

1 **Barium in seawater**

2 **Dissolved distribution, relationship to silicon, and barite saturation state** 3 **determined using machine learning**

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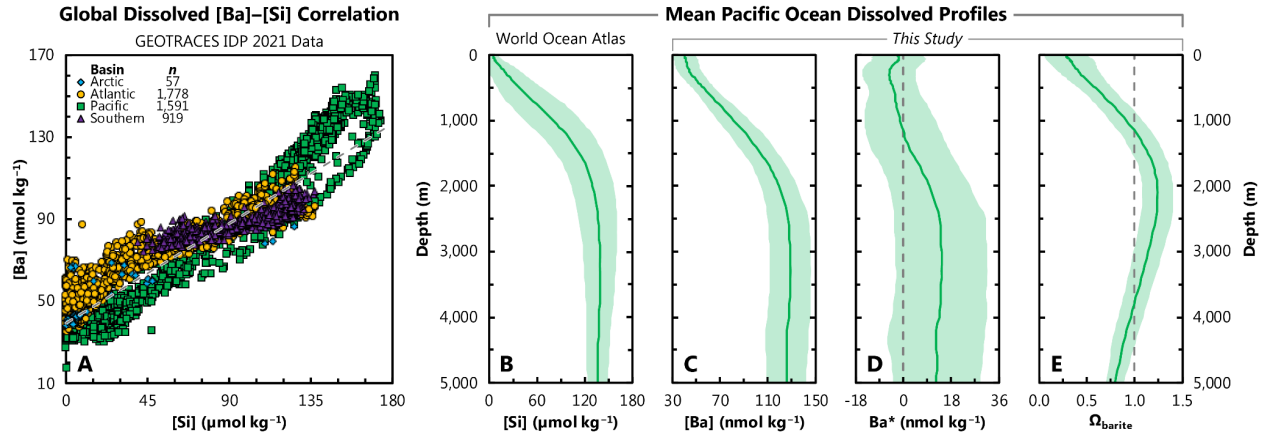
13 **Abstract**

14 Barium is widely used as a proxy for dissolved silicon and particulate organic carbon fluxes in
15 seawater. However, these proxy applications are limited by insufficient knowledge of the dissolved
16 distribution of Ba ([Ba]). For example, there is significant spatial variability in the barium–silicon
17 relationship, and ocean chemistry may influence sedimentary Ba preservation. To help address
18 these issues, we developed 4,095 models for predicting [Ba] using Gaussian Progress Regression
19 Machine Learning. These models were trained to predict [Ba] from standard oceanographic
20 observations using GEOTRACES data from the Arctic, Atlantic, Pacific, and Southern Oceans.
21 Trained models were then validated by comparing predictions against withheld [Ba] data from the
22 Indian Ocean. We find that a model trained using depth, temperature, salinity, as well as dissolved
23 dioxygen, phosphate, nitrate, and silicate can accurately predict [Ba] in the Indian Ocean with a
24 mean absolute percentage deviation of 6.0 %. We use this model to simulate [Ba] on a global basis
25 using these same seven predictors in the World Ocean Atlas. The resulting [Ba] distribution
26 constrains the Ba budget of the ocean to $122(\pm 7) \times 10^{12}$ mol and reveals systematic variability in
27 the barium–silicon relationship. We also calculate the saturation state of seawater with respect to
28 barite. In addition to revealing systematic spatial and vertical variations, our results show that the
29 ocean below 1,000 m is at equilibrium with respect to barite. We describe a number of possible
30 applications for our model output, ranging from use in biogeochemical models to paleoproxy
31 calibration. Our approach demonstrates the utility of machine learning to accurately simulate the
32 distributions of tracers in the sea and provides a framework that could be extended to other trace
33 elements.

34 **1. Introduction**

35 Barium (Ba) is a Group II trace metal that is widely applied in studies of modern and ancient
36 marine biogeochemistry, despite lacking a recognized biochemical function (e.g., Horner &
37 Crockford, 2021). These applications of Ba are based on two empirical correlations relating to its
38 dissolved and particulate cycles. The first correlation relates to the dissolved concentration of Ba,
39 hereafter [Ba], which is strongly correlated with that of the algal nutrient silicon (Si; as dissolved
40 silicic acid; Fig. 1; Chan et al., 1977). Unlike [Si], ambient [Ba] concentrations are faithfully
41 recorded by a number of marine carbonates, such as planktonic (e.g., Hönisch et al., 2011) and
42 benthic foraminifera (e.g., Lea & Boyle, 1990), surface- (e.g., Gonneea et al., 2017) and deep-sea
43 corals (e.g., Anagnostou et al., 2011; LaVigne et al., 2011), and mollusks (e.g., Komagoe et al.,
44 2018). Preservation of these signals means that the Ba content of carbonates can be related to the
45 Ba content of seawater and, by extension, that of Si. Accordingly, the Ba–Si proxy has been applied
46 to understand ocean nutrient dynamics on decadal (e.g., Lea et al., 1989) to millennial timescales
47 (e.g., Stewart et al., 2021).

48 The nutrient-like distribution of dissolved Ba in seawater is thought to be sustained by the second
49 empirical correlation, relating to cycling of particulate Ba. Particulate Ba in seawater occurs mostly
50 in the form of discrete, micron-sized crystals of the mineral barite ($\text{BaSO}_4(\text{s})$, barium sulfate; e.g.,
51 Dehairs et al., 1980; Stroobants et al., 1991). Pelagic BaSO_4 is an ubiquitous component of marine
52 particulate matter (e.g., Light & Norris, 2021) and constitutes the principal removal flux of
53 dissolved Ba from seawater (Paytan & Kastner, 1996). Pelagic BaSO_4 is thought to precipitate
54 within ephemeral particle-associated microenvironments that develop during the microbial
55 oxidation of sinking organic matter (e.g., Chow & Goldberg, 1960; Bishop, 1988). The flux of
56 particulate BaSO_4 to the seafloor is correlated with the flux of exported organic matter (e.g.,
57 Dymond et al., 1992; Eagle et al., 2003; Serno et al., 2014; Hayes et al., 2021). This correlation
58 means that the accumulation rate of sedimentary BaSO_4 —or its main constituent, Ba—can be used
59 to trace patterns of past organic matter export on timescales ranging from millenia to millions of
60 years (e.g., Bains et al., 2000; Paytan & Griffith, 2007; Schmitz, 1987; Schroeder et al., 1997).



61 **Figure 1. Distribution of barium in seawater.** **A.** Property–property plot showing the 4,345 co-located,
 62 core-feature complete dissolved data used in ML model training (Sect. 2). Sample locations shown in Figure
 63 2. Dashed line shows best-fit linear regression through these data, whereby $[Ba] = 0.54 \cdot [Si] + 39.3$. Panels
 64 **B., C., D.,** and **E.** show average Pacific Ocean dissolved depth profiles of $[Si]$, $[Ba]$, Ba^* , and Ω_{barite} ,
 65 respectively. Solid line denotes the arithmetic mean and the shaded region encompasses one standard
 66 deviation either side of the mean. Dashed line indicates $Ba^* = 0$ (**D**) and $\Omega_{barite} = 1$ (**E**).

67 While the Ba-based proxies are valuable, their applications are potentially limited by insufficient
 68 knowledge of the distribution of $[Ba]$. For example, there is significant vertical and spatial
 69 variability in the Ba–Si relationship (Sect. 3.3.; Fig. 1), which we quantify using Ba^* (barium-star;
 70 e.g., Horner et al., 2015):

$$71 \quad Ba^* = [Ba]_{in\ situ} - [Ba]_{predicted} \quad [Eq. 1]$$

72 where $[Ba]_{predicted}$ is based on the Ba–Si linear regression (Fig. 1):

$$73 \quad [Ba]_{predicted} = 0.54 \cdot [Si]_{in\ situ} + 39.3 \quad [Eq. 2]$$

74 Here, $[Si]_{in\ situ}$ has units of $\mu\text{mol kg}^{-1}$ and $[Ba]_{predicted}$ nmol kg^{-1} ; therefore, Ba^* also has units of
 75 nmol kg^{-1} . The vertical profile of Ba^* is rarely conservative (Fig. 1D) and these variations could
 76 introduce uncertainty in the reconstruction of $[Si]$ using Ba.

77 The relationship between sedimentary $BaSO_4$ accumulation rates and productivity also contains a
 78 significant degree of scatter (e.g., Serno et al., 2014; Hayes et al., 2021). Some of this scatter may
 79 relate to variability in $BaSO_4$ preservation, which is at least partially sensitive to ambient saturation

80 state, Ω_{barite} (e.g., Schenau et al., 2001; Singh et al., 2020; Fig. 1). The saturation state of a parcel
81 of water with respect to BaSO_4 is defined as:

$$82 \quad \Omega_{\text{barite}} = Q / K_{\text{sp}} \quad [\text{Eq. 3}]$$

83 where Q is the Ba and sulfate ion product and K_{sp} is the *in situ* BaSO_4 solubility product. Discerning
84 the importance of Ω_{barite} on BaSO_4 preservation has hitherto been challenging owing to the sparsity
85 of *in situ* [Ba] measurements. Accurately determining the global distribution of [Ba] would be
86 valuable for geochemists and oceanographers, and would enable a more thorough investigation of
87 the effects of preservation on BaSO_4 fluxes and refinement of the Ba–Si nutrient proxy.

88 A powerful way of interrogating oceanic element distributions is through modeling. Broadly, there
89 are two modeling approaches relevant for simulating [Ba]: mechanistic (i.e., theory driven) and
90 statistical modeling (i.e., data driven; e.g., Glover et al., 2011). In mechanistic or process-based
91 modeling, model outputs are derived from sets of underlying equations that are based on
92 fundamental theory. As such, mechanistic model outputs can be interrogated to obtain
93 understanding of processes and their sensitivities. However, creating a mechanistic model of the
94 marine Ba cycle requires embedding a biogeochemical model of BaSO_4 cycling within a
95 computationally expensive global circulation model. Although the computational cost associated
96 with building mechanistic models has been reduced by the development of ocean circulation
97 inverse models (e.g., DeVries, 2014; John et al., 2020), this approach still requires detailed
98 parametrizations of the marine Ba cycle, which do not currently exist. In contrast, statistical models
99 are based on extracting patterns from existing data and using those relationships to make
100 predictions. Statistical models encompass a wide variety of approaches ranging from regression
101 analysis to machine learning (ML). Of particular interest to our study are ML models, which can
102 make predictions without any explicit parameterizations of causal relationships. Machine learning
103 models are computationally efficient and can be highly accurate, though they offer limited
104 interpretability. Machine learning is increasingly being used to solve problems in Earth and
105 environmental sciences, including simulating the dissolved distribution of tracers in the sea (e.g.,
106 for cadmium, Roshan & DeVries, 2021; copper, Roshan et al., 2020; iodine, Sherwen et al. 2019;
107 nitrogen isotopes of nitrate, Rafter et al., 2019; and zinc, Roshan et al., 2018).

108 The goal of this study is to obtain an accurate simulation of [Ba], which ML makes possible even
109 in the absence of a process-level understanding of the marine Ba cycle. We tested thousands of
110 ML models that were trained using quality-controlled GEOTRACES data from the Arctic,
111 Atlantic, Pacific, and Southern Oceans, supplemented by Argo, satellite chlorophyll, and
112 bathymetry data products (Sect. 2.). Models were tested for their accuracy by simulating [Ba] in
113 the Indian Ocean and comparing predictions against observations made between 1977–2013. Since
114 no Indian Ocean data were seen by any of the models during training, we are able to identify
115 models with high generalization performance (Sect. 2.). We then identify an optimal set of
116 predictor variables, calculate model uncertainties, and simulate [Ba], Ba*, and Ω_{barite} on a global
117 basis (Sect. 5.). This result will be valuable for researchers interested in marine Ba cycling, and
118 demonstrates the utility of ML to tackle problems in marine biogeochemistry.

119 **2. Training and testing data**

120 Machine learning algorithms are adept at making accurate predictions of a target variable by
121 identifying relationships between variables within large data sets. However, making accurate
122 predictions first requires that a ML algorithm is trained on existing observations of that variable
123 alongside a number of other parameters. These other parameters, hereafter termed features, are an
124 important part of model training; features should encode information that may help the ML
125 algorithm predict [Ba], otherwise their inclusion may diminish model performance. Features
126 should also be well characterized in the global ocean, which allows ML models to make predictions
127 in regions beyond the initial training dataset. We selected 12 model features by considering the
128 tradeoff between feature availability and presumed predictive power (Table 1). While testing more
129 features may have resulted in a more accurate final model, we found that many observations of
130 [Ba] did not have corresponding data for multiple features; thus, including more features would
131 have meant fewer training data. Moreover, we find that including more than nine features can
132 actually diminish model performance. As such, we did not evaluate the predictive power of other
133 features beyond the 12 initially selected.

134 **Table 1. List of oceanographic parameters selected as model features.** The features tested were
 135 selected based on their presumed predictive power and geospatial coverage.

#	Parameter Name	Abbreviation	Units	Coverage*
1	Latitude	Lat.	degrees north (°N)	–
2	Longitude	Long.	degrees east (°E)	–
3	Sample collection depth	<i>z</i>	meters (m)	–
4	Temperature	<i>T</i>	degrees Celsius (°C)	97.44%
5	Salinity	<i>S</i>	unitless, but often written in ‘units’ of PSU or PSS	97.44%
6	Dissolved oxygen	[O ₂]	μmol kg ⁻¹	97.44%
7	Dissolved nitrate	[NO ₃]	μmol kg ⁻¹	97.44%
8	Dissolved phosphate	[PO ₄]	μmol kg ⁻¹	97.44%
9	Dissolved silicon (as silicic acid)	[Si]	μmol kg ⁻¹	97.44%
10	Maximum monthly mean mixed-layer depth	MLD	meters (m)	88.20%
11	Mean average annual surface chlorophyll	Chl. <i>a</i>	mg m ⁻³	93.95%
12	Bathymetry	Bathy.	meters (m)	100%

*Coverage values represent the percentage of data points within the World Ocean Atlas 2018 grid that have available data for a given parameter. Latitude, longitude, and depth have 100 % coverage as these features define the grid itself.

136 The 12 features used to predict [Ba] and their associated data sources are summarized in Table 1
 137 and described below. The first three features (latitude, longitude, depth) record geospatial
 138 information that defines the location of an observation in three-dimensional space. To avoid
 139 numerical discontinuities, latitude and longitude were introduced into the model as a
 140 hyperparameter consisting of the cosine and sine of their respective values (in radians). Data for
 141 features 1–3 were included in the sample metadata. Features 4–9 encode physical (temperature,
 142 salinity) and chemical (oxygen, nutrients) information that is routinely measured alongside [Ba].
 143 These data were generally available for the same bottle as the [Ba] measurements; however, when
 144 that was not the case, nutrient data were taken from the corresponding location during a separate

145 cast, or, in the case of oxygen, from linearly interpolated sensor data. The final three features are
146 independent of depth, meaning that all samples within a given vertical profile exhibit the same
147 value for MLD (mixed-layer depth), sea-surface chlorophyll *a*, and bathymetry. Features 10–12
148 were drawn from several data sources. A climatology of MLD (feature 10) was compiled using
149 the Argo database (Holte et al., 2017). We selected maximum monthly mean MLD as the feature
150 of interest, as this appears to be the spatiotemporal scale most relevant for influencing [Ba]
151 distributions (Bates et al., 2017). Feature 11 represents a blended SeaWiFS and MODIS
152 climatology of chlorophyll *a* that was obtained from the Copernicus Marine Environment
153 Monitoring Service (CMEMS, 2021). We calculated the mean annual chlorophyll *a* for each grid
154 cell in the data product and log transformed the data to reduce parameter weighting (e.g., Rafter et
155 al., 2019). Data for MLD and chlorophyll *a* were extracted at the location of [Ba] observations
156 using nearest-neighbor interpolation and their values logged in the master record. Bathymetric
157 information (feature 12) was extracted from one of two sources. Our preferred source was the
158 sample metadata, which generally included a value for bathymetry. For samples lacking
159 bathymetric information, we used nearest-neighbor interpolation to extract a value from the
160 *ETOPO5* Global Relief Model (National Geophysical Data Center, 1993). Occasionally, the
161 *ETOPO5*-extracted bathymetry was shallower than the deepest observation of [Ba] in a given
162 vertical profile. In such cases, the bathymetry logged in the master record was set to 1.01 times the
163 depth of the deepest observation in that profile.

164 The [Ba] data from the Indian Ocean were collected from several, primarily pre-GEOTRACES
165 sources (Table 2). As such, these data were generally incomplete for the 12 features used to train
166 the ML models. Rather than using a mixture of *in situ* and interpolated data, we decided to
167 interpolate all Indian Ocean data for parameters 4–12. Data for parameters 4–9 were linearly
168 interpolated from the nearest vertical profile in the World Ocean Atlas 2018 (WOA; Boyer et al.,
169 2018; García et al., 2018a; 2018b; Locarnini et al., 2018; Zweng et al., 2018) and values for MLD
170 and chlorophyll *a* were extracted from the aforementioned data products using nearest-neighbor
171 interpolation. Bathymetric information was obtained from either the WOA or *ETOPO5*. For the
172 vast majority of most samples, bathymetry was taken as the arithmetic mean of the maximum
173 depth of the nearest vertical profile in the WOA and the depth at the standard level below. For
174 example, if the maximum depth at a station was 950 m, the bathymetry was recorded as 975 m,
175 which is the mean of levels 46 (950 m) and 47 (1,000 m). For profiles with a maximum depth of

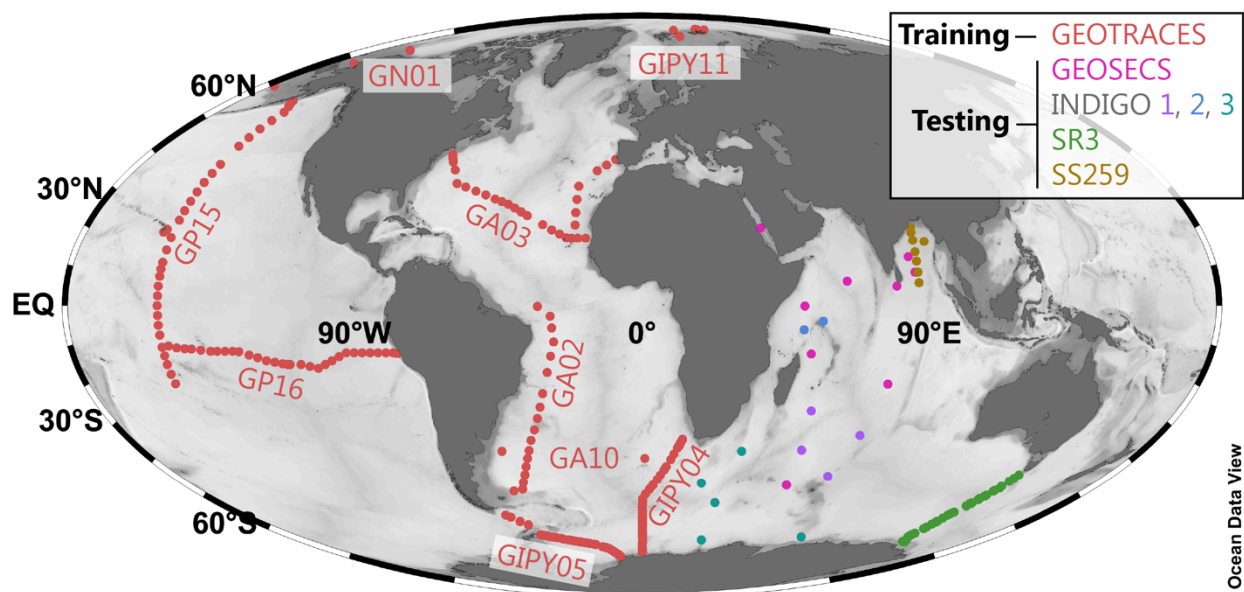
176 5,500 m—level 102, the lowest in the WOA—bathymetry was recorded as either 5,550 m or the
 177 nearest-neighbor interpolated value from *ETOPO5*, whichever was deeper.

178 **Table 2. Data sources.** Information regarding the source of [Ba] incorporated into the master record.

Purpose	Region	Expedition ID	Data source	Data Originators (if unpublished)
Model training	South Atlantic (Meridional)	GA02	GEOTRACES IDP 2017 (Schlitzer et al., 2018)	Jose M. Godoy
	North Atlantic (Zonal)	GA03	Rahman et al., 2022	
	South Atlantic (Zonal)	GA10	Horner et al., 2015; Bates et al., 2017; Hsieh & Henderson, 2017; Bridgestock et al., 2018	
	Southern Ocean (Meridional)	GIPY04	GEOTRACES IDP 2017 (Schlitzer et al., 2018)	Frank Dehairs
	Southern Ocean (Zonal)	GIPY05	Hoppema et al., 2010	
	Arctic	GIPY11	Roeske et al., 2012	
		GN01	Whitmore et al., 2022	
	Pacific (Meridional)	GP15	GEOTRACES IDP 2021 (GEOTRACES IDP Group, 2021)	Laura Whitmore, Melissa Gilbert, Emilie Le Roy, Tristan Horner, Alan Shiller
	Subtropical South Pacific (Zonal)	GP16	Rahman et al., 2022	
Model testing	Indian Ocean	GEOSECS	Craig & Turekian (1980)	
		INDIGO 1	Jeandel et al. (1996)	
		INDIGO 2		
		INDIGO 3		
		SR3	Jacquet et al. (2004)	
		SS259	Singh et al. (2013)	

179

180 This data ingestion process resulted in a master record containing 5,502 observations of [Ba] that
 181 also contained a corresponding value for all 12 core features (Table 1). The record was then split
 182 into a Pareto partition: the first partition was used for ML model training (4,345 observations, 79
 183 % of data; Fig. 1A) and the second for model testing (1,157 data; 21 %). This partitioning was
 184 determined based on the basin from which the sample was collected; data from the Arctic, Atlantic,
 185 Pacific, and Southern Oceans were used in model training, whereas the 1,157 [Ba] data from the
 186 Indian Ocean were reserved for model testing (Table 2; Fig. 2). This location-based separation of
 187 training and testing data was chosen to minimize overfitting, which can occur when the training–
 188 testing separation is randomly assigned (see Sect. 3.2.).



189 **Figure 2. Geographical distribution of the training and testing data.** The 4,345 core-feature complete
 190 training data (red; Fig. 1) are from the GEOTRACES 2021 Intermediate Data Product (GEOTRACES IDP
 191 Group, 2021); GEOTRACES expedition identifiers are noted next to each section. The $n = 1,157$ testing
 192 data from the Indian Ocean are color-coded by expedition. Data sources listed in Table 2.

193 3. Methods

194 In the following subsections we discuss details of the specific ML algorithm that was used for
 195 model development (Sect. 3.1.), explain the model training and testing process (Sect. 3.2.), and
 196 describe how a global prediction of [Ba] was obtained and interrogated (Sect. 3.3.).

197 **3.1. Algorithm selection and training**

198 We opted for supervised ML using a Gaussian Process Regression learner, implemented in
199 MATLAB. This particular ML algorithm is non-parametric, kernel-based, and probabilistic, which
200 means that it does not make strong assumptions about the mapping function, can handle
201 nonlinearities, and takes into account the effect of random occurrences when making predictions.
202 Gaussian Process Regression algorithms are widely used in geostatistics, where it is often referred
203 to as ‘kriging’ (e.g., Cressie, 1993; Rasmussen & Williams, 2006; Glover et al., 2011). This type
204 of algorithm is ideal when working with continuous data that also contains a certain level of noise,
205 such as from measurement uncertainty or oceanographic variation. The MATLAB function,
206 `fitrgp`, was used for model training. A full list of the parameter selections used in `fitrgp` is
207 provided in Table S1. All predictors were normalized and standardized to have a mean of zero and
208 a standard deviation of unity. This process places all parameters on the same relative range and
209 reduces scale dependencies.

210 A selection of the training data were used to train 4,095 different machine learning models with
211 the goal of finding a model that could accurately simulate the global distribution of [Ba]. The
212 number of models derives from the number of features investigated; each model uses a unique
213 combination of the 12 features in Table 1 and our testing followed a factorial design whereby each
214 feature was either enabled or disabled. This design yields a total of 2^{12} unique feature combinations
215 (i.e., $\text{levels}^{\text{features}}$); however, since it is not possible to train a model with no features enabled, the
216 final number of unique, trainable, ML models with ≥ 1 features is $2^{12}-1=4,095$. The full experiment
217 list is provided in Section 6. Each of the 4,095 models was trained using the same data and with
218 the same function parameters described in Table S1.

219 **3.2. Assessing model performance**

220 Model performance—accuracy and generalizability—was assessed during two phases: training
221 and testing. During model training, the 4,345 observations of [Ba] from the Arctic, Atlantic,
222 Pacific, and Southern Oceans were randomly split into two folds: a training fold containing 80 %
223 of the observations, and a holdout fold containing the other 20 %. Model accuracy was assessed
224 by comparing model-predicted [Ba] against observed [Ba] for the 20 % of the data in the holdout

225 fold. We then performed additional testing to establish model generalizability. A significant
226 problem in supervised ML, and particularly Gaussian Process Regression learning, is overfitting:
227 models may fit the noise in the training data, leading to poor generalization performance
228 (Rasmussen & Williams, 2006). Since our goal was to develop a global model of [Ba] using
229 regional training data, we deemed it especially important to identify generalizable models.
230 Generalizable models were identified through a testing process involving regional cross-
231 validation; each trained model was used to predict [Ba] for the 1,157 samples from the Indian
232 Ocean and model predictions were again compared against observations. Importantly, no [Ba] data
233 from the Indian Ocean were seen by any of the models during training. This process helped to
234 identify models that may have been overfit to the training data and can further be used to calculate
235 generalization errors (Sect. 4.1).

236 The accuracy of trained models was determined by comparing ML model predictions against
237 withheld data and calculating the mean absolute error (MAE) and mean absolute percentage error
238 (MAPE), defined as:

$$239 \quad \text{MAE} = \frac{\sum_{i=1}^n |[Ba]_{\text{predicted}} - [Ba]_{\text{observed}}|}{n} \quad [\text{Eq. 4}]$$

240 and:

$$241 \quad \text{MAPE} = \frac{100\%}{n} \sum_{i=1}^n \left| \frac{[Ba]_{\text{predicted}} - [Ba]_{\text{observed}}}{[Ba]_{\text{observed}}} \right| \quad [\text{Eq. 5}]$$

242 respectively, where n is the sample size.

243 Models with lower accuracy exhibit higher errors, whereas models with high accuracy have lower
244 errors. We calculated MAE and MAPE for every possible feature combination, which enables
245 quantification of how specific features affect model performance. Likewise, we calculated errors
246 for each model on predictions made during training (i.e., for the holdout fold) and during model
247 testing (i.e., during regional cross-validation; Fig. 3). This information is used to quantify
248 generalization performance; low errors for both training and testing indicate models that are both
249 accurate and generalizable, whereas models with low training errors and high testing errors might
250 indicate models that are overfit to the training data.

251 **3.3. Global predictions**

252 A select number of models with low MAE and MAPE were used to simulate [Ba] on a global
253 basis. The process by which we selected these models is described in Section 5.1. Global
254 simulations were performed on the same grid as the WOA, which was also used as the data source
255 for features 1–9 (Boyer et al., 2018). The WOA is a $1^\circ \times 1^\circ$ resolution data product with around
256 41,000 stations that contain up to 102 depth levels spanning 0–5,500 m in 5, 25, 50, or 100 m
257 increments. Data for features 10–12 (MLD, chlorophyll *a*, and bathymetry) were also resampled
258 to the WOA grid using the same sources and interpolation methods as described for the Indian
259 Ocean testing data in Section 2. Model outputs were visualized using Ocean Data View software
260 (ODV; Figs. 5–8; Schlitzer, 2023).

261 A selection of the most accurate models of [Ba] were then used to simulate Ba* and Ω_{barite} . Star
262 tracers, such as Ba*, are valuable for illustrating processes that influence the cycling of elements
263 in the ocean. First defined for N–P decoupling (N*; Gruber & Sarmiento, 1997) star tracers show
264 variations whenever there are differences in the sources and sinks of the two elements being
265 compared. If there are no differences in sources and sinks, the tracer will show conservative
266 behavior because both elements share the same circulation. Barium-star is based on Ba–Si
267 decoupling and was first defined by Horner et al. (2015). The definition of Ba* is shown in
268 Equations 1 and 2. The coefficients in Equation 2 are based on data from the GEOTRACES 2021
269 Intermediate Data Product and specifically the subset of these data shown in Figure 1. These
270 coefficients differ from previous formulations of Ba* that were based primarily on [Ba] and [Si]
271 data from the Southern and Atlantic Oceans (e.g., Horner et al., 2015; Bates et al., 2017). The
272 global distribution of Ba* was determined by calculating $[\text{Ba}]_{\text{predicted}}$ (Eq. 2) using $[\text{Si}]_{\text{in situ}}$ from
273 the WOA 2018 (García et al., 2018b). The values of $[\text{Ba}]_{\text{in situ}}$ was taken from the ML model output
274 and $[\text{Ba}]_{\text{predicted}}$ was subtracted from this to yield Ba* (Eq. 1).

275 Values of Ω_{barite} were computed using the method described by Rushdi et al. (2000), summarized
276 in Equation 3. The numerator, Q , represents the *in situ* Ba and sulfate ion product and, in this
277 formulation, depends only on [Ba] and $[\text{SO}_4^{2-}]$ molality. The denominator, K_{sp} , depends on T , S ,
278 and z (i.e., pressure) and is calculated in two steps: *in situ* T and S are used to calculate the
279 stoichiometric solubility product and then this value is modified by calculating the effect of
280 pressure on partial molal volume and compressibility, which are functions of T and z . As with the

281 calculation of Ba^* , values of $[Ba]_{in\ situ}$ were obtained from ML models and co-located data for T ,
282 S , and z were extracted from the WOA (Locarnini et al., 2018; Zweng et al., 2018). Sulfate
283 concentrations were assumed to be conservative with respect to S using $[SO_4^{2-}] = 29.26\text{ mmol kg}^{-1}$
284 ¹ when salinity = 35 PSU. This latter assumption likely breaks down in certain environments (e.g.,
285 where $[SO_4^{2-}]$ reduction occurs); as such, our model is not used to predict Ω_{barite} in restricted
286 basins, such as the Black Sea or Caspian Sea. Given that our estimates of Ω_{barite} exhibit a MAE of
287 0.08 (Appendix), we believe that values of Ω_{barite} between 0.92 and 1.08 are indicative of ‘perfect’
288 saturation with respect to $BaSO_4$.

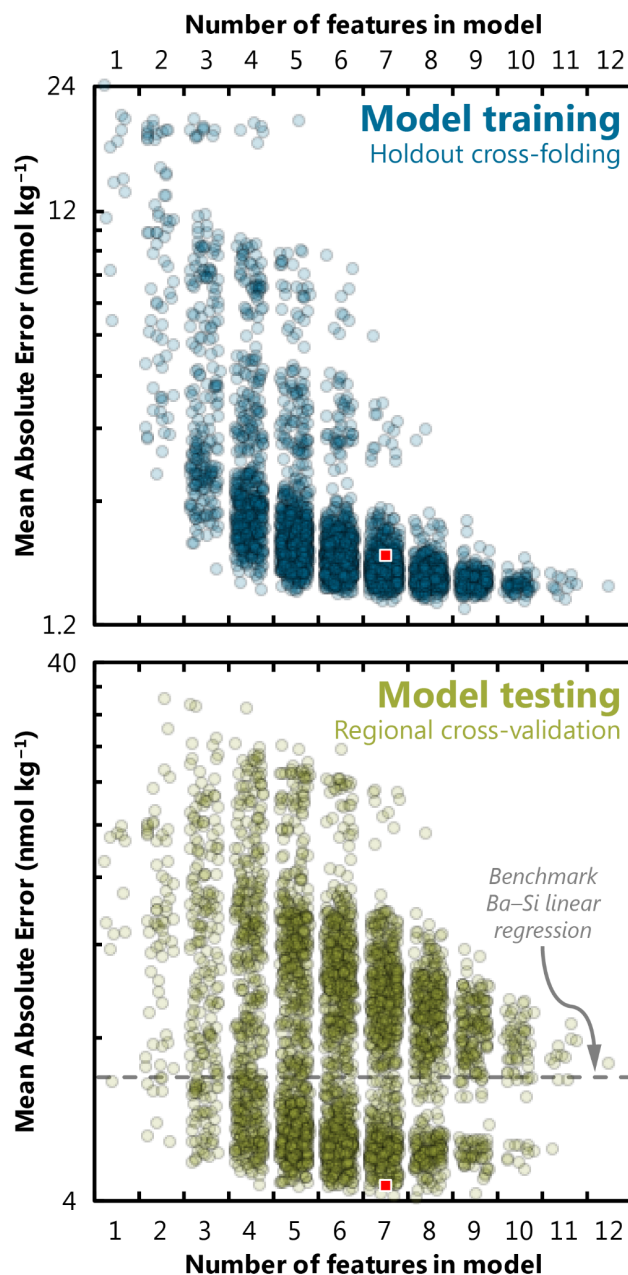
289 Output from the most accurate ML models was then used to calculate mean $[Ba]$ and Ω_{barite} for
290 each basin, for a series of prescribed depth bins, and for the global ocean. This calculation was
291 performed by weighting each cell in the model output by its volume, which ensures a fair
292 comparison between any two points in the model output. We then subdivided the global ocean into
293 five sub-basins: Arctic, Atlantic, Indian, Pacific, and Southern. Basin boundaries were defined as
294 per Eakins & Sharman (2010), though we merged the Mediterranean and Baltic Seas into the
295 Atlantic and considered the South China Sea as part of the Pacific Ocean. Neither $[Ba]$ nor Ω_{barite}
296 were simulated in the Black or Caspian Seas and thus these regions are not included in the global
297 mean calculations.

298 **4. Results**

299 **4.1. Factors affecting model accuracy**

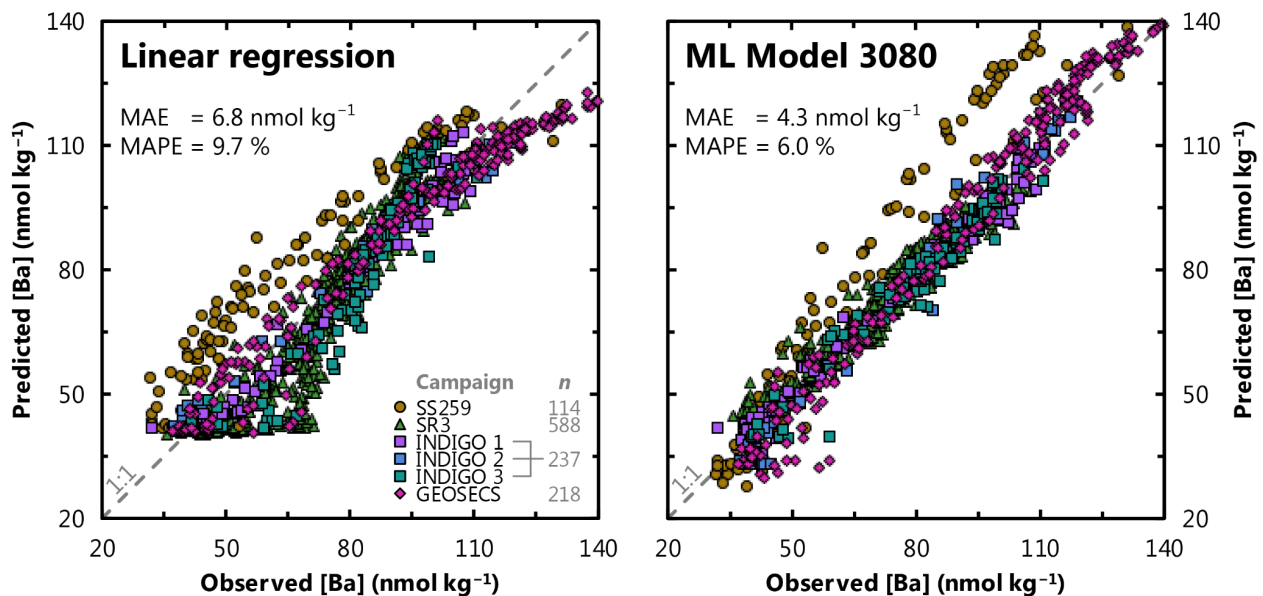
300 Here we examine how model performance is influenced by the number and nature of features
301 included during training. We consider model performance in terms of accuracy and
302 generalizability, which we quantify using MAE (Eq. 4). We first explore how the number of
303 features influences model performance (Fig. 3). Here we see that increasing the number of features
304 generally improves the accuracy of trained models; however, the response differs depending on
305 whether accuracy is calculated based on comparison to the holdout fold (i.e., during model
306 training) or to the withheld Indian Ocean data (i.e., during model testing). When considering only
307 the holdout fold, trained models predict $[Ba]$ with a high level of accuracy—the mean, median,
308 and most-accurate trained models achieve a MAE of 2.4, 1.7, and 1.3 nmol kg^{-1} , respectively.
309 Similarly, increasing the number of features almost always improves model accuracy; the MAE of

310 the most accurate model for a given number of features decreases from 6.5 to 1.3 nmol kg⁻¹ as the
311 number of features is increased from one to nine, at which point MAE plateaus between 1.4–1.5
312 nmol kg⁻¹ for models with 10–12 features (Fig. 3A).



313 **Figure 3. Effect of feature addition on ML model accuracy.** Accuracy was quantified for each of the
314 4,095 trained models and quantified here using MAE (note log scale, which differs between panels). The
315 accuracy of trained models is shown for random holdout cross-validation during training (top) and for
316 regional cross-validation during testing (bottom). Square indicates the performance of our favored predictor
317 model, #3080 (see Fig. 4, Sect. 5.1). The accuracy of the Ba–Si linear regression benchmark is shown as
318 a dashed line in the lower panel (MAE = 6.8 nmol kg⁻¹). To illustrate data density, points have been
319 randomly positioned within their respective bin and plotted with 80 % transparency.

320 Moving to the regional cross-validation, the overall performance of models is lower; the same
 321 4,095 trained models achieve a mean, median, and most-accurate MAE for the Indian Ocean
 322 dataset of 8.8, 7.9, and 4.0 nmol kg⁻¹, respectively. For comparison, if [Ba] was estimated for these
 323 same 1,157 Indian Ocean samples using the linear [Ba]–[Si] relationship (Fig. 1) and ambient [Si]
 324 as the only predictor, this linear model would achieve a MAE of 6.8 nmol kg⁻¹. Thus, there are
 325 1,687 ML models that achieve a superior accuracy to existing methods for estimating [Ba],
 326 offering an improvement of as much as 41 % (Fig. 4). However, regional cross-validation also
 327 shows that the addition of more features may, in fact, degrade model performance. The MAE of
 328 the most accurate model for a given number of features decreases from 6.6 to 4.0 nmol kg⁻¹ as the
 329 number of features is increased from one to eight. As the number of features is increased from 9–
 330 12, the MAE of the most-accurate models increases monotonically from 4.1 to 7.1 nmol kg⁻¹. The
 331 overall lower performance of trained models during regional cross validation—and the observation
 332 that many of the feature-rich models perform worse than models with fewer features—is indicative
 333 of certain models being over-fit to the training data. Together, these observations suggest that the
 334 optimum number of features needed to accurately predict [Ba] is between six and nine.



335 **Figure 4. Comparison of existing and ML methods to estimate [Ba] in seawater.** Left panel shows the
 336 performance benchmark: predicted [Ba] for the Indian Ocean testing data using the [Ba]–[Si] linear
 337 regression and ambient [Si] as the sole predictor. Right panel shows predicted [Ba] using ML model 3080,
 338 which improves on existing methods by more than 37 %. Perfect correspondence between predictions and
 339 observations is indicated by the dashed line marked '1:1.' Data locations and sources are shown in Fig. 2

340 and Table 2, respectively; n refers to the number of testing data for each campaign. Mean Absolute Error
 341 (MAE; Eq. 4) and Mean Absolute Percentage Error (MAPE; Eq. 5) are noted for both models.

342 We also evaluated the nature of the predictors used to estimate [Ba]. The full factorial experiment
 343 design enables us to perform comparisons between all models that contained a certain feature and
 344 all of those that did not (Sect. 3.1). We quantified the effect of adding a feature by comparing the
 345 absolute and percentage change in MAE relative to the mean MAE of the two sets of models. This
 346 comparison was performed three times: for all 4,095 models based on the holdout cross-folded
 347 training data, for all models using the regionally cross-validated testing data, and again for the
 348 testing data, but only considering those 1,687 models that achieved a superior accuracy compared
 349 to the [Ba]–[Si] linear regression model (Table 3).

350 **Table 3. Feature addition analysis.** Effect of each feature on model performance for Training and Testing
 351 datasets. Model performance is quantified using MAE, thus all columns have units of nmol kg^{-1} unless
 352 otherwise shown. The Testing analysis is further subdivided into a comparison of all models and ‘good’
 353 models, meaning those that achieved superior accuracy than the Ba–Si linear regression (Fig. 1).

Feature	Training			Testing						
	All models ($n = 4,095$)			All models ($n = 4,095$)			Good models ($n = 1,687$)			
	Mean MAE of models with feature	Mean MAE of models without feature	Relative change in MAE	Mean MAE of models with feature	Mean MAE of models without feature	Relative change in MAE	Mean MAE of models with feature	Mean MAE of models without feature	Relative change in MAE	Share of models with feature
[Si]	1.71	3.03	-56%	7.08	10.6	-39%	5.06	5.50	-8.3%	63%
z	1.83	2.90	-45%	7.94	9.70	-20%	5.05	5.44	-7.4%	55%
[O ₂]	2.03	2.71	-29%	8.25	9.39	-13%	5.14	5.33	-3.8%	54%
T	1.78	2.96	-50%	7.61	10.0	-27%	5.17	5.31	-2.8%	59%
[NO ₃]	2.09	2.65	-24%	8.27	9.36	-12%	5.16	5.30	-2.7%	53%
[PO ₄]	2.11	2.63	-22%	8.24	9.40	-13%	5.17	5.30	-2.4%	53%
S	2.02	2.72	-29%	8.67	8.97	-3.5%	5.23	5.23	0.0%	53%
Bathy.	2.30	2.44	-6.1%	8.55	9.08	-6.0%	5.23	5.22	0.2%	51%
Chl.	2.25	2.48	-10%	8.67	8.97	-3.5%	5.24	5.22	0.4%	50%
MLD	2.31	2.43	-4.8%	8.69	8.95	-3.0%	5.24	5.21	0.5%	50%
Lat.	2.16	2.58	-18%	8.13	9.51	-16%	5.32	5.11	4.0%	54%
Long.	2.17	2.57	-17%	11.4	6.24	58%	6.45	5.19	22%	3%

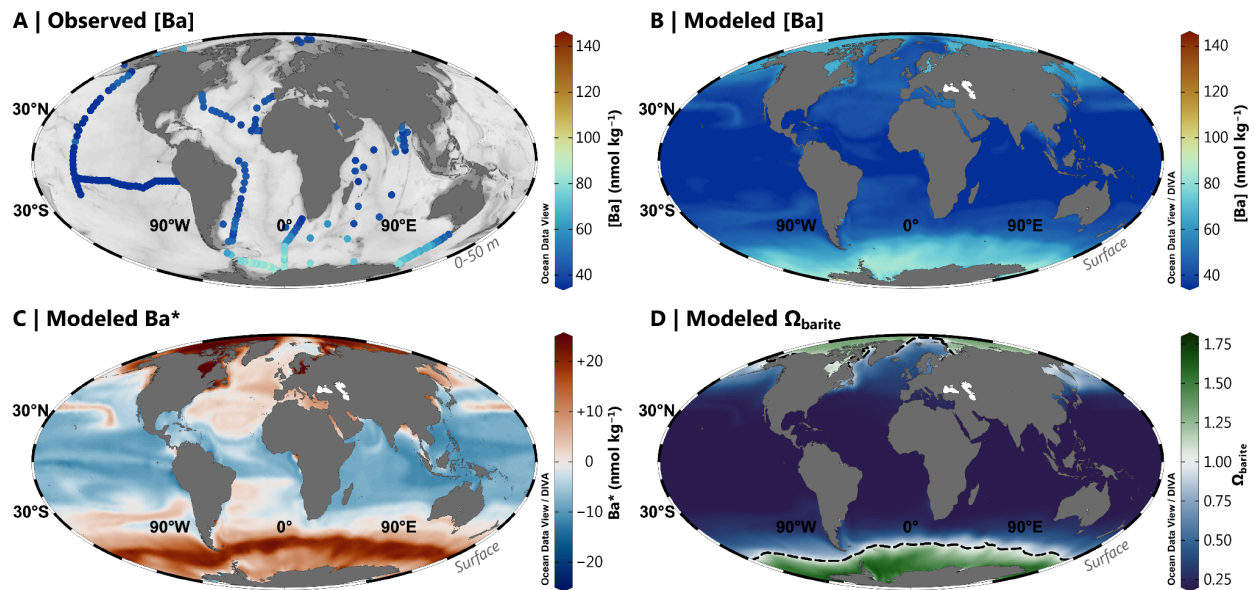
354 This analysis yields three main results. When considering only the holdout cross-folded training
 355 data, the addition of any of the 12 features improves model performance by between -4.8 and -56
 356 %. Excepting longitude, similar across-the-board improvements were observed when considering
 357 only the testing data, though the improvements for most features were more modest (between -3.0
 358 and -39 %). If considering only the ‘good’ models, six features improved model performance by

359 -2.4 and -8.3 % ($[\text{PO}_4]$, $[\text{NO}_3]$, T , $[\text{O}_2]$, z , and $[\text{Si}]$), five degraded model performance by $+0.2$ to
 360 $+22$ % (bathy., Chl. a, MLD, lat., and long.), and salinity had no significant effect (Table 3).

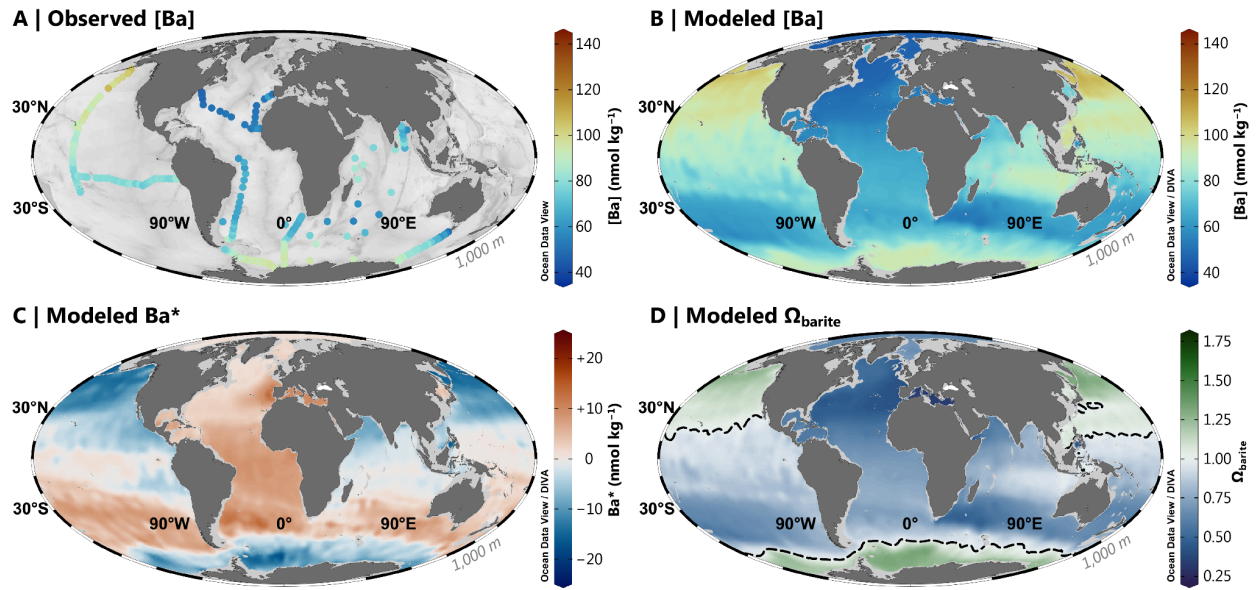
361 Overall, our results indicate that between six and nine features will result in an accurate and
 362 generalizable ML model of $[\text{Ba}]$, and that $[\text{PO}_4]$, $[\text{NO}_3]$, T , $[\text{O}_2]$, z , $[\text{Si}]$, and possibly S , are likely
 363 to be included as predictors in such a model.

364 4.2. Model outputs

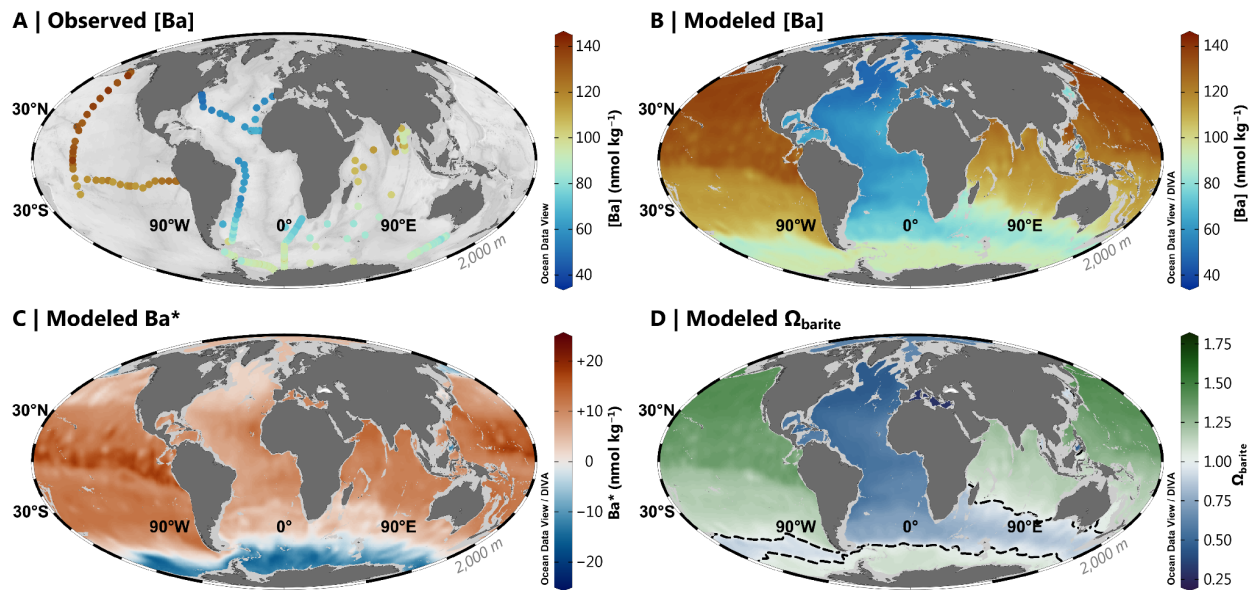
365 Almost 1,700 models achieved superior accuracy compared to the Ba–Si linear regression
 366 benchmark of 6.8 nmol kg^{-1} . We winnow this list to a single model, #3080, in the next section.
 367 We henceforth refer to model #3080 as our favored predictor model, which achieves a MAE of
 368 4.3 nmol kg^{-1} using z , T , S , $[\text{O}_2]$, $[\text{PO}_4]$, $[\text{NO}_3]$, and $[\text{Si}]$ as predictors (Fig. 4). Model #3080 is used
 369 to simulate $[\text{Ba}]$, Ba^* , and Ω_{barite} on a global basis and to calculate whole-ocean averages. Surface
 370 plots showing the model outputs for the sea surface, 1,000 m, 2,000 m, and 4,000 m are shown in
 371 Figures 5, 6, 7, and 8, respectively.



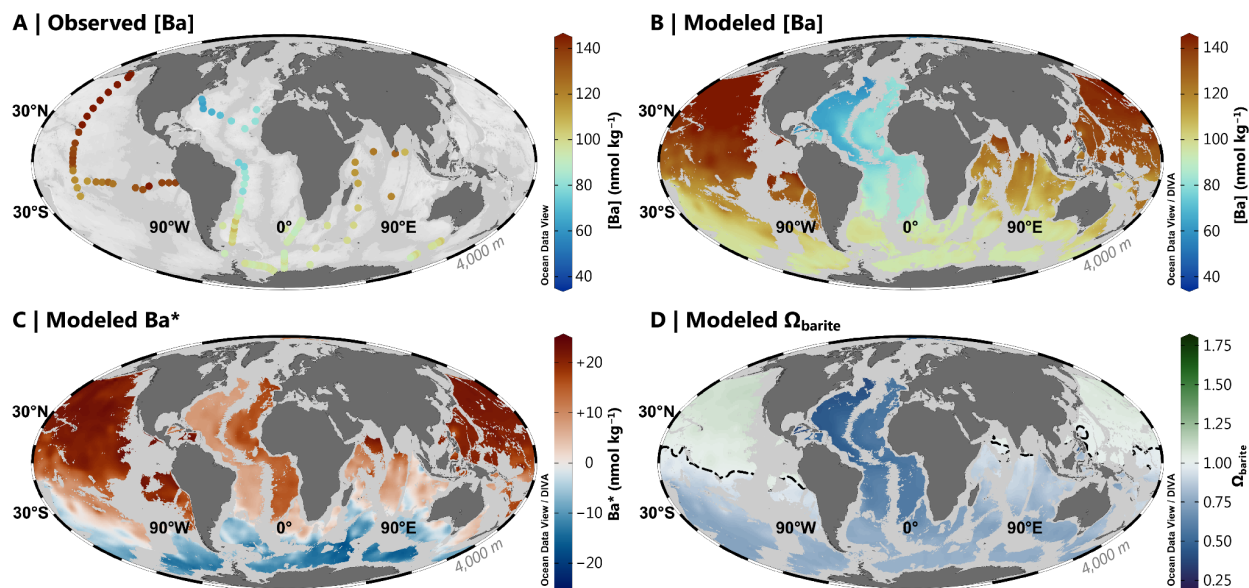
372 **Figure 5. Barium at the sea surface.** Observed $[\text{Ba}]$ between 0–50 m (A); Model 3080 $[\text{Ba}]$ (B), Ba^* (C),
 373 and Ω_{barite} (D). The dashed line in Panel D indicates the BaSO_4 saturation horizon (i.e., $\Omega_{\text{barite}} = 1.0$). Panels
 374 A and B use the *roma* color map, whereas Panels C and D use *vik* and *cork*, respectively (Crameri, 2018).
 375 Color palettes and parameter ranges are the same for the respective panels in Figure 6–8.



376 **Figure 6. Barium at 1,000 m.** Observed [Ba] (A); Model 3080 [Ba] (B), Ba* (C), and Ω_{barite} (D). The dashed
 377 line in Panel D indicates the BaSO₄ saturation horizon.



378
 379 **Figure 7. Barium at 2,000 m.** Observed [Ba] (A); Model 3080 [Ba] (B), Ba* (C), and Ω_{barite} (D). The dashed
 380 line in Panel D indicates the BaSO₄ saturation horizon.



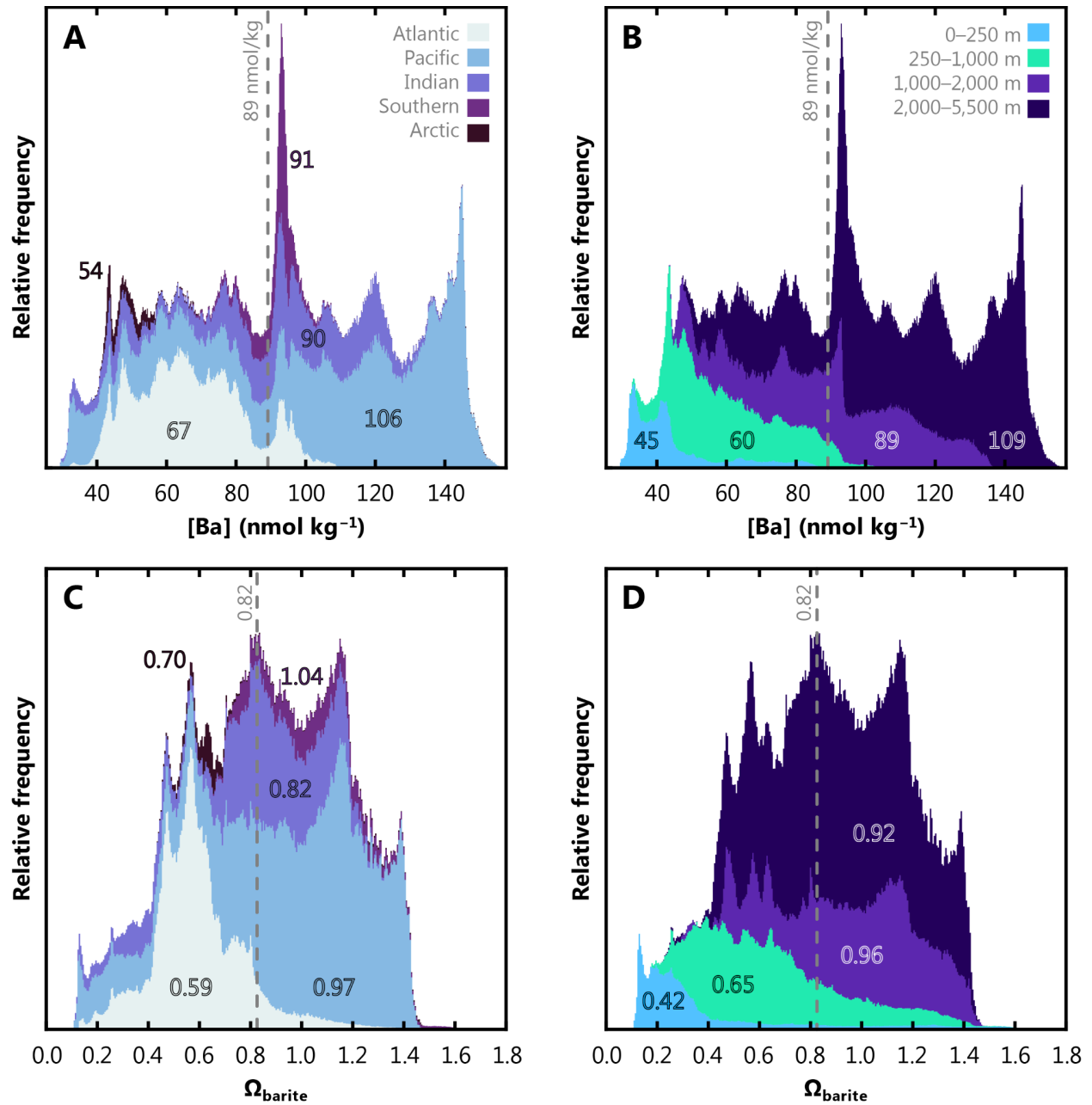
381 **Figure 7. Barium at 4,000 m.** Observed [Ba] (A); Model 3080 [Ba] (B), Ba* (C), and Ω_{barite} (D). The dashed
 382 line in Panel D indicates the BaSO₄ saturation horizon.

383 Model #3080 contains 3,302,570 predictions for each of [Ba], Ba*, and Ω_{barite} (Sect. 6). Assuming
 384 that the MAPE and MAE are good estimates of the prediction error, we estimate that modeled [Ba]
 385 and Ba* have uncertainties of 6.0 % and 4.3 nmol kg⁻¹, respectively. Uncertainties on Ω_{barite} were
 386 estimated by comparison to literature data, which yields a MAE of 0.08. These estimates are
 387 discussed in more detail in Section 5.2 and the Appendix.

388 Modeled [Ba] ranges from 26.2–156.8 nmol kg⁻¹ and the data exhibit an unweighted mean of 72.0
 389 nmol kg⁻¹. The range of model #3080 predictions is within the range of [Ba] encountered in the
 390 4,345 training data (17.1–159.8 nmol kg⁻¹). This is an important consideration when assessing the
 391 accuracy of Gaussian Process Regression models, and we provide additional discussion of this
 392 point in the Supplement. Based on our formulation (Eqs. 1, 2), Ba* varies from -27.2 to +27.9
 393 nmol kg⁻¹ and possesses an unweighted mean of +2.4 nmol kg⁻¹. Values of Ω_{barite} vary from 0.11
 394 to 1.70 and exhibit an unweighted mean of 0.75. To account for the different volumes represented
 395 by each cell in the WOA grid, we constructed a volume-weighted mean of [Ba] and Ω_{barite} for the
 396 ocean as a whole, for each ocean basin, and for a series of prescribed depth bins (Fig. 9). Looking
 397 at the ocean as a whole, the probability density function of [Ba] roughly resembles a uniform
 398 distribution, with a mean ocean [Ba] of 89 nmol kg⁻¹ (Fig. 9A). Within this mean is considerable

399 spatial and vertical variation. For example, the Arctic Ocean exhibits the lowest volume-weighted
400 mean [Ba] of 54 nmol kg⁻¹, whereas mean Pacific [Ba] = 106 nmol kg⁻¹. The Indian Ocean exhibits
401 a similar mean [Ba] (90 nmol kg⁻¹) to the mean of the global ocean. Shallower than 1,000 m, [Ba]
402 infrequently exceeds 100 nmol kg⁻¹, whereas concentrations <45 nmol kg⁻¹ are rare below 1,000
403 m (Fig. 9B).

404 The probability density function of volume-weighted Ω_{barite} is more similar to a normal
405 distribution, albeit with a slight negative skew. Volume-weighted mean oceanic Ω_{barite} is 0.82. The
406 Arctic, Atlantic, and Indian Oceans are, on average, undersaturated with respect to BaSO₄, all
407 exhibiting mean $\Omega_{\text{barite}} \leq 0.82$. In contrast, the Pacific and Southern Oceans are within uncertainty
408 of saturation, with mean Ω_{barite} of 0.97 and 1.04, respectively (Fig. 9C). Values of $\Omega_{\text{barite}} < 0.2$ are
409 mostly restricted to the upper 250 m, whilst values of Ω_{barite} exceeding 1.5 are exceptionally rare,
410 found only in the upper 1,000 m of the Southern Ocean. Lastly, Ω_{barite} tends to increase between
411 the 0–250 m, 250–1,000 m, and 1,000–2,000 m depth bins, increasing from 0.42, to 0.65, and 0.96,
412 respectively. Average Ω_{barite} in the deepest bin (2,000–5,500 m) is slightly lower, with a mean
413 value of 0.92 (Fig. 9D). Given the accuracy of our model-derived Ω_{barite} predictions (0.08 to 0.10),
414 the ocean between 1,000–5,500 m is within uncertainty of BaSO₄ equilibrium.



415 **Figure 9. Stacked, volume-weighted histograms showing the relative frequency distribution of**
 416 **dissolved [Ba] (A, B) and Ω_{barite} (C, D) in the global ocean.** The left column shows data grouped by basin
 417 and the right column shows data grouped by a prescribed depth bin. Numbers in each panel display the
 418 mean property value for that bin. Dashed line shows the global mean.

419 **5. Discussion**

420 **5.1. Identification of the optimal predictor model**

421 Choosing a single, optimal model configuration is challenging given the sheer number of skillful
422 ML models. Below we winnow the list from 4,095 to a single model (#3080). We base our
423 winnowing primarily on the results of the regional cross-validation performed in the Indian Ocean,
424 rather than from the errors determined from random holdout cross folding of the training data. We
425 believe that there are three strong reasons for winnowing in this way. First, Gaussian Process
426 Regression Learners tend to fit the noise in the training data, meaning that the training error is
427 significantly lower than the generalization error (Rasmussen & Williams, 2006). Indeed, trained
428 models showed overall lower performance during testing compared to training, which we believe
429 is evidence of overfitting (Fig. 3, Table 3). Second, a generalizable global model should be able to
430 make predictions in regions where it has not already learned anything about the target variable.
431 Our regional cross-validation approach satisfies this consideration since no Indian Ocean data were
432 seen during model training. Third, the Indian Ocean is an ideal basin for testing as it exhibits the
433 full diversity of features expected to influence [Ba] (riverine inputs, oxygen-minimum zones,
434 coastal upwelling, etc.) and constitutes ≈ 20 % of the global ocean volume. Likewise, the Indian
435 Ocean captures most of the range in [Ba] seen elsewhere in the ocean (Fig. 9); this likely reflects
436 the input of Atlantic waters through the Aughulas leakage, transport of old Pacific waters via the
437 Indonesian Throughflow, and northward spreading of mode and intermediate waters from the
438 Southern Ocean. We thus assume that the Indian Ocean testing errors are a good approximation of
439 the generalization error, which we now use to winnow the list of models.

440 Our results show that 1,687 of the 4,095 ML models (41 %) produce more accurate predictions of
441 [Ba] than the benchmark Ba–Si linear regression using [Si] as the sole predictor (Fig. 3, Table 3).
442 We focus our winnowing on these 1,687 models as they are superior to existing methods for
443 estimating [Ba] in seawater. Focusing only on these ‘good’ models reveals significant differences
444 in the information content of the 12 features tested. For example, the inclusion of spatial
445 information in the form of latitude and longitude significantly degrades mean model performance
446 by between +4.0 and +22 %, respectively. While bathymetry, chlorophyll *a*, and mixed-layer depth
447 exhibited only minor influences, they were nonetheless deleterious to mean model performance
448 by between +0.2 to +0.5 % (Table 3). Only [PO₄], [NO₃], *T*, [O₂], *z*, and [Si] consistently improved

449 the mean ML model, which corresponds to model #3112 (testing MAE of 4.3 nmol kg⁻¹).
450 However, visual inspection of model #3112 output reveals that it does not reproduce expected
451 near-shore surface plumes of elevated [Ba] close to certain major rivers (see Supplement). Though
452 volumetrically minor, riverine inputs are a geochemically important component of the marine Ba
453 cycle, and the existence of nearshore Ba plumes underpins a major proxy application of Ba. Near-
454 shore riverine influence is easily discerned by low *S*; we thus explored output from model #3080,
455 which is identical to model #3112, but includes *S* as a seventh feature during training. Models
456 #3080 and #3112 exhibit identical statistical performance for the testing data (MAE = 4.3 nmol
457 kg⁻¹; Fig. S1) and make similar predictions for mean marine [Ba] and Ω_{barite} (89 nmol kg⁻¹ and
458 0.82, respectively; see Supplement). The similar performance of the two models is consistent with
459 *S* exerting a near-negligible impact on overall model performance (Table 3). Despite this small
460 effect, model #3080 is better able to reproduce riverine [Ba] plumes compared to model #3112
461 (see Supplement). We therefore consider model #3080 to be our best estimate of marine [Ba].
462 Model #3080 achieves a MAPE of 6.0 %, which represents a 39 % improvement over existing
463 methods to estimate [Ba] (Fig. 4). We henceforth consider model #3080 as our optimal predictor
464 model, which we use to simulate [Ba], Ba*, and Ω_{barite} in Figures 5–9.

465 **5.2. Model validation**

466 We now explore the validity of model #3080 in terms of its oceanographic consistency, the sources
467 of uncertainty that affect its accuracy, and potential limitations of the model output. We find that
468 model #3080 reproduces the major known features of the marine [Ba] distribution and makes
469 testable predictions for regions that are yet to be sampled.

470 *5.2.1. Visual inspection of model output*

471 Visual inspection of model output is an important component of data analysis considering the
472 limits of statistical tests (see e.g., Anscombe, 1973). Models may produce statistically satisfactory
473 fits to the testing data, but the oceanic realism of the output is also important to consider. Modeled
474 [Ba] should display patterns consistent with related oceanographic properties and exhibit smooth

475 vertical and spatial variations (Boyle & Edmond, 1975). Predicted [Ba] from model #3080 does
476 indeed show smooth and systematic spatial and vertical variations that also resembles sparse
477 observations (Figs. 4–8).

478 Model #3080 also shows systematic increases in [Ba] close to land, especially near the mouths of
479 major rivers (Fig. 4). This is reassuring given that elevated sea-surface [Ba] close to rivers is both
480 widely reported and one of the major proxy applications of Ba: reconstructing spatiotemporal
481 patterns of terrestrial runoff by measuring the Ba:Ca ratio of carbonates (e.g., Sinclair &
482 McCulloch, 2004; LaVigne et al., 2016). For example, model #3080 correctly identifies elevated
483 [Ba] near the Ganges–Brahmaputra (Singh et al., 2013), Río de la Plata (GEOTRACES IDP
484 Group, 2021), and Yangtze outflows (Cao et al., 2021). Model #3080 also predicts elevated sea-
485 surface [Ba] in the Gulf of Guinea where several rivers discharge, including the Niger River; the
486 Eastern Tropical Atlantic associated with the Congo River (Edmond et al., 1978; Zhang et al.,
487 2023); and in the Gulf of St. Lawrence (St. Lawrence River; see Supplement for additional details
488 and figures). Except for the Congo River, these predictions of elevated near-shore [Ba] await
489 corroboration. Interestingly, model #3080 does not predict elevated [Ba] at all major river mouths;
490 neither the Mississippi nor Amazon Rivers are associated with significant increases in sea-surface
491 [Ba] (see Supplement). The reasons for the lack of elevated [Ba] near the outflow of these two
492 rivers is less clear. It is possible that the model is simply inaccurate in these regions, though we
493 have no particular reason to believe that this is the case. Alternatively, it may reflect seasonal
494 variations in Ba release that are not captured by our mean annual model (e.g., Joung & Shiller,
495 2014). It could also indicate that these particular rivers are not major *net* sources of Ba to the
496 surface ocean, which might be the case if dissolved Ba is being retained in the catchment (e.g.,
497 Charbonnier et al., 2020) or estuary (e.g., Coffey et al., 1997).

498 Overall, model #3080 makes accurate, oceanographically consistent predictions of [Ba] in the
499 Indian Ocean using input data from the WOA. Model #3080 also makes a number of testable
500 predictions of [Ba] in regions lacking direct observations. Given that these predictions were made
501 using the same model and the same WOA inputs, we believe that it is reasonable to assume that
502 model #3080 output is an accurate representation of mean annual global [Ba].

503 5.2.2. Quantifying uncertainties

504 We now describe and, where possible, quantify two possible sources of uncertainty to our ML
505 model output. Before doing so, we describe how uncertainty is quantified as well as the uncertainty
506 of existing approaches. Certain ML models, such as Gaussian Process Regression, offer low
507 interpretability, meaning it is not possible to assess uncertainty using a conventional error
508 propagation. Thus, all model uncertainties are assessed *post hoc*, by comparing predictions against
509 observations. Our preferred metrics are MAE and MAPE (Eqs. 4, 5). Existing approaches for
510 estimating [Ba] result in a wide range of uncertainties. At the low end, the uncertainty associated
511 with measuring [Ba] in seawater represents a fundamental limit to the accuracy of any model. A
512 number of analysts report measurement uncertainties in the range of 1–2 % (e.g., Pyle et al., 2018;
513 Cao et al., 2020). This level of intra-laboratory uncertainty is typical for [Ba] data obtained using
514 isotope dilution–inductively coupled plasma mass spectrometry, and applies to GEOTRACES-era
515 datasets and to much of the training data from the Indian Ocean. However, intra-laboratory
516 uncertainty is typically much smaller than inter-laboratory uncertainty, which is often between 6–
517 9 % (e.g., Hathorne et al., 2013). At the upper end, the benchmark Ba–Si linear regression achieves
518 a MAPE of 9.7 % in the Indian Ocean (Fig. 4). Thus, useful ML models of [Ba] should achieve
519 MAPE between 1–10 %. Indeed, our favored predictor model, #3080, achieves a MAPE of 6.0 %.

520 Now we consider two factors that contribute to the observed 6.0 % uncertainty: realization
521 uncertainty and uncertainties in the training data. The realization uncertainty stems from the fact
522 that two models trained on the same training dataset—even with the exact same subset of model
523 features—will produce slightly different predictions. This is due to the holdout cross-folding
524 process used during model training, which partitions the training dataset into random subsets (
525 Sect. 3.1.). Thus, the training process results in a slightly different trained model each time the
526 model is realized. We quantified the realization uncertainty by training select models 100 times
527 and calculating the relative standard deviation of the different predictions of [Ba] for the 3.3
528 million values in the output. This uncertainty is small; the median, mean, and maximum realization
529 uncertainty was 0.03 %, 0.04 %, and 0.32 % variability in modeled [Ba].

530 Next we consider uncertainties in the training data. As noted above, many labs report uncertainties
531 on [Ba] measurements of 1–2 %, while inter-laboratory differences may be up to a factor of five
532 larger. However, this does not consider any uncertainties associated with the other physical and

533 chemical features used to predict [Ba]. In general, these supporting measurement uncertainties
534 should be small: all overboard sensors are regularly calibrated and biogeochemical properties in
535 GEOTRACES are determined using established methods that are based on GO-SHIP best practices
536 (Hood et al., 2010). Moreover, all GEOTRACES sections include crossover stations that are
537 intended to facilitate intercalibration of all parameters, including those used here to predict [Ba]
538 (Fig. 2; Cutter, 2013). The WOA, MLD, Chl. *a*, and bathymetry data products are similarly
539 subjected to stringent quality review and so we consider it unlikely that these data contribute
540 systematic biases. We believe that the most likely source of uncertainty relates to the fact that all
541 predictor information used for model testing in the Indian Ocean was derived from time-averaged
542 data products, whereas [Ba] was derived from *in situ* measurements. We made this decision
543 because the *in situ* data were incomplete for all 12 core features (Table 1), and this would have
544 necessitated interpolation for some features and not others. Since all models were tested using the
545 same predictor information, the comparison process should avoid systematic errors, though this
546 does not preclude temporal variability, described next.

547 5.2.3. Other considerations

548 We now consider four other factors that potentially contribute to the uncertainty of the model
549 output: short- and long-term temporal variations, limitations of ML, and uncertainties regarding
550 the thermodynamic properties of BaSO₄. Short-timescale variability in [Ba] may affect how
551 models were evaluated, though this effect is difficult to quantify. In principle, the trained models
552 should be able to resolve seasonal variations in [Ba] since they were trained on *in situ* physical and
553 chemical data. In contrast, model predictions in the Indian Ocean were made using annual average
554 physical and chemical conditions and then evaluated by comparing these predictions against *in*
555 *situ* [Ba]. The temporal mismatch between Indian Ocean observations and predictions is unlikely
556 to be significant in the deep ocean, where seasonal variations are minor and the Ba residence time
557 is longest (e.g., Hayes et al., 2018). Seasonal variations are, however, likely to matter more for the
558 surface ocean. We were able to minimize some of the impact of these uncertainties by using long-
559 term averages of Chl. *a* and the maximum monthly mean MLD during model training and testing.
560 Significant seasonal mismatches for other parameters are unavoidable given that [Ba] data are too
561 sparse to develop a time-resolved model. We suspect that these variations are most likely to be

562 significant for boundary sources rather than biogeochemical cycling of Ba; significant
563 biogeochemical drawdown of surface [Ba] over seasonal timescales appears to be rare (e.g., Esser
564 & Volpe, 2002), whereas there are large seasonal variations in river discharge that impact near-
565 shore [Ba] (e.g., Samanta & Dalai, 2016). These suspicions could be tested using a model with
566 better than $1 \times 1^\circ$ spatial resolution, which—in theory—is possible with model #3080, so long as
567 similarly high-resolution data are provided for the six predictors utilized by this model (z , T , S ,
568 $[O_2]$, $[PO_4]$, $[NO_3]$, and $[Si]$). While it is challenging to precisely quantify seasonal uncertainties,
569 we note that model #3080 performs well at low [Ba], which is found mostly near the surface, where
570 seasonal variations should exhibit the largest effects. Likewise, seasonal variations will have only
571 a minor effect on our calculations of global mean [Ba] or Ω_{barite} (Fig. 8).

572 Long-term variability in [Ba] may also influence model performance, since the testing data from
573 the Indian Ocean were collected between 1977 (GEOSECS) and 2008 (SS259; Fig. 2). If secular
574 changes in Indian Ocean [Ba] were occurring, we might expect models to make accurate
575 predictions for some datasets at the expense of others. In contrast, we note that model #3080
576 reproduces all testing datasets similarly well, with the exception of a subset of samples from SS259
577 in the deep Bay of Bengal. Here we observe that model #3080 predicts 18 % higher [Ba] than
578 observed by Singh et al. (2013) for the 42 samples between 1,000–3,000 m (Figs. 4B; 7A, B).
579 Interestingly, model #3080 correctly predicts [Ba] at nearby GEOSECS stations 445 and 446, also
580 in the Bay of Bengal, sampled some 31 years prior to SS259. We briefly consider three possibilities
581 for the origin of this regional model–data discrepancy. It may derive from the fact that model
582 #3080 does not include the features needed to correctly predict [Ba] in these samples. We view
583 this as the least likely possibility as model #3080 performs well for other samples from the northern
584 Indian Ocean, including samples shallower than 1,000 m from Singh et al. (2013). Another
585 possibility is that it could reflect an 18 % decrease in [Ba] in the deep Bay of Bengal since the
586 GEOSECS survey in the 1970's. Lastly, it could reflect differences in how *in situ* [Ba] was
587 measured, noting that Singh et al. (2013) opted for standard addition instead of isotope dilution.
588 We currently lack the data needed to confidently distinguish between these latter two possibilities.

589 A third factor concerns the limitations of ML itself. We note that no trained model was able to
590 achieve a MAPE better than $\sim 6\%$. This 6 % value may represent one of three things. First, it may
591 point toward an intrinsic limitation of Gaussian Process Regression. Other types of ML, such as

592 Decision Trees or Artificial Neural Networks, may be able to achieve superior accuracy, though
593 this was not investigated. Second, it may indicate that the 12 features investigated provide
594 insufficient information about [Ba] to achieve higher accuracy. We view this as unlikely given that
595 our earlier analysis showed that only six–nine features were needed to accurately simulate [Ba]
596 and that the 12 features tested have proved useful in other studies simulating dissolved tracer
597 distributions (e.g., Rafter et al., 2019; Sherwen et al., 2019; Roshan & DeVries, 2021). However,
598 this does not rule out the existence of other features beyond the 12 that we tested that are more
599 useful for predicting [Ba], only that we did not investigate them. Third, it is possible that the lowest
600 MAPE of ~6 % reflects the current limit of inter-laboratory uncertainty in determining [Ba]. We
601 note that inter-laboratory uncertainties of 6–9 % were reported for the measurement of Ba:Ca in
602 carbonates ($n = 10$ labs; Hathorne et al., 2013). If the ~6 % MAPE derives from inter-laboratory
603 uncertainty, it is unlikely that further model refinements will improve the accuracy of [Ba]
604 predictions: the fundamental limitation is the data, not the model.

605 A final source of uncertainty concerns the computation of Ω_{barite} , which contains two further
606 sources of uncertainty: the thermodynamic model and the solubility coefficients used to calculate
607 K_{sp} . We calculated Ω_{barite} based on the computation described by Rushdi et al. (2000), and our
608 approach yields similar values to their study and several others (e.g., Jeandel et al., 1996; Monnin
609 et al., 1999; see Appendix). The model used by Rushdi et al. (2000) is based on BaSO₄ solubility
610 data from Raju & Atkinson (1988), who note good agreement with the thermodynamic data of
611 Blount (1977). These solubility data were obtained based on experimentation with lab-made,
612 coarse-grained BaSO₄, which is unlikely to be wholly representative of the microcrystalline BaSO₄
613 precipitates found in seawater. Thus, the absolute values of Ω_{barite} calculated here may be subject
614 to eventual revision; however, the vertical (Fig. 1), spatial (Figs. 4–8), and whole-ocean (Fig. 9)
615 trends in Ω_{barite} are robust. Should new thermodynamic data for marine-relevant micron-sized
616 pelagic BaSO₄ become available, updated maps of Ω_{barite} could be recalculated using model #3080-
617 derived [Ba] data. Given the nature of these uncertainties, we opted to calculate prediction
618 uncertainties for Ω_{barite} empirically by comparison to literature data (see Appendix). This yields a
619 value between 0.08 and 0.10, similar to the 10 % prediction error reported by Monnin et al. (1999).

620 We can calculate Ω_{barite} to a high degree of precision; however, there are numerous uncertainties
621 pertaining to ML-predicted [Ba], the BaSO_4 solubility coefficients used to calculate K_{sp} , and the
622 thermodynamic model used in the computation of Ω_{barite} (Sect. 5.2.). Thus,

623 **5.3. Barium in seawater: A global perspective**

624 Here we provide an overview of the main model features in [Ba], Ba^* and Ω_{barite} , then outline three
625 possible applications of the model output.

626 *5.3.1. Dissolved distribution of [Ba]*

627 Model #3080 predictions show several interesting features in [Ba] (Figs. 5–8). The model
628 reproduces the expected nutrient-like distribution of [Ba] (Fig. 1C) and shows a general increase
629 in [Ba] along the Meridional Overturning Circulation: volume-weighted mean [Ba] increases from
630 67 to 90 to 106 nmol kg^{-1} from the Atlantic to Indian to the Pacific Ocean, respectively. The model
631 also predicts some variation in shallow [Ba] that follows major surface-water currents, such as a
632 region of elevated [Ba] associated with the North Pacific Current, as well as low [Ba] in the western
633 North Atlantic associated with the Gulf Stream (Fig. 5B; Talley et al., 2011). However, these
634 features and the processes driving them await corroboration.

635 Considering the ocean as a whole, we can use our model to calculate the total Ba inventory of
636 seawater. Using the mean oceanic [Ba] of 89 nmol kg^{-1} and multiplying by the mass of seawater
637 (1.37×10^{21} kg) yields a total inventory of 122 ± 7 Tmol Ba, whereby the uncertainty is based on the
638 MAPE of model #3080 (6.0 %). This estimate of the total oceanic Ba inventory is between 11–21
639 % lower than existing estimates of 145 Tmol Ba (Dickens et al., 2003; Carter et al., 2020). Given
640 the range of probable global marine Ba fluxes between 18 (Paytan & Kastner, 1996) and 44 Gmol
641 Ba yr^{-1} (Rahman et al., 2022), our inventory estimate places the mean residence time of Ba in
642 seawater between 2,600–7,200 years.

643 5.3.2. *The Ba–Si relationship*

644 We now quantify spatial and vertical variations in the Ba–Si relationship, which we explore using
645 Ba*. Star tracers, such as Ba*, highlight the processes affecting the distribution of a tracer by
646 comparing it to another tracer that shares the same circulation (Gruber & Sarmiento, 1997). The
647 concept has since been extended to study the processes affecting the distributions of many other
648 bioactive elements, including Si (Si*, relative to N; Sarimento et al., 2004), cadmium (Cd*,
649 relative to P; Baars et al., 2014), zinc (Zn*, relative to Si; Wyatt et al., 2014). First defined by
650 Horner et al. (2015) for Ba, Ba* is analogous to other star tracers: it is a measure of Ba–Si
651 decoupling whereby larger values indicate larger Ba–Si deviations relative to expected mean ocean
652 behavior. Vertical or spatial differences in Ba and Si sources or sinks will drive variations in Ba*,
653 as will any Ba:Si fractionation occurring during their combined cycling. Conversely, if all Ba and
654 Si cycling occurs in the same places (and with a fixed Ba:Si ratio), no Ba–Si decoupling will occur
655 and Ba* will exhibit conservative behavior. Since Ba and Si are cycled by different processes *and*
656 there are large vertical and spatial variations in the intensity of these processes (e.g., Bishop, 1989),
657 significant variations in Ba* are possible. We now explore these variations.

658 In the surface ocean, patterns of Ba* generally resemble those of [Ba] (Fig. 4). In large parts of
659 the ocean, surface [Si] approaches 0 $\mu\text{mol kg}^{-1}$; thus, variations in Ba* derive mostly from
660 variations in [Ba]. This is most evident when examining regions with significant terrestrial input
661 of Ba, such as from major rivers (Sect. 5.2.1) and from rivers and continental shelves in the Arctic
662 (e.g., Guay & Falkner, 1998; Whitmore et al., 2022; Fig. 5A). The Southern Ocean also exhibits
663 positive Ba*, though we suspect the mechanism is different. Here we observe a belt of waters with
664 positive Ba* $\approx +20 \text{ nmol kg}^{-1}$ centered on the Polar Frontal Zone—the region between the Antarctic
665 Polar Front and the Subantarctic Front (Orsi et al., 1995; Fig. 5A). Silicic acid is intensely stripped
666 from waters that transit northward through this region (e.g., Sarmiento et al., 2004), potentially
667 contributing to elevated Ba* at the sea surface. Dissolved [Ba] and Ba* then decrease to the north
668 of the Subantarctic front, partly driven by extensive particulate Ba formation in the frontal region
669 (e.g., Bishop, 1989).

670 At 1,000 m, the Atlantic, South Pacific, and southern Indian Oceans exhibit positive Ba* around
671 $+10 \text{ nmol kg}^{-1}$, whereas the North Pacific, Southern, and northern Indian Oceans are negative
672 between -10 to -20 nmol kg^{-1} (Fig. 6C). The positive anomalies are likely related to the northward

673 spreading of southern-sourced intermediate waters that originate within the Polar Frontal Zone and
674 carry positive Ba* into the low latitudes (e.g., Bates et al., 2017). In the Atlantic, these values are
675 carried all the way to the north of the basin and return as North Atlantic Deep Water with only
676 minor modifications to Ba* ($\approx +10$ nmol kg⁻¹; Figs. 6C, 7C, 8C). Negative Ba* in the North Pacific,
677 Southern, and northern Indian Ocean at 1,000 m likely reflects a mixture of hydrographic processes
678 and *in situ* processes. For example, the extensive region of negative Ba* in the North Pacific is
679 closely associated with North Pacific Intermediate Water, which originates in the Sea of Okhotsk
680 (Talley, 1991). While the specific mechanism sustaining this particular Ba* feature is unknown, it
681 most possibly reflects a combination of preferential removal of Ba relative to Si in the source water
682 formation region (such as from particulate Ba formation) and weak vertical mixing in the
683 subsurface North Pacific relative to lateral transports (e.g., Kawabe & Fujio, 2010). We suspect
684 that the negative Ba* values seen above 1,000 m in the northern Indian Ocean originate through
685 processes occurring internally within this basin, as the majority of the Indian Ocean below 1,000
686 m exhibits positive Ba*. A possible mechanism for these shallow negative Ba* anomalies may
687 relate to the relatively weak overturning transports (Talley, 2008) and strong particulate Ba cycle
688 north of 30 °S (Singh et al., 2013), though this awaits more detailed investigation.

689 Lastly, the Southern Ocean exhibits negative Ba* between -10 and -20 nmol kg⁻¹ from ≈ 200 m
690 water depth to the seafloor. These negative anomalies in Ba* appear to be associated with
691 Circumpolar Deep Water and, below that, Antarctic Bottom Water; the influence of the latter can
692 also be seen in near-bottom negative Ba* in the South Pacific, southern Indian, and South Atlantic
693 Oceans (Fig. 8C). As with the other basins, the origin of the negative Ba* waters in the Southern
694 Ocean likely reflects a combination of *in situ* and circulation-related phenomena. For example, in
695 the Southern Ocean, Si is only stripped at the very surface, whereas particulate Ba formation is
696 thought to be greatest in the mesopelagic (i.e., between 200–1,000 m; e.g., Stroobants et al. 1991).
697 Barite formation is generally considered to be related to the regeneration of particulate organic
698 matter (e.g., Chow & Goldberg, 1960), whereby the former consumes Ba and the latter releases
699 Si. Thus, intense organic matter remineralization and associated pelagic BaSO₄ precipitation could
700 contribute to negative Ba* in the mesopelagic Southern Ocean. Similarly, the Si cycle in the
701 Southern Ocean tends to ‘trap’ a significant fraction of the global Si inventory in the waters
702 circulating close to Antarctica (e.g., Holzer et al., 2014). Since the calculation of Ba* depends on

703 both [Ba] and [Si], waters with elevated [Si] will exhibit lower Ba* whether or not there is elevated
704 Ba removal.

705 By 2,000 m, almost all of the ocean north of 50 °S exhibits positive Ba* (Fig. 7C). By 4,000 m,
706 the areal extent of the positive-Ba* waters shrinks to encompass the area north of 30 °S (Fig. 8C).
707 Despite covering a smaller area, the abyssal ocean exhibits the most positive Ba* values outside
708 of the surface of the Southern Ocean. The reasons for elevated and increasing Ba* between the
709 deep and abyssal oceans likely reflects a mixture of local and regional processes, and we offer two
710 speculative explanations for these patterns. First, Si trapping in the Southern Ocean potentially
711 renders most of the deep ocean away from Antarctica deficient in Si relative to Ba. Thus, much of
712 the ocean may exhibit more positive Ba* than the deep circum-Antarctic region due to processes
713 unrelated to Ba cycling. Second, the most positive Ba* values are generally found close to the
714 seafloor, rather than the mid-depths, especially in the North Pacific, the Peru and Chile Basins,
715 and the Philippine Sea. This may indicate a mechanism that preferentially removes Ba (relative to
716 Si) from the mid-depths, or input of Ba (relative to Si) close to the seafloor.

717 Systematic variations in Ba* arise due to differences in the marine biogeochemical cycles of Ba
718 and Si. While, in some cases, the specific drivers of these variations remains unresolved, our model
719 identifies multiple hotspots of Ba–Si decoupling that warrant additional study.

720 5.3.3. Barite saturation state of seawater

721 Here we show that our approach can predict Ω_{barite} with an MAE of 0.08, that our output is in
722 agreement with published values, and that the deep ocean, below 1,000 m, is at saturation with
723 respect to BaSO₄. By comparison to literature data, we estimate that our model achieves a typical
724 prediction uncertainty on Ω_{barite} of 0.08 (see Appendix). Accordingly, values of Ω_{barite} between
725 0.92–1.08 can be considered as ‘BaSO₄ saturated,’ whereas values of $\Omega_{\text{barite}} < 0.92$ or > 1.08 indicate
726 under- or super-saturation, respectively. Global patterns in Ω_{barite} derived using our model are
727 similar to those reported by Monnin et al. (1999) and Rushdi et al. (2000). Readers looking for
728 detailed basin-by-basin descriptions of Ω_{barite} are directed to those studies. Briefly our model shows
729 that, excepting the high latitudes, the surface ocean is undersaturated with respect to BaSO₄ (i.e.,
730 $\Omega_{\text{barite}} < 0.92$). The lowest values of Ω_{barite} in the open ocean are observed in the hot, salty cores of

731 the Subtropical Gyres (Ω_{barite} between 0.1–0.2; Fig. 5D). Conversely, the cold and fresh polar
732 regions exhibit supersaturation at the sea surface, though there are important differences between
733 the Southern and Arctic Oceans. The Southern Ocean exhibits BaSO_4 saturation to depths around
734 2,000 m, whereas the Arctic Ocean switches to undersaturated conditions below the halocline
735 (~250 m). At 1,000 m, most of the North Pacific achieves saturation (or slight supersaturation)
736 with respect to BaSO_4 (Fig. 6D) and at 2,000 m almost all of the ocean exhibits $\Omega_{\text{barite}} > 0.92$. The
737 main exceptions to this are the Atlantic Ocean, which is undersaturated at all depths, and the
738 southern Indian Ocean between 35–50 °S (Fig. 7D). The South Pacific and Indian Oceans return
739 to undersaturated conditions by 4,000 m, whereas parts of the North Pacific remain saturated to
740 the seafloor (Fig. 8D). From a global perspective, the oceans are slightly undersaturated with
741 respect to BaSO_4 : volume-weighted mean $\Omega_{\text{barite}} = 0.82$; however, the ocean between 1,000–5,500
742 m exhibits $\Omega_{\text{barite}} \geq 0.92$ (Fig. 9). This result implies that the deep ocean, as a whole, is close to
743 chemical equilibrium with respect to BaSO_4 .

744 5.3.4. Model applications

745 In the spirit of maximizing model utility, we suggest three possible uses for model #3080 outputs.
746 First, the outputs can be used for model intercomparison and intercalibration. For example, a
747 number of statistical models, such as Optimum Multiparameter Optimization, have been
748 successfully used to study Ba cycling in the North Atlantic (Le Roy et al., 2018; Rahman et al.,
749 2022), Southeast Pacific (Rahman et al., 2022), and Mediterranean Sea (Jullion et al., 2017). These
750 models can apportion the relative contributions of *in situ* biogeochemical cycling and conservative
751 mixing to observed [Ba]; however, accurate quantification of these processes requires *a priori*
752 knowledge of end-member water mass [Ba], which model #3080 can provide. Our model could
753 also be used to benchmark output from process-based models, such as Ocean Circulation Inverse
754 Models (e.g., John et al., 2020; Roshan & DeVries, 2021). Second, the output can be used for
755 interpolation purposes. Many groups investigated Ba partitioning into various types of marine
756 carbonates (see Sect. 1 for examples); however, these investigations are sometimes performed
757 without a co-located measurement of [Ba]. In these cases output from model #3080 could be used
758 to help calibrate specific substrates, such as deep-sea corals or benthic forams. This also avoids
759 the potential for circular reasoning whereby [Si] is used to estimate [Ba], which is then

760 reconstructed from the Ba:Ca ratio of carbonates to estimate [Si]. Third, the model output makes
761 testable predictions for regions of the ocean that have yet to be sampled by GEOTRACES-style
762 surveys. Several of these regions, such as the Southern Ocean, exhibit with sharp lateral and
763 vertical gradients in [Ba], Ba*, and Ω_{barite} . Such gradients should be considered prime targets for
764 future process-oriented studies of marine Ba cycling.

765 **6. Data availability**

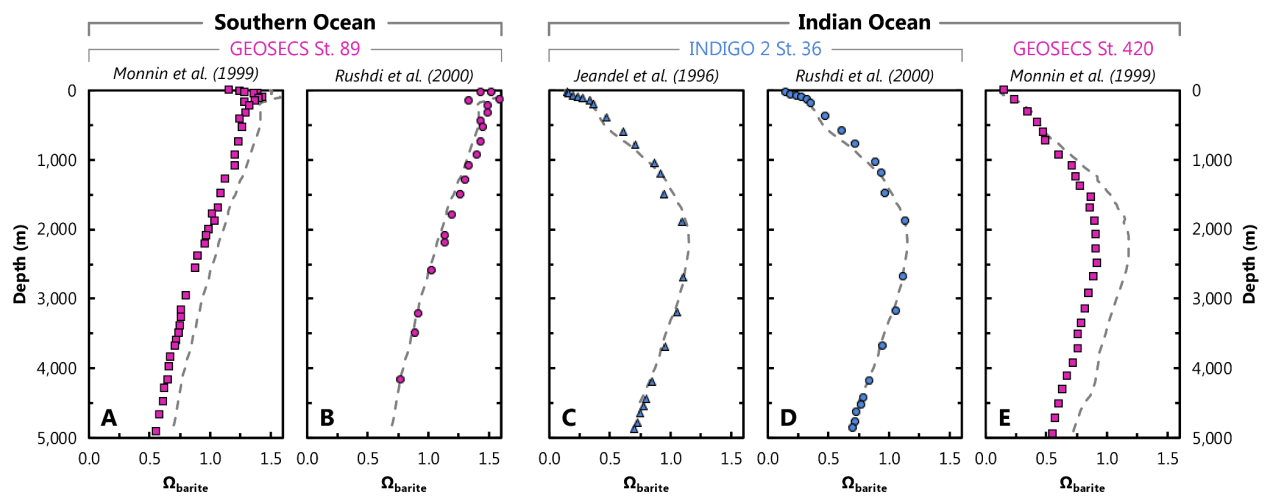
766 Data described in this manuscript can be accessed at the *Biological and Chemical Oceanography*
767 *Data Management Office* under data doi:10.26008/1912/bco-dmo.885506.1 (Horner & Mete,
768 2023).

769 **7. Conclusions**

770 This study presents a spatially and vertically resolved global model of [Ba] determined using
771 Gaussian Process Regression machine learning. The model reproduces several known features of
772 the marine [Ba] distribution and makes testable predictions in regions that are yet to be sampled.
773 Analysis of the model output reveals the mean oceanic [Ba] is 89 nmol kg⁻¹, implying a total
774 marine Ba inventory of 122±7 Tmol. Using predictors from the World Ocean Atlas, we also
775 estimate the global distribution of Ba* and Ω_{barite} . Both properties exhibit significant gradients that
776 can be systematically investigated in future studies. The mean oceanic Ω_{barite} is 0.82, though
777 between 1,000–5,500 m the mean is ≥ 0.92 , implying that the deep ocean is at equilibrium with
778 respect to BaSO₄. Our model output should prove valuable in studies of Ba biogeochemistry,
779 specifically for statistical- and process-based model validation, calibrating sedimentary archives,
780 and for identifying promising regions for further study. More broadly, our study demonstrates the
781 utility of using machine learning to accurately simulate the distributions of trace elements in
782 seawater. With minor adjustments, our approach could be employed to make predictions for other
783 dissolved tracers in the sea.

784 **Appendix**

785 Here we compare our results with published profiles of Ω_{barite} . Our results were calculated using
 786 the thermodynamic model of Rusdi et al. (2000), model #3080 [Ba], and WOA T , S , and pressure.
 787 Literature profiles of Ω_{barite} were calculated using one of three different thermodynamic models
 788 and *in situ* observations of [Ba], T , S , and pressure. In general, there is strong agreement between
 789 modeled and *in situ* Ω_{barite} whereby our model reproduces the shape of published profiles (Fig.
 790 A1). There are, however, some small systematic offsets between the various approaches, and we
 791 suspect that these derive from differences in the underlying thermodynamic models.



792

793 **Figure A1. Comparison of literature- (symbols) and Model #3080-derived (dashed line) values of**
 794 **Ω_{barite} .** Panels **A** and **B** show profiles of Ω_{barite} at GEOSECS St. 89 (60°0' S, 0°2' E). The other panels are
 795 from the Indian Ocean: **C** and **D** are from INDIGO 2 St. 36 (6°9' S, 50°55' E) and **E** from GEOSECS St.
 796 420 (0°3' S, 50°55' E), some ≈675 km north of INDIGO 2 St. 36.

797 We compare our model output with literature data Ω_{barite} at two locations in two basins (Fig. A1).
 798 These locations were chosen to ensure a fair comparison between studies; at each location, at least
 799 two studies calculated profiles of Ω_{barite} using the same underlying *in situ* data for [Ba], T , S , and
 800 pressure. Thus, any differences in modeled Ω_{barite} should derive from the thermodynamic model
 801 and not the input data. Likewise, literature profiles at these locations were based on calculations
 802 for pure, rather than strontian, BaSO_4 , as in our study. Published profiles of Ω_{barite} were extracted
 803 graphically from each study using *WebPlotDigitizer* (Rohatgi, 2022). This extraction process may

804 introduce some minor scatter in the literature data, though this is relatively minor compared to the
805 range of variation in Ω_{barite} .

806 First, we examine profiles of Ω_{barite} reported for GEOSECS St. 89 in the Southern Ocean (Fig. A1;
807 Monnin et al., 1999; Rushdi et al., 2000). Modeled and published profiles show supersaturation in
808 the surface ocean and undersaturation below 2,000–2,500 m. Profiles from Rushdi et al. (2000)
809 show excellent agreement with Ω_{barite} calculated from model #3080 [Ba] and WOA T , S , and
810 pressure, with our output offset by a MAE of 0.06 ($n = 22$). Given that we use the same
811 thermodynamic model as Rushdi et al. (2000), the overall excellent agreement with their study is
812 not surprising. However, the result is nonetheless reassuring since our study uses mean annual
813 values for the various inputs, whereas Rushdi et al. (2000) utilized *in situ* data. There is a slightly
814 larger offset between our profile of Ω_{barite} and that calculated by Monnin et al. (1999), with our
815 respective profile exhibiting an MAE of 0.13 ($n = 41$). This most likely reflects differences in the
816 underlying thermodynamic model and not the *in situ* data since our model reproduces the same
817 overall profile shape as Monnin et al. (1999). Likewise, both Monnin et al. (1999) and Rushdi et
818 al. (2000) used the same *in situ* input data and their results are highly comparable, albeit with an
819 offset similar to that between our results and Monnin et al. (1999).

820 Next we examine profiles of Ω_{barite} in the Indian Ocean for samples from INDIGO 2 St. 36 (Fig.
821 A1; Jeandel et al., 1996; Rushdi et al., 2000). Profiles of Ω_{barite} show undersaturation at the surface,
822 moderate supersaturation between 2,000–3,500 m, then return to undersaturated conditions down
823 to the seafloor. Our profile shows overall excellent agreement with that of Jeandel et al. (1996),
824 whereby a comparison of Ω_{barite} yields a MAE of of 0.03 ($n = 21$). Our profile shows similarly
825 good agreement with Rushdi et al. (2000), whereby a comparison between our respective values
826 of Ω_{barite} yields a MAE of 0.04 ($n = 20$).

827 We also compared our results with data from St. 420 of GEOSECS (Monnin et al., 1999), which
828 is located ≈ 675 km north of INDIGO 2 St. 36 (Fig. 2). As with data from the Southern Ocean
829 (GEOSECS St. 89), our profile data are offset to higher Ω_{barite} than those of Monnin et al. (1999),
830 with slightly larger MAE of 0.16 ($n = 29$). However, our modeled Ω_{barite} is generally in much closer
831 agreement with Monnin et al. (1999) above 1,100 m than below, equivalent to a MAE of 0.04 (n
832 = 8) and 0.21 ($n = 21$), respectively. In this case it is more challenging to ascribe a unique cause

833 of the differences in calculated Ω_{barite} ; these offsets could relate to differences in the predictors or
834 the thermodynamic model.

835 We can use these comparisons to estimate the prediction uncertainty on our model-derived values
836 of Ω_{barite} . The MAE of the 133 comparisons shown in Fig. A1 yields a value of 0.10. However,
837 there are different numbers of points in each profile; we thus believe it is more appropriate to
838 average the MAE calculated for each of the five profiles, which yields a value of 0.08. Both values
839 are similar to the 10 % prediction uncertainty reported by Monnin et al. (1999).

840 Overall, our ML-derived profiles of Ω_{barite} show excellent agreement with *in situ* data, both in
841 terms of profile shape and values of Ω_{barite} . We use this comparison to estimate the prediction
842 uncertainty on ML-derived values of Ω_{barite} , which we calculate as being between 0.08 and 0.10.
843 Should a revised thermodynamic model and/or improved BaSO_4 solubility coefficients become
844 available, a new grid of Ω_{barite} could be calculated using Model #3080 [Ba] and WOA T , S , and
845 pressure data.

846 **Author contributions**

847 Project conceptualization and funding acquisition by T.J.H. Data curation, formal analysis,
848 investigation, and methodology by O.Z.M., A.V.S., H.H.K., and T.J.H. Data visualization by
849 A.V.S. and T.J.H. Software provided by O.Z.M., A.V.S., H.H.K., and A.G.D. Writing (original
850 draft) by O.Z.M. and T.J.H.; review and editing by A.V.S., H.H.K., A.G.D., L.M.W., A.M.S.,
851 M.G., and W.D.L.

852 **Competing interests**

853 The authors declare that they have no conflict of interest.

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