountered in the 390 4,345 training data (17.1–159.8 nmol kg⁻¹). This is an important consideration when assessing the 391 accuracy of Gaussian Process Regression models, and we provide additional discussion of this point in the Supplement. Based on our formulation (Eqs. 1, 2), Ba* varies from -27.2 to +27.9 392 nmol kg⁻¹ and possesses an unweighted mean of +2.4 nmol kg⁻¹. Values of Ω_{barite} vary from 0.11 393 394 to 1.70 and exhibit an unweighted mean of 0.75. To account for the different volumes represented 395 by each cell in the WOA grid, we constructed a volume-weighted mean of [Ba] and Ω_{barite} for the 396 ocean as a whole, for each ocean basin, and for a series of prescribed depth bins (Fig. 9). Looking 397 at the ocean as a whole, the probability density function of [Ba] roughly resembles a uniform distribution, with a mean ocean [Ba] of 89 nmol kg⁻¹ (Fig. 9A). Within this mean is considerable 398

- spatial and vertical variation. For example, the Arctic Ocean exhibits the lowest volume-weighted mean [Ba] of 54 nmol kg⁻¹, whereas mean Pacific [Ba] = 106 nmol kg⁻¹. The Indian Ocean exhibits a similar mean [Ba] (90 nmol kg⁻¹) to the mean of the global ocean. Shallower than 1,000 m, [Ba] infrequently exceeds 100 nmol kg⁻¹, whereas concentrations <45 nmol kg⁻¹ are rare below 1,000 m (Fig. 9B).
- 404 The probability density function of volume-weighted Ω_{barite} is more similar to a normal 405 distribution, albeit with a slight negative skew. Volume-weighted mean oceanic Ω_{barite} is 0.82. The 406 Arctic, Atlantic, and Indian Oceans are, on average, undersaturated with respect to BaSO₄, all exhibiting mean $\Omega_{\text{barite}} \leq 0.82$. In contrast, the Pacific and Southern Oceans are within uncertainty 407 408 of saturation, with mean Ω_{barite} of 0.97 and 1.04, respectively (Fig. 9C). Values of $\Omega_{\text{barite}} < 0.2$ are 409 mostly restricted to the upper 250 m, whilst values of Ω_{barite} exceeding 1.5 are exceptionally rare, 410 found only in the upper 1,000 m of the Southern Ocean. Lastly, Ω_{barite} tends to increase between 411 the 0–250 m, 250–1,000 m, and 1,000–2,000 m depth bins, increasing from 0.42, to 0.65, and 0.96, 412 respectively. Average Ω_{barite} in the deepest bin (2,000–5,500 m) is slightly lower, with a mean value of 0.92 (Fig. 9D). Given the accuracy of our model-derived Ω_{barite} predictions (0.08 to 0.10), 413 414 the ocean between 1,000–5,500 m is within uncertainty of BaSO₄ equilibrium.





419 **5. Discussion**

420 **5.1.** Identification of the optimal predictor model

421 Choosing a single, optimal model configuration is challenging given the sheer number of skillful 422 ML models. Below we winnow the list from 4,095 to a single model (#3080). We base our 423 winnowing primarily on the results of the regional cross-validation performed in the Indian Ocean, 424 rather than from the errors determined from random holdout cross folding of the training data. We 425 believe that there are three strong reasons for winnowing in this way. First, Gaussian Process 426 Regression Learners tend to fit the noise in the training data, meaning that the training error is 427 significantly lower than the generalization error (Rasmussen & Williams, 2006). Indeed, trained 428 models showed overall lower performance during testing compared to training, which we believe 429 is evidence of overfitting (Fig. 3, Table 3). Second, a generalizable global model should be able to 430 make predictions in regions where it has not already learned anything about the target variable. 431 Our regional cross-validation approach satisfies this consideration since no Indian Ocean data were 432 seen during model training. Third, the Indian Ocean is an ideal basin for testing as it exhibits the 433 full diversity of features expected to influence [Ba] (riverine inputs, oxygen-minimum zones, 434 coastal upwelling, etc.) and constitutes ≈ 20 % of the global ocean volume. Likewise, the Indian 435 Ocean captures most of the range in [Ba] seen elsewhere in the ocean (Fig. 9); this likely reflects 436 the input of Atlantic waters through the Aughulas leakage, transport of old Pacific waters via the 437 Indonesian Throughflow, and northward spreading of mode and intermediate waters from the 438 Southern Ocean. We thus assume that the Indian Ocean testing errors are a good approximation of 439 the generalization error, which we now use to winnow the list of models.

440 Our results show that 1,687 of the 4,095 ML models (41 %) produce more accurate predictions of 441 [Ba] than the benchmark Ba–Si linear regression using [Si] as the sole predictor (Fig. 3, Table 3). 442 We focus our winnowing on these 1,687 models as they are superior to existing methods for 443 estimating [Ba] in seawater. Focusing only on these 'good' models reveals significant differences 444 in the information content of the 12 features tested. For example, the inclusion of spatial 445 information in the form of latitude and longitude significantly degrades mean model performance 446 by between +4.0 and +22 %, respectively. While bathymetry, chlorophyll a, and mixed-layer depth 447 exhibited only minor influences, they were nonetheless deleterious to mean model performance 448 by between +0.2 to +0.5 % (Table 3). Only [PO₄], [NO₃], T, [O₂], z, and [Si] consistently improved

449 the mean ML model, which corresponds to model #3112 (testing MAE of 4.3 nmol kg⁻¹). 450 However, visual inspection of model #3112 output reveals that it does not reproduce expected 451 near-shore surface plumes of elevated [Ba] close to certain major rivers (see Supplement). Though 452 volumetrically minor, riverine inputs are a geochemically important component of the marine Ba 453 cycle, and the existence of nearshore Ba plumes underpins a major proxy application of Ba. Near-454 shore riverine influence is easily discerned by low S; we thus explored output from model #3080, 455 which is identical to model #3112, but includes S as a seventh feature during training. Models 456 #3080 and #3112 exhibit identical statistical performance for the testing data (MAE = 4.3 nmol 457 kg⁻¹; Fig. S1) and make similar predictions for mean marine [Ba] and Ω_{barite} (89 nmol kg⁻¹ and 458 0.82, respectively; see Supplement). The similar performance of the two models is consistent with 459 S exerting a near-negligible impact on overall model performance (Table 3). Despite this small 460 effect, model #3080 is better able to reproduce riverine [Ba] plumes compared to model #3112 461 (see Supplement). We therefore consider model #3080 to be our best estimate of marine [Ba]. 462 Model #3080 achieves a MAPE of 6.0 %, which represents a 39 % improvement over existing 463 methods to estimate [Ba] (Fig. 4). We henceforth consider model #3080 as our optimal predictor 464 model, which we use to simulate [Ba], Ba*, and Ω_{barite} in Figures 5–9.

465 **5.2. Model validation**

We now explore the validity of model #3080 in terms of its oceanographic consistency, the sources of uncertainty that affect its accuracy, and potential limitations of the model output. We find that model #3080 reproduces the major known features of the marine [Ba] distribution and makes testable predictions for regions that are yet to be sampled.

470 5.2.1. Visual inspection of model output

Visual inspection of model output is an important component of data analysis considering the
limits of statistical tests (see e.g., Anscombe, 1973). Models may produce statistically satisfactory
fits to the testing data, but the oceanic realism of the output is also important to consider. Modeled
[Ba] should display patterns consistent with related oceanographic properties and exhibit smooth

vertical and spatial variations (Boyle & Edmond, 1975). Predicted [Ba] from model #3080 does
indeed show smooth and systematic spatial and vertical variations that also resembles sparse
observations (Figs. 4–8).

478 Model #3080 also shows systematic increases in [Ba] close to land, especially near the mouths of 479 major rivers (Fig. 4). This is reassuring given that elevated sea-surface [Ba] close to rivers is both 480 widely reported and one of the major proxy applications of Ba: reconstructing spatiotemporal 481 patterns of terrestrial runoff by measuring the Ba:Ca ratio of carbonates (e.g., Sinclair & 482 McCulloch, 2004; LaVigne et al., 2016). For example, model #3080 correctly identifies elevated 483 [Ba] near the Ganges-Brahmaputra (Singh et al., 2013), Río de la Plata (GEOTRACES IDP 484 Group, 2021), and Yangtze outflows (Cao et al., 2021). Model #3080 also predicts elevated sea-485 surface [Ba] in the Gulf of Guinea where several rivers discharge, including the Niger River; the 486 Eastern Tropical Atlantic associated with the Congo River (Edmond et al., 1978; Zhang et al., 487 2023); and in the Gulf of St. Lawrence (St. Lawrence River; see Supplement for additional details 488 and figures). Except for the Congo River, these predictions of elevated near-shore [Ba] await 489 corroboration. Interestingly, model #3080 does not predict elevated [Ba] at all major river mouths; 490 neither the Mississippi nor Amazon Rivers are associated with significant increases in sea-surface 491 [Ba] (see Supplement). The reasons for the lack of elevated [Ba] near the outflow of these two 492 rivers is less clear. It is possible that the model is simply inaccurate in these regions, though we 493 have no particular reason to believe that this is the case. Alternatively, it may reflect seasonal 494 variations in Ba release that are not captured by our mean annual model (e.g., Joung & Shiller, 495 2014). It could also indicate that these particular rivers are not major net sources of Ba to the 496 surface ocean, which might be the case if dissolved Ba is being retained in the catchment (e.g., 497 Charbonnier et al., 2020) or estuary (e.g., Coffey et al., 1997).

498 Overall, model #3080 makes accurate, oceanographically consistent predictions of [Ba] in the 499 Indian Ocean using input data from the WOA. Model #3080 also makes a number of testable 500 predictions of [Ba] in regions lacking direct observations. Given that these predictions were made 501 using the same model and the same WOA inputs, we believe that it is reasonable to assume that 502 model #3080 output is an accurate representation of mean annual global [Ba].

503 5.2.2. Quantifying uncertainties

504 We now describe and, where possible, quantify two possible sources of uncertainty to our ML 505 model output. Before doing so, we describe how uncertainty is quantified as well as the uncertainty 506 of existing approaches. Certain ML models, such as Gaussian Process Regression, offer low 507 interpretability, meaning it is not possible to assess uncertainty using a conventional error 508 propagation. Thus, all model uncertainties are assessed *post hoc*, by comparing predictions against 509 observations. Our preferred metrics are MAE and MAPE (Eqs. 4, 5). Existing approaches for 510 estimating [Ba] result in a wide range of uncertainties. At the low end, the uncertainty associated 511 with measuring [Ba] in seawater represents a fundamental limit to the accuracy of any model. A 512 number of analysts report measurement uncertainties in the range of 1-2 % (e.g., Pyle et al., 2018; 513 Cao et al., 2020). This level of intra-laboratory uncertainty is typical for [Ba] data obtained using 514 isotope dilution-inductively coupled plasma mass spectrometry, and applies to GEOTRACES-era 515 datasets and to much of the training data from the Indian Ocean. However, intra-laboratory 516 uncertainty is typically much smaller than inter-laboratory uncertainty, which is often between 6– 517 9% (e.g., Hathorne et al., 2013). At the upper end, the benchmark Ba–Si linear regression achieves 518 a MAPE of 9.7 % in the Indian Ocean (Fig. 4). Thus, useful ML models of [Ba] should achieve 519 MAPE between 1–10 %. Indeed, our favored predictor model, #3080, achieves a MAPE of 6.0 %.

Now we consider two factors that contribute to the observed 6.0 % uncertainty: realization 520 521 uncertainty and uncertainties in the training data. The realization uncertainty stems from the fact 522 that two models trained on the same training dataset—even with the exact same subset of model 523 features—will produce slightly different predictions. This is due to the holdout cross-folding 524 process used during model training, which partitions the training dataset into random subsets (525 Sect. 3.1.). Thus, the training process results in a slightly different trained model each time the 526 model is realized. We quantified the realization uncertainty by training select models 100 times 527 and calculating the relative standard deviation of the different predictions of [Ba] for the 3.3 528 million values in the output. This uncertainty is small; the median, mean, and maximum realization 529 uncertainty was 0.03 %, 0.04 %, and 0.32 % variability in modeled [Ba].

Next we consider uncertainties in the training data. As noted above, many labs report uncertainties on [Ba] measurements of 1-2 %, while inter-laboratory differences may be up to a factor of five larger. However, this does not consider any uncertainties associated with the other physical and 533 chemical features used to predict [Ba]. In general, these supporting measurement uncertainties 534 should be small: all overboard sensors are regularly calibrated and biogeochemical properties in 535 GEOTRACES are determined using established methods that are based on GO-SHIP best practices 536 (Hood et al., 2010). Moreover, all GEOTRACES sections include crossover stations that are 537 intended to facilitate intercalibration of all parameters, including those used here to predict [Ba] 538 (Fig. 2; Cutter, 2013). The WOA, MLD, Chl. a, and bathymetry data products are similarly 539 subjected to stringent quality review and so we consider it unlikely that these data contribute 540 systematic biases. We believe that the most likely source of uncertainty relates to the fact that all 541 predictor information used for model testing in the Indian Ocean was derived from time-averaged 542 data products, whereas [Ba] was derived from in situ measurements. We made this decision 543 because the *in situ* data were incomplete for all 12 core features (Table 1), and this would have 544 necessitated interpolation for some features and not others. Since all models were tested using the 545 same predictor information, the comparison process should avoid systematic errors, though this 546 does not preclude temporal variability, described next.

547 5.2.3. Other considerations

548 We now consider four other factors that potentially contribute to the uncertainty of the model 549 output: short- and long-term temporal variations, limitations of ML, and uncertainties regarding 550 the thermodynamic properties of BaSO₄. Short-timescale variability in [Ba] may affect how 551 models were evaluated, though this effect is difficult to quantify. In principle, the trained models 552 should be able to resolve seasonal variations in [Ba] since they were trained on *in situ* physical and 553 chemical data. In contrast, model predictions in the Indian Ocean were made using annual average 554 physical and chemical conditions and then evaluated by comparing these predictions against *in* 555 situ [Ba]. The temporal mismatch between Indian Ocean observations and predictions is unlikely 556 to be significant in the deep ocean, where seasonal variations are minor and the Ba residence time 557 is longest (e.g., Hayes et al., 2018). Seasonal variations are, however, likely to matter more for the 558 surface ocean. We were able to minimize some of the impact of these uncertainties by using long-559 term averages of Chl. a and the maximum monthly mean MLD during model training and testing. 560 Significant seasonal mismatches for other parameters are unavoidable given that [Ba] data are too 561 sparse to develop a time-resolved model. We suspect that these variations are most likely to be 562 significant for boundary sources rather than biogeochemical cycling of Ba; significant 563 biogeochemical drawdown of surface [Ba] over seasonal timescales appears to be rare (e.g., Esser 564 & Volpe, 2002), whereas there are large seasonal variations in river discharge that impact near-565 shore [Ba] (e.g., Samanta & Dalai, 2016). These suspicions could be tested using a model with 566 better than 1×1° spatial resolution, which—in theory—is possible with model #3080, so long as 567 similarly high-resolution data are provided for the six predictors utilized by this model (z, T, S, 568 [O₂], [PO₄], [NO₃], and [Si]). While it is challenging to precisely quantify seasonal uncertainties, 569 we note that model #3080 performs well at low [Ba], which is found mostly near the surface, where 570 seasonal variations should exhibit the largest effects. Likewise, seasonal variations will have only 571 a minor effect on our calculations of global mean [Ba] or Ω_{barite} (Fig. 8).

572 Long-term variability in [Ba] may also influence model performance, since the testing data from 573 the Indian Ocean were collected between 1977 (GEOSECS) and 2008 (SS259; Fig. 2). If secular 574 changes in Indian Ocean [Ba] were occurring, we might expect models to make accurate 575 predictions for some datasets at the expense of others. In contrast, we note that model #3080 576 reproduces all testing datasets similarly well, with the exception of a subset of samples from SS259 577 in the deep Bay of Bengal. Here we observe that model #3080 predicts 18 % higher [Ba] than 578 observed by Singh et al. (2013) for the 42 samples between 1,000–3,000 m (Figs. 4B; 7A, B). 579 Interestingly, model #3080 correctly predicts [Ba] at nearby GEOSECS stations 445 and 446, also 580 in the Bay of Bengal, sampled some 31 years prior to SS259. We briefly consider three possibilities 581 for the origin of this regional model-data discrepancy. It may derive from the fact that model 582 #3080 does not include the features needed to correctly predict [Ba] in these samples. We view 583 this as the least likely possibility as model #3080 performs well for other samples from the northern 584 Indian Ocean, including samples shallower than 1,000 m from Singh et al. (2013). Another 585 possibility is that it could reflect an 18 % decrease in [Ba] in the deep Bay of Bengal since the 586 GEOSECS survey in the 1970's. Lastly, it could reflect differences in how in situ [Ba] was 587 measured, noting that Singh et al. (2013) opted for standard addition instead of isotope dilution. 588 We currently lack the data needed to confidently distinguish between these latter two possibilities.

589 A third factor concerns the limitations of ML itself. We note that no trained model was able to 590

achieve a MAPE better than ~6 %. This 6 % value may represent one of three things. First, it may

591 point toward an intrinsic limitation of Gaussian Process Regression. Other types of ML, such as 592 Decision Trees or Artificial Neural Networks, may be able to achieve superior accuracy, though 593 this was not investigated. Second, it may indicate that the 12 features investigated provide 594 insufficient information about [Ba] to achieve higher accuracy. We view this as unlikely given that 595 our earlier analysis showed that only six-nine features were needed to accurately simulate [Ba] 596 and that the 12 features tested have proved useful in other studies simulating dissolved tracer 597 distributions (e.g., Rafter et al., 2019; Sherwen et al., 2019; Roshan & DeVries, 2021). However, 598 this does not rule out the existence of other features beyond the 12 that we tested that are more 599 useful for predicting [Ba], only that we did not investigate them. Third, it is possible that the lowest 600 MAPE of ~ 6 % reflects the current limit of inter-laboratory uncertainty in determining [Ba]. We note that inter-laboratory uncertainties of 6–9 % were reported for the measurement of Ba:Ca in 601 602 carbonates (n = 10 labs; Hathorne et al., 2013). If the ~6 % MAPE derives from inter-laboratory 603 uncertainty, it is unlikely that further model refinements will improve the accuracy of [Ba] 604 predictions: the fundamental limitation is the data, not the model.

605 A final source of uncertainty concerns the computation of Ω_{barite} , which contains two further 606 sources of uncertainty: the thermodynamic model and the solubility coefficients used to calculate 607 $K_{\rm sp.}$ We calculated $\Omega_{\rm barite}$ based on the computation described by Rushdi et al. (2000), and our 608 approach yields similar values to their study and several others (e.g., Jeandel et al., 1996; Monnin 609 et al., 1999; see Appendix). The model used by Rushdi et al. (2000) is based on BaSO₄ solubility 610 data from Raju & Atkinson (1988), who note good agreement with the thermodynamic data of 611 Blount (1977). These solubility data were obtained based on experimentation with lab-made, 612 coarse-grained BaSO₄, which is unlikely to be wholly representative of the microcrystalline BaSO₄ precipitates found in seawater. Thus, the absolute values of Ω_{barite} calculated here may be subject 613 614 to eventual revision; however, the vertical (Fig. 1), spatial (Figs. 4–8), and whole-ocean (Fig. 9) 615 trends in Ω_{barite} are robust. Should new thermodynamic data for marine-relevant micron-sized 616 pelagic BaSO₄ become available, updated maps of Ω_{barite} could be recalculated using model #3080-617 derived [Ba] data. Given the nature of these uncertainties, we opted to calculate prediction 618 uncertainties for Ω_{barite} empirically by comparison to literature data (see Appendix). This yields a 619 value between 0.08 and 0.10, similar to the 10 % prediction error reported by Monnin et al. (1999).

620 **5.3. Barium in seawater: A global perspective**

621 Here we provide an overview of the main model features in [Ba], Ba* and Ω_{barite} , then outline three 622 possible applications of the model output.

623 5.3.1. Dissolved distribution of [Ba]

Model #3080 predictions show several interesting features in [Ba] (Figs. 5-8). The model 624 625 reproduces the expected nutrient-like distribution of [Ba] (Fig. 1C) and shows a general increase 626 in [Ba] along the Meridional Overturning Circulation: volume-weighted mean [Ba] increases from 67 to 90 to 106 nmol kg⁻¹ from the Atlantic to Indian to the Pacific Ocean, respectively. The model 627 628 also predicts some variation in shallow [Ba] that follows major surface-water currents, such as a 629 region of elevated [Ba] associated with the North Pacific Current, as well as low [Ba] in the western 630 North Atlantic associated with the Gulf Stream (Fig. 5B; Talley et al., 2011). However, these 631 features and the processes driving them await corroboration.

632 Considering the ocean as a whole, we can use our model to calculate the total Ba inventory of seawater. Using the mean oceanic [Ba] of 89 nmol kg⁻¹ and multiplying by the mass of seawater 633 634 $(1.37 \times 10^{21} \text{ kg})$ yields a total inventory of 122 ± 7 Tmol Ba, whereby the uncertainty is based on the 635 MAPE of model #3080 (6.0 %). This estimate of the total oceanic Ba inventory is between 11–21 636 % lower than existing estimates of 145 Tmol Ba (Dickens et al., 2003; Carter et al., 2020). Given 637 the range of probable global marine Ba fluxes between 18 (Paytan & Kastner, 1996) and 44 Gmol Ba vr⁻¹ (Rahman et al., 2022), our inventory estimate places the mean residence time of Ba in 638 639 seawater between 2,600–7,200 years.

640 5.3.2. The Ba–Si relationship

We now quantify spatial and vertical variations in the Ba–Si relationship, which we explore using Ba*. Star tracers, such as Ba*, highlight the processes affecting the distribution of a tracer by comparing it to another tracer that shares the same circulation (Gruber & Sarmiento, 1997). The concept has since been extended to study the processes affecting the distributions of many other bioactive elements, including Si (Si*, relative to N; Sarimento et al., 2004), cadmium (Cd*, 646 relative to P; Baars et al., 2014), zinc (Zn*, relative to Si; Wyatt et al., 2014). First defined by 647 Horner et al. (2015) for Ba, Ba* is analogous to other star tracers; it is a measure of Ba–Si 648 decoupling whereby larger values indicate larger Ba-Si deviations relative to expected mean ocean 649 behavior. Vertical or spatial differences in Ba and Si sources or sinks will drive variations in Ba^{*}, 650 as will any Ba:Si fractionation occurring during their combined cycling. Conversely, if all Ba and 651 Si cycling occurs in the same places (and with a fixed Ba:Si ratio), no Ba–Si decoupling will occur 652 and Ba* will exhibit conservative behavior. Since Ba and Si are cycled by different processes and 653 there are large vertical and spatial variations in the intensity of these processes (e.g., Bishop, 1989), 654 significant variations in Ba* are possible. We now explore these variations.

655 In the surface ocean, patterns of Ba* generally resemble those of [Ba] (Fig. 4). In large parts of the ocean, surface [Si] approaches 0 μ mol kg⁻¹; thus, variations in Ba* derive mostly from 656 657 variations in [Ba]. This is most evident when examining regions with significant terrestrial input 658 of Ba, such as from major rivers (Sect. 5.2.1) and from rivers and continental shelves in the Arctic 659 (e.g., Guay & Falkner, 1998; Whitmore et al., 2022; Fig. 5A). The Southern Ocean also exhibits 660 positive Ba*, though we suspect the mechanism is different. Here we observe a belt of waters with positive Ba* \approx +20 nmol kg⁻¹ centered on the Polar Frontal Zone—the region between the Antarctic 661 662 Polar Front and the Subantarctic Front (Orsi et al., 1995; Fig. 5A). Silicic acid is intensely stripped 663 from waters that transit northward through this region (e.g., Sarmiento et al., 2004), potentially 664 contributing to elevated Ba* at the sea surface. Dissolved [Ba] and Ba* then decrease to the north 665 of the Subantarctic front, partly driven by extensive particulate Ba formation in the frontal region 666 (e.g., Bishop, 1989).

667 At 1,000 m, the Atlantic, South Pacific, and southern Indian Oceans exhibit positive Ba* around +10 nmol kg⁻¹, whereas the North Pacific, Southern, and northern Indian Oceans are negative 668 between -10 to -20 nmol kg⁻¹ (Fig. 6C). The positive anomalies are likely related to the northward 669 670 spreading of southern-sourced intermediate waters that originate within the Polar Frontal Zone and 671 carry positive Ba* into the low latitudes (e.g., Bates et al., 2017). In the Atlantic, these values are 672 carried all the way to the north of the basin and return as North Atlantic Deep Water with only minor modifications to Ba* (\approx +10 nmol kg⁻¹; Figs. 6C, 7C, 8C). Negative Ba* in the North Pacific, 673 674 Southern, and northern Indian Ocean at 1,000 m likely reflects a mixture of hydrographic processes 675 and *in situ* processes. For example, the extensive region of negative Ba* in the North Pacific is

676 closely associated with North Pacific Intermediate Water, which originates in the Sea of Okhotsk 677 (Talley, 1991). While the specific mechanism sustaining this particular Ba* feature is unknown, it 678 most possibly reflects a combination of preferential removal of Ba relative to Si in the source water 679 formation region (such as from particulate Ba formation) and weak vertical mixing in the 680 subsurface North Pacific relative to lateral transports (e.g., Kawabe & Fujio, 2010). We suspect 681 that the negative Ba* values seen above 1,000 m in the northern Indian Ocean originate through 682 processes occurring internally within this basin, as the majority of the Indian Ocean below 1,000 683 m exhibits positive Ba*. A possible mechanism for these shallow negative Ba* anomalies may 684 relate to the relatively weak overturning transports (Talley, 2008) and strong particulate Ba cycle 685 north of 30 °S (Singh et al., 2013), though this awaits more detailed investigation.

686 Lastly, the Southern Ocean exhibits negative Ba* between -10 and -20 nmol kg⁻¹ from ≈ 200 m 687 water depth to the seafloor. These negative anomalies in Ba* appear to be associated with 688 Circumpolar Deep Water and, below that, Antarctic Bottom Water; the influence of the latter can 689 also be seen in near-bottom negative Ba* in the South Pacific, southern Indian, and South Atlantic 690 Oceans (Fig. 8C). As with the other basins, the origin of the negative Ba* waters in the Southern 691 Ocean likely reflects a combination of *in situ* and circulation-related phenomena. For example, in 692 the Southern Ocean, Si is only stripped at the very surface, whereas particulate Ba formation is 693 thought to be greatest in the mesopelagic (i.e., between 200–1,000 m; e.g., Stroobants et al. 1991). 694 Barite formation is generally considered to be related to the regeneration of particulate organic 695 matter (e.g., Chow & Goldberg, 1960), whereby the former consumes Ba and the latter releases 696 Si. Thus, intense organic matter remineralization and associated pelagic BaSO₄ precipitation could 697 contribute to negative Ba* in the mesopelagic Southern Ocean. Similarly, the Si cycle in the 698 Southern Ocean tends to 'trap' a significant fraction of the global Si inventory in the waters 699 circulating close to Antartica (e.g., Holzer et al., 2014). Since the calculation of Ba* depends on 700 both [Ba] and [Si], waters with elevated [Si] will exhibit lower Ba* whether or not there is elevated 701 Ba removal.

By 2,000 m, almost all of the ocean north of 50 °S exhibits positive Ba* (Fig. 7C). By 4,000 m,
the areal extent of the positive-Ba* waters shrinks to encompass the area north of 30 °S (Fig. 8C).
Despite covering a smaller area, the abyssal ocean exhibits the most positive Ba* values outside
of the surface of the Southern Ocean. The reasons for elevated and increasing Ba* between the

706 deep and abyssal oceans likely reflects a mixture of local and regional processes, and we offer two 707 speculative explanations for these patterns. First, Si trapping in the Southern Ocean potentially 708 renders most of the deep ocean away from Antarctica deficient in Si relative to Ba. Thus, much of 709 the ocean may exhibit more positive Ba* than the deep circum-Antarctic region due to processes 710 unrelated to Ba cycling. Second, the most positive Ba* values are generally found close to the 711 seafloor, rather than the mid-depths, especially in the North Pacific, the Peru and Chile Basins, 712 and the Philippine Sea. This may indicate a mechanism that preferentially removes Ba (relative to 713 Si) from the mid-depths, or input of Ba (relative to Si) close to the seafloor.

Systematic variations in Ba* arise due to differences in the marine biogeochemical cycles of Ba
and Si. While, in some cases, the specific drivers of these variations remains unresolved, our model
identifies multiple hotspots of Ba–Si decoupling that warrant additional study.

717 5.3.3. Barite saturation state of seawater

718 Here we show that our approach can predict Ω_{barite} with an MAE of 0.08, that our output is in 719 agreement with published values, and that the deep ocean, below 1,000 m, is at saturation with 720 respect to BaSO₄. By comparison to literature data, we estimate that our model achieves a typical 721 prediction uncertainty on Ω_{barite} of 0.08 (see Appendix). Accordingly, values of Ω_{barite} between 722 0.92–1.08 can be considered as 'BaSO₄ saturated,' whereas values of $\Omega_{\text{barite}} < 0.92 \text{ or} > 1.08$ indicate 723 under- or super-saturation, respectively. Global patterns in Ω_{barite} derived using our model are 724 similar to those reported by Monnin et al. (1999) and Rushdi et al. (2000). Readers looking for 725 detailed basin-by-basin descriptions of Ω_{barite} are directed to those studies. Briefly our model shows 726 that, excepting the high latitudes, the surface ocean is undersaturated with respect to BaSO₄ (i.e., $\Omega_{\text{barite}} < 0.92$). The lowest values of Ω_{barite} in the open ocean are observed in the hot, salty cores of 727 728 the Subtropical Gyres (Ω_{barite} between 0.1–0.2; Fig. 5D). Conversely, the cold and fresh polar 729 regions exhibit supersaturation at the sea surface, though there are important differences between 730 the Southern and Arctic Oceans. The Southern Ocean exhibits BaSO₄ saturation to depths around 731 2,000 m, whereas the Arctic Ocean switches to undersaturated conditions below the halocline 732 (~250 m). At 1,000 m, most of the North Pacific achieves saturation (or slight supersaturation) with respect to BaSO₄ (Fig. 6D) and at 2,000 m almost all of the ocean exhibits $\Omega_{\text{barite}} > 0.92$. The 733

main exceptions to this are the Atlantic Ocean, which is undersaturated at all depths, and the southern Indian Ocean between 35–50 °S (Fig. 7D). The South Pacific and Indian Oceans return to undersaturated conditions by 4,000 m, whereas parts of the North Pacific remain saturated to the seafloor (Fig. 8D). From a global perspective, the oceans are slightly undersaturated with respect to BaSO₄: volume-weighted mean $\Omega_{\text{barite}} = 0.82$; however, the ocean between 1,000–5,500 m exhibits $\Omega_{\text{barite}} \ge 0.92$ (Fig. 9). This result implies that the deep ocean, as a whole, is close to chemical equilibrium with respect to BaSO₄.

741 5.3.4. Model applications

742 In the spirit of maximizing model utility, we suggest three possible uses for model #3080 outputs. 743 First, the outputs can be used for model intercomparison and intercalibration. For example, a 744 number of statistical models, such as Optimum Multiparameter Optimization, have been 745 successfully used to study Ba cycling in the North Atlantic (Le Roy et al., 2018; Rahman et al., 746 2022), Southeast Pacific (Rahman et al., 2022), and Mediterranean Sea (Jullion et al., 2017). These 747 models can apportion the relative contributions of *in situ* biogeochemical cycling and conservative 748 mixing to observed [Ba]; however, accurate quantification of these processes requires a priori 749 knowledge of end-member water mass [Ba], which model #3080 can provide. Our model could 750 also be used to benchmark output from process-based models, such as Ocean Circulation Inverse 751 Models (e.g., John et al., 2020; Roshan & DeVries, 2021). Second, the output can be used for 752 interpolation purposes. Many groups investigated Ba partitioning into various types of marine 753 carbonates (see Sect. 1 for examples); however, these investigations are sometimes performed 754 without a co-located measurement of [Ba]. In these cases output from model #3080 could be used 755 to help calibrate specific substrates, such as deep-sea corals or benthic forams. This also avoids 756 the potential for circular reasoning whereby [Si] is used to estimate [Ba], which is then 757 reconstructed from the Ba:Ca ratio of carbonates to estimate [Si]. Third, the model output makes 758 testable predictions for regions of the ocean that have yet to be sampled by GEOTRACES-style 759 surveys. Several of these regions, such as the Southern Ocean, exhibit with sharp lateral and 760 vertical gradients in [Ba], Ba^{*}, and Ω_{barite} . Such gradients should be considered prime targets for 761 future process-oriented studies of marine Ba cycling.

762 6. Data availability

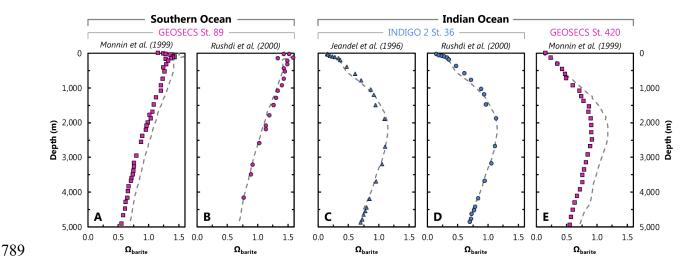
Data described in this manuscript can be accessed at the *Biological and Chemical Oceanography Data Management Office* under data doi:10.26008/1912/bco-dmo.885506.2 (Horner & Mete,
2023).

766 **7. Conclusions**

767 This study presents a spatially and vertically resolved global model of [Ba] determined using 768 Gaussian Process Regression machine learning. The model reproduces several known features of 769 the marine [Ba] distribution and makes testable predictions in regions that are yet to be sampled. 770 Analysis of the model output reveals the mean oceanic [Ba] is 89 nmol kg⁻¹, implying a total 771 marine Ba inventory of 122±7 Tmol. Using predictors from the World Ocean Atlas, we also 772 estimate the global distribution of Ba^{*} and $\Omega_{\text{barite.}}$ Both properties exhibit significant gradients that 773 can be systematically investigated in future studies. The mean oceanic Ω_{barite} is 0.82, though 774 between 1,000–5,500 m the mean is ≥ 0.92 , implying that the deep ocean is at equilibrium with 775 respect to BaSO₄. Our model output should prove valuable in studies of Ba biogeochemistry, 776 specifically for statistical- and process-based model validation, calibrating sedimentary archives, 777 and for identifying promising regions for further study. More broadly, our study demonstrates the 778 utility of using machine learning to accurately simulate the distributions of trace elements in 779 seawater. With minor adjustments, our approach could be employed to make predictions for other 780 dissolved tracers in the sea.

781 Appendix

Here we compare our results with published profiles of Ω_{barite} . Our results were calculated using the thermodynamic model of Rusdi et al. (2000), model #3080 [Ba], and WOA *T*, *S*, and pressure. Literature profiles of Ω_{barite} were calculated using one of three different thermodynamic models and *in situ* observations of [Ba], *T*, *S*, and pressure. In general, there is strong agreement between modeled and *in situ* Ω_{barite} whereby our model reproduces the shape of published profiles (Fig. A1). There are, however, some small systematic offsets between the various approaches, and we suspect that these derive from differences in the underlying thermodynamic models.



790Figure A1. Comparison of literature- (symbols) and Model #3080-derived (dashed line) values of791 $Ω_{barite}$. Panels A and B show profiles of $Ω_{barite}$ at GEOSECS St. 89 (60°0' S, 0°2' E). The other panels are792from the Indian Ocean: C and D are from INDIGO 2 St. 36 (6°9' S, 50°55' E) and E from GEOSECS St.793420 (0°3' S, 50°55' E), some ≈675 km north of INDIGO 2 St. 36.

We compare our model output with literature data Ω_{barite} at two locations in two basins (Fig. A1). These locations were chosen to ensure a fair comparison between studies; at each location, at least two studies calculated profiles of Ω_{barite} using the same underlying *in situ* data for [Ba], *T*, *S*, and pressure. Thus, any differences in modeled Ω_{barite} should derive from the thermodynamic model and not the input data. Likewise, literature profiles at these locations were based on calculations for pure, rather than strontian, BaSO₄, as in our study. Published profiles of Ω_{barite} were extracted graphically from each study using *WebPlotDigitizer* (Rohatgi, 2022). This extraction process may 801 introduce some minor scatter in the literature data, though this is relatively minor compared to the 802 range of variation in Ω_{barite} .

803 First, we examine profiles of Ω_{barite} reported for GEOSECS St. 89 in the Southern Ocean (Fig. A1; 804 Monnin et al., 1999; Rushdi et al., 2000). Modeled and published profiles show supersaturation in 805 the surface ocean and undersaturation below 2,000–2,500 m. Profiles from Rushdi et al. (2000) 806 show excellent agreement with Ω_{barite} calculated from model #3080 [Ba] and WOA T, S, and 807 pressure, with our output offset by a MAE of 0.06 (n = 22). Given that we use the same 808 thermodynamic model as Rushdi et al. (2000), the overall excellent agreement with their study is 809 not surprising. However, the result is nonetheless reassuring since our study uses mean annual 810 values for the various inputs, whereas Rushdi et al. (2000) utilized *in situ* data. There is a slightly 811 larger offset between our profile of Ω_{barite} and that calculated by Monnin et al. (1999), with our respective profile exhibiting an MAE of 0.13 (n = 41). This most likely reflects differences in the 812 813 underlying thermodynamic model and not the *in situ* data since our model reproduces the same 814 overall profile shape as Monnin et al. (1999). Likewise, both Monnin et al. (1999) and Rushdi et 815 al. (2000) used the same *in situ* input data and their results are highly comparable, albeit with an 816 offset similar to that between our results and Monnin et al. (1999).

Next we examine profiles of Ω_{barite} in the Indian Ocean for samples from INDIGO 2 St. 36 (Fig. A1; Jeandel et al., 1996; Rushdi et al., 2000). Profiles of Ω_{barite} show undersaturation at the surface, moderate supersaturation between 2,000–3,500 m, then return to undersaturated conditions down to the seafloor. Our profile shows overall excellent agreement with that of Jeandel et al. (1996), whereby a comparison of Ω_{barite} yields a MAE of of 0.03 (n = 21). Our profile shows similarly good agreement with Rushdi et al. (2000), whereby a comparison between our respective values of Ω_{barite} yields a MAE of 0.04 (n = 20).

We also compared our results with data from St. 420 of GEOSECS (Monnin et al., 1999), which is located ≈ 675 km north of INDIGO 2 St. 36 (Fig. 2). As with data from the Southern Ocean (GEOSECS St. 89), our profile data are offset to higher Ω_{barite} than those of Monnin et al. (1999), with slightly larger MAE of 0.16 (n = 29). However, our modeled Ω_{barite} is generally in much closer agreement with Monnin et al. (1999) above 1,100 m than below, equivalent to a MAE of 0.04 (n = 8) and 0.21 (n = 21), respectively. In this case it is more challenging to ascribe a unique cause

- 830 of the differences in calculated Ω_{barite} ; these offsets could relate to differences in the predictors or 831 the thermodynamic model.
- 832 We can use these comparisons to estimate the prediction uncertainty on our model-derived values
- 833 of Ω_{barite} . The MAE of the 133 comparisons shown in Fig. A1 yields a value of 0.10. However,
- there are different numbers of points in each profile; we thus believe it is more appropriate to
- average the MAE calculated for each of the five profiles, which yields a value of 0.08. Both values
- are similar to the 10 % prediction uncertainty reported by Monnin et al. (1999).
- 837 Overall, our ML-derived profiles of Ω_{barite} show excellent agreement with *in situ* data, both in
- 838 terms of profile shape and values of $\Omega_{\text{barite.}}$ We use this comparison to estimate the prediction
- 839 uncertainty on ML-derived values of Ω_{barite} , which we calculate as being between 0.08 and 0.10.
- 840 Should a revised thermodynamic model and/or improved BaSO₄ solubility coefficients become
- available, a new grid of Ω_{barite} could be calculated using Model #3080 [Ba] and WOA *T*, *S*, and
- 842 pressure data.

843 Author contributions

844 Project conceptualization and funding acquisition by T.J.H. Data curation, formal analysis,

- 845 investigation, and methodology by O.Z.M., A.V.S., H.H.K, and T.J.H. Data visualization by
- A.V.S. and T.J.H. Software provided by O.Z.M., A.V.S., H.H.K., and A.G.D. Writing (original draft) by O.Z.M. and T.J.H.; review and editing by A.V.S., H.H.K., A.G.D., L.M.W., A.M.S.,
- M_{C} and M_{D} Unit M_{D} Unit M_{D} Unit M_{D} M_{D} M
- 848 M.G., and W.D.L.

849 **Competing interests**

850 The authors declare that they have no conflict of interest.

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