



Supplement of
State of Wildfires 2023–2024

Matthew W. Jones et al.

Correspondence to: Matthew W. Jones (matthew.w.jones@uea.ac.uk), Douglas I. Kelley (doukel@ceh.ac.uk), Chantelle A. Burton (chantelle.burton@metoffice.gov.uk), and Francesca Di Giuseppe (francesca.digiuseppe@ecmwf.int)

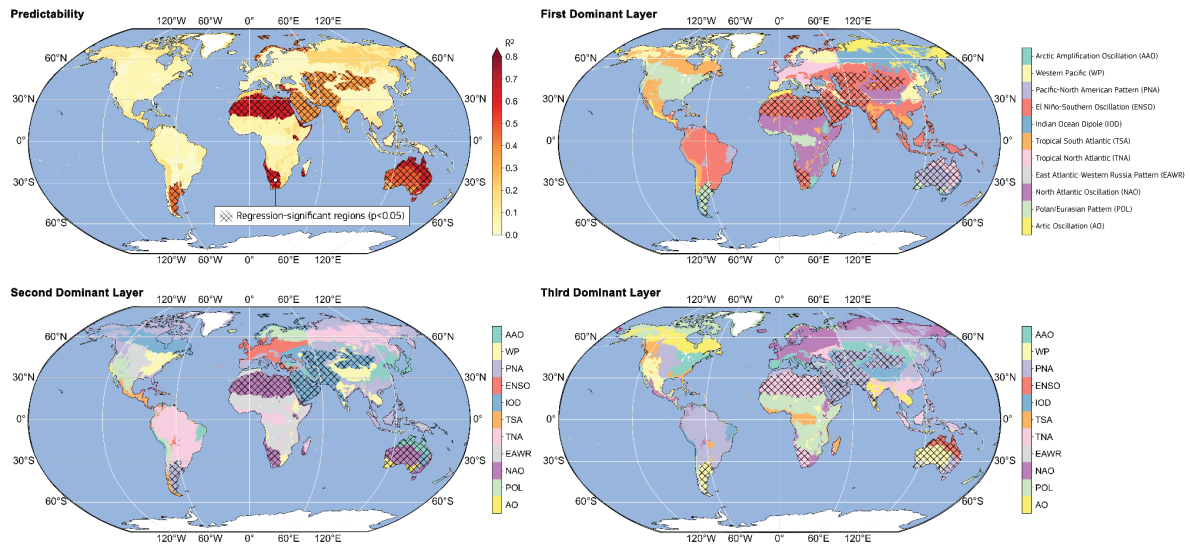
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1 **Supplementary Figures**

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3 Figures referred to in the main text.

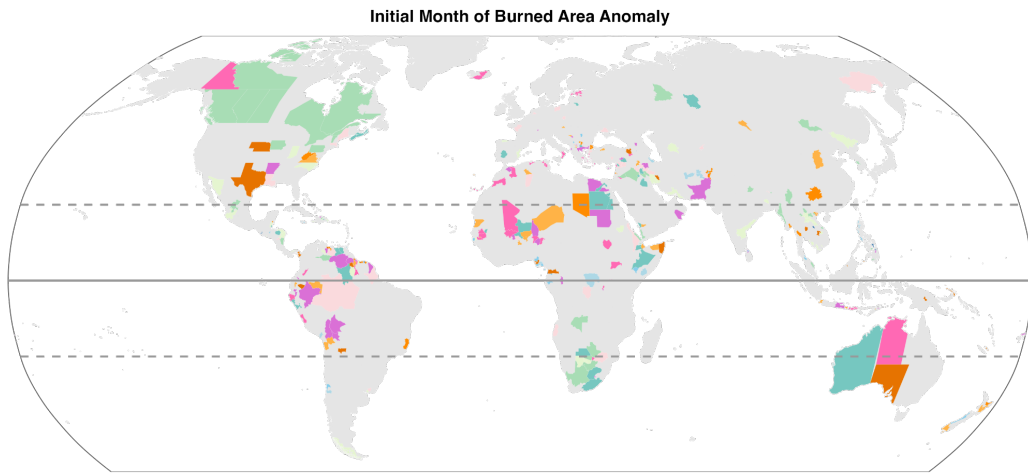
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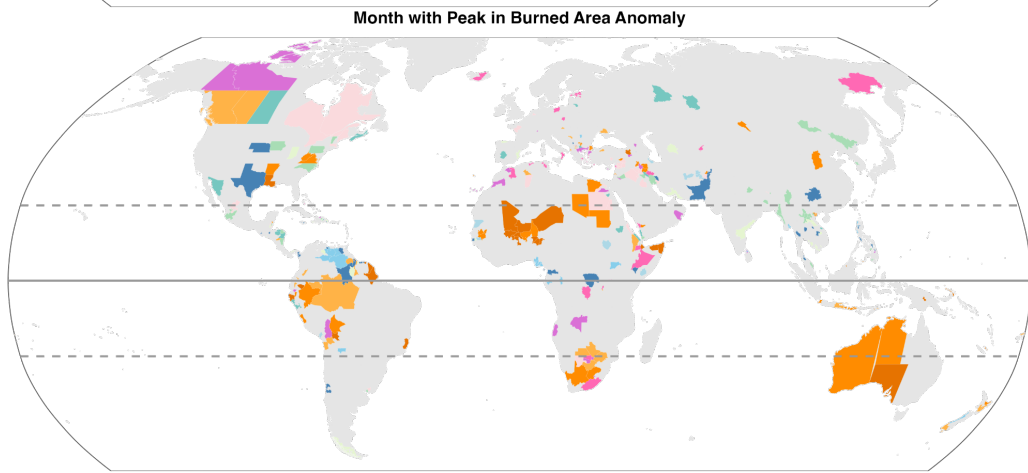
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6 **Figure S1:** Influence of modes of climate variability on Burned Area (BA) in the period 2001-
7 2020. The figure maps the mode of climate variability with a dominant (and secondary, and
8 tertiary) influence on interannual variability in BA. The coefficient of covariation (R^2) value
9 linking BA to each mode is also shown. The modes included are: Antarctic Oscillation (AAO);
10 West Pacific (WP) pattern; Pacific-North American (PNA) pattern; El Niño-Southern
11 Oscillation (ENSO); Indian Ocean Dipole (IOD); tropical South Atlantic (TSA) pattern; tropical
12 North Atlantic (TNA) pattern; East Atlantic/Western Russia (EAWR) pattern; North Atlantic
13 Oscillation (NAO); Polar-Eurasian (POL) pattern; and the Arctic Oscillation (AO).
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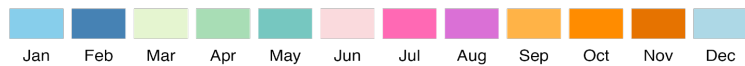
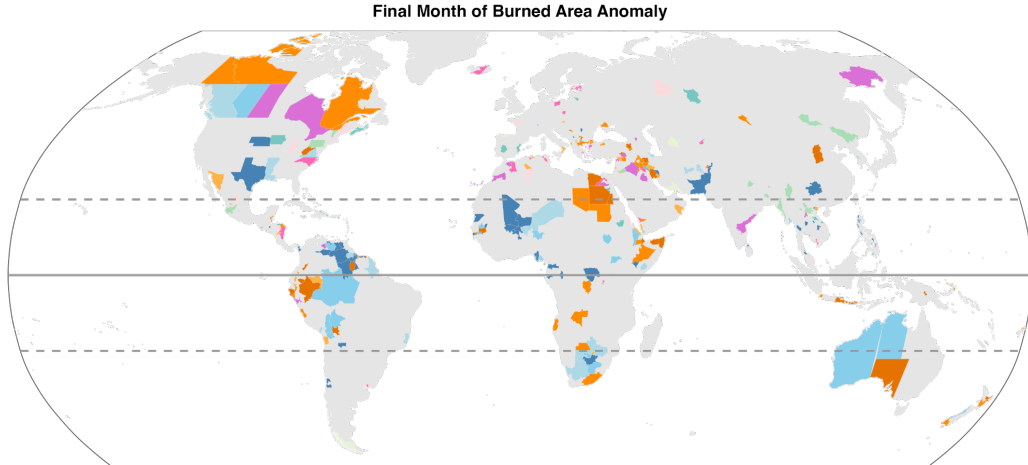
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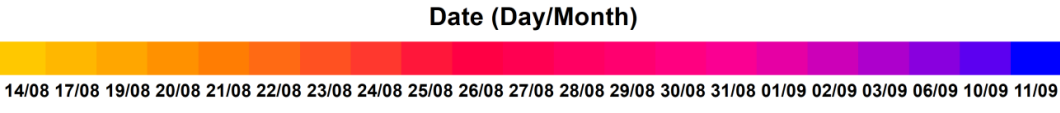
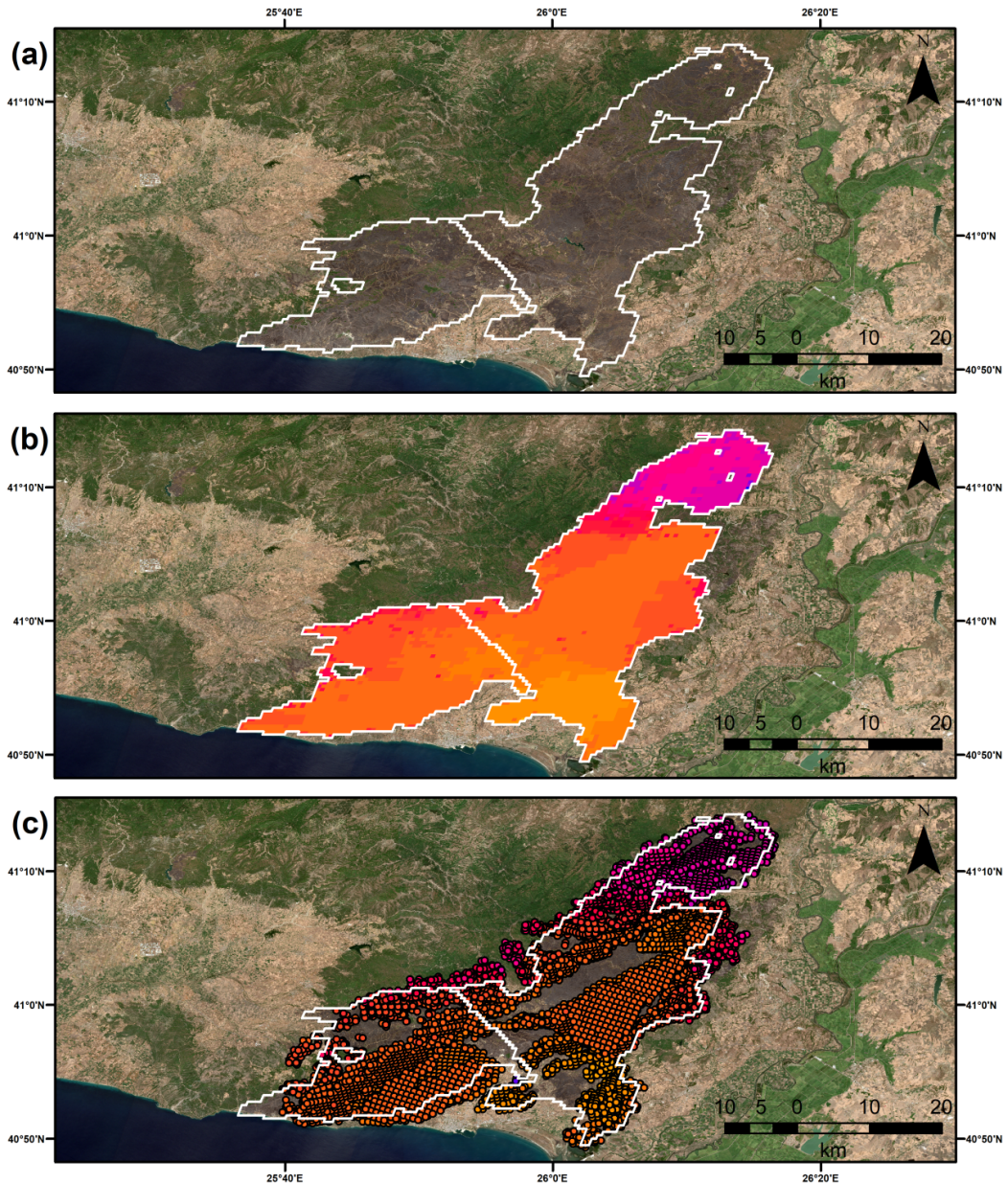


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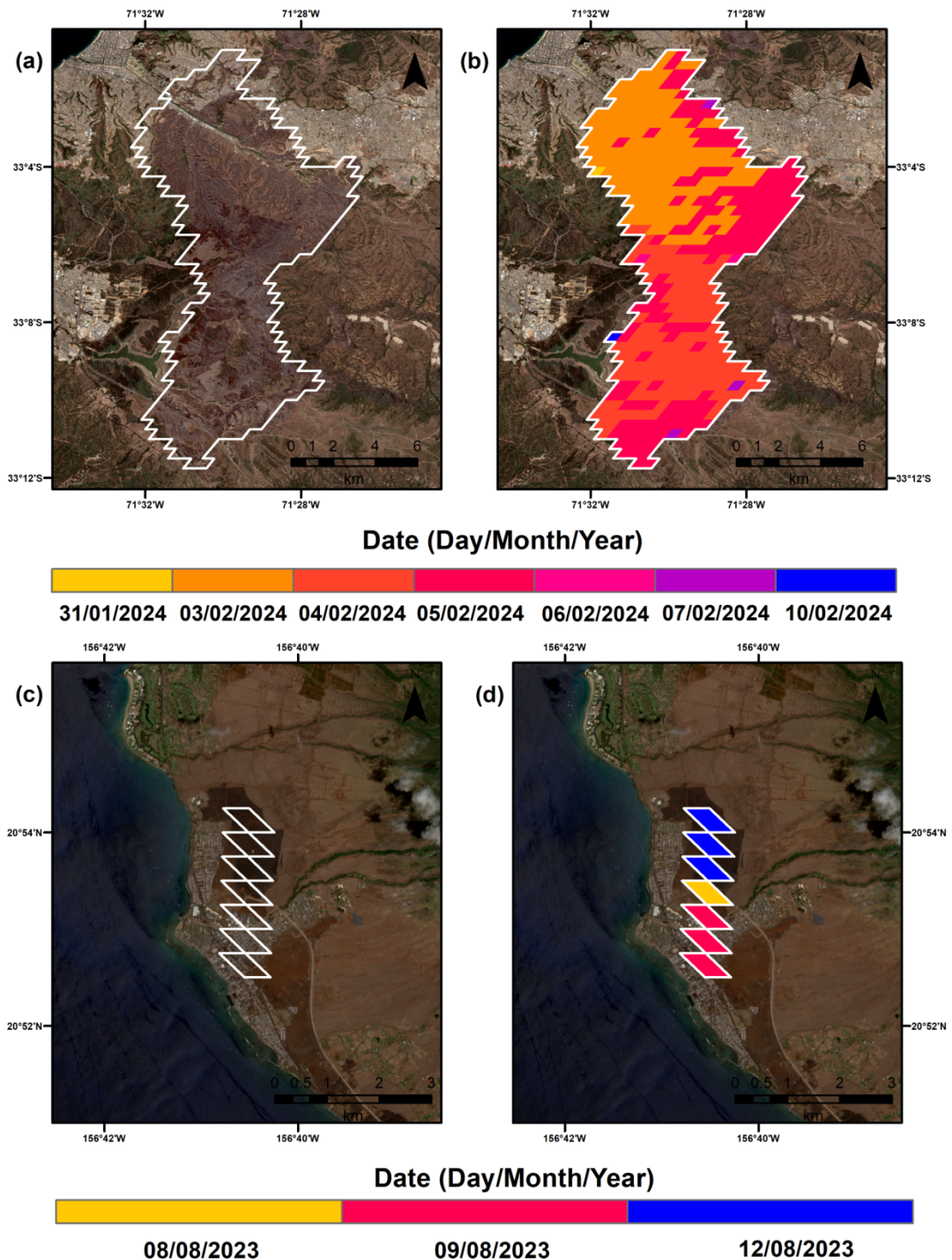
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Figure S2: (Top panel) first month, **(middle panel)** peak month, and **(lower panel)** final month of positive BA anomalies at Global Administrative Level 1 during March 2023-February 2024. Peak anomalies are identified relative to the monthly climatology in 2002-2023. The first and final months of the BA anomaly incorporate the period when BA was continuously above the climatological mean. Graduated colours are separated seasonally.



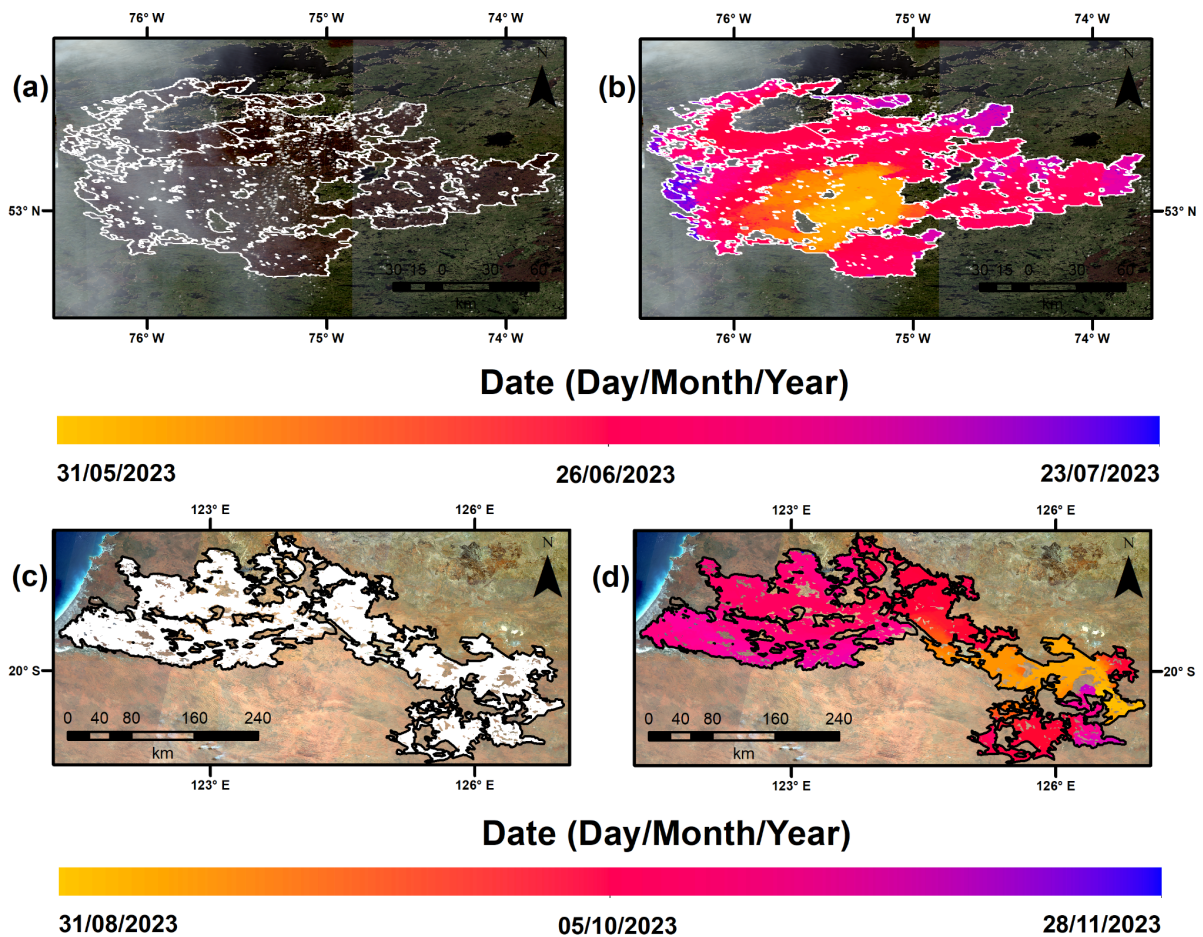
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Figure S3: Perimeter and daily progression of the largest fire ever recorded in the EU (Xanthopoulos et al., 2024; EU Science Hub, 2023), near Alexandroupolis in Macedonia and Thrace, Greece. Panel (a) shows a Sentinel-2 true colour composite image (10 m resolution) from 12th September 2023, the day after the fire ceased to grow. The darker colour of recently-burned surfaces contrasts with green unburned forests in surrounding areas. Overlaying the image are lines marking the perimeter of the Alexandroupolis fire from the Global Fire Atlas. Panel (b) additionally shows the burn date according to the MODIS BA dataset MCD64A1 (500 m resolution), and for comparison panel (c) shows the burn date from active fire detections from the Visible Infrared Imaging Radiometer Suite (VIIRS) sensor (375 m resolution; Schroeder et al., 2014).

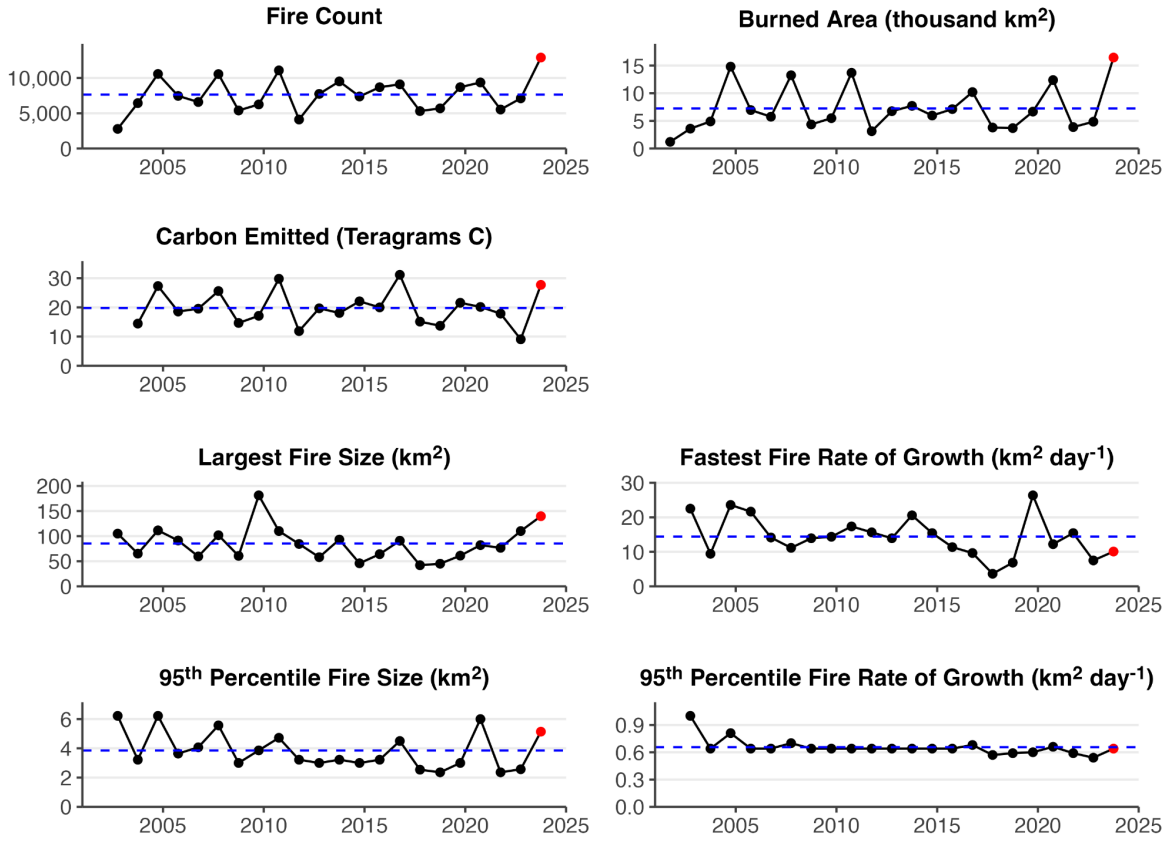


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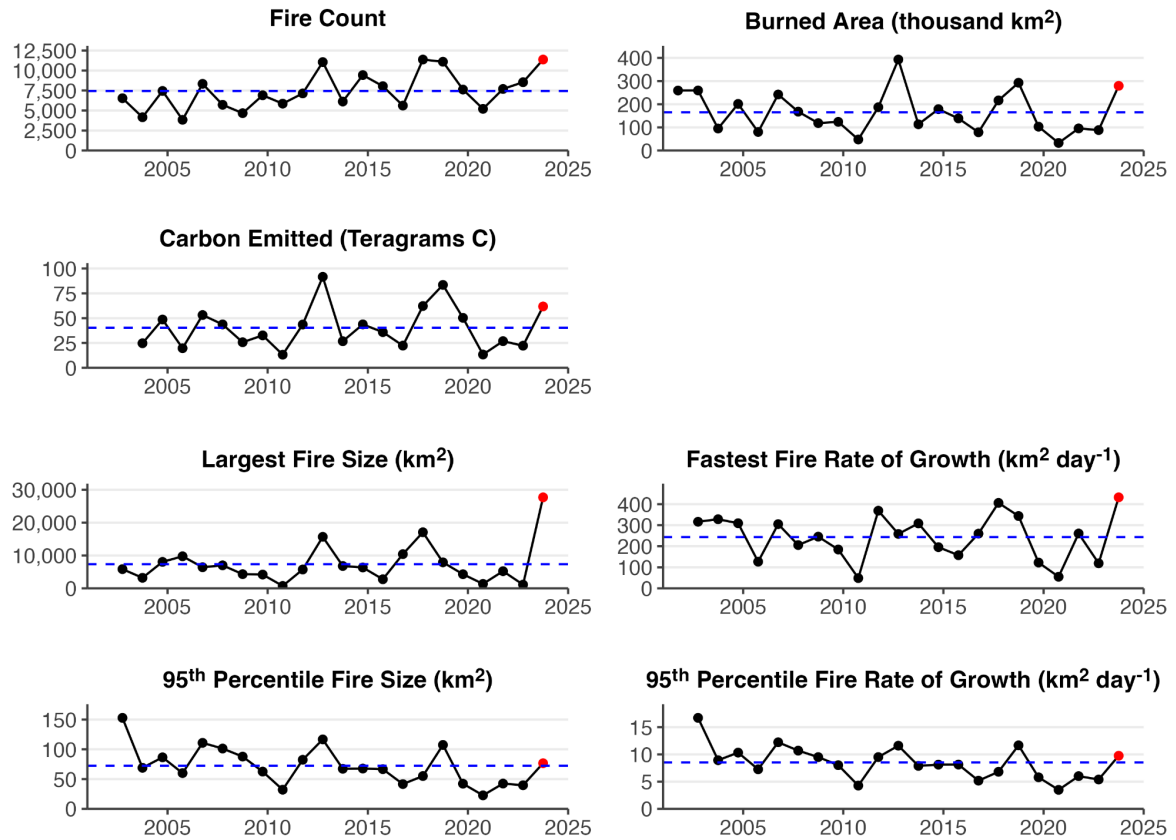
Figure S4: Perimeter and daily progression of extreme individual fires in (a-b) Valparaíso, Chile, and (c-d) Lahaina, Hawai'i. Panels (a) and (c) show Sentinel-2 true colour composite images (10 m resolution) from 8th March 2024 and 18th August 2023, on the first cloud-free day after each fire. Overlaying the image are lines marking the perimeter of the impactful fire events from the Global Fire Atlas. Panels (b) and (d) additionally show the burn date according to the MODIS BA dataset MCD64A1 (500 m resolution).



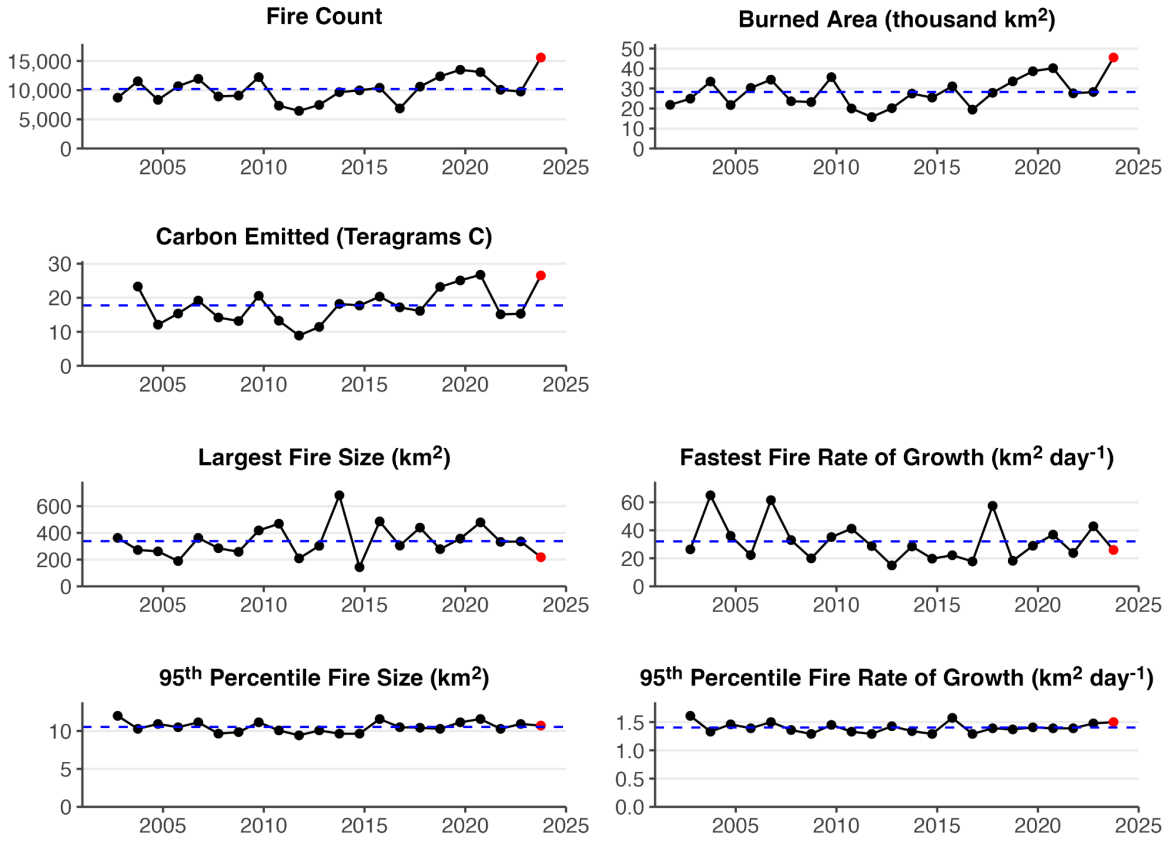
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 47 **Figure S5:** Perimeter and daily progression of extreme individual fires **(a-b)**
 48 near La Grande Reservoir in Quebec, Canada, and **(c-d)** in the Great Sandy Desert and Anna Plains,
 49 Australia. Panels **(a)** and **(c)** show Sentinel-2 true colour composite images (10 m resolution)
 50 based on observations in the periods 25/04/2023-25/08/2023 and 02/09/2023 to 08/09/2023,
 51 respectively. Overlaying the image **(a)** are white lines marking the perimeter of the La Grande
 52 fire according to the Global Fire Atlas. Overlaying the image **(c)** are white areas marking the
 53 area burned by the La Grande fire according to the Global Fire Atlas, and black lines marking
 54 the wildfire perimeter from the Department of Fire and Emergency Services in Western
 55 Australia. Panels **(b)** and **(d)** additionally show the burn date according to the MODIS BA
 56 dataset MCD64A1 (500 m resolution).
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 59 **Figure S6:** Summary of the 2023-2024 fire season in Lao PDR. Time series of annual fire
 60 count, BA, C emissions, PM2.5 emissions, 95th percentile fire size, fastest daily rate of growth,
 61 and 95th percentile fire daily rate of growth. Black dots show annual values prior to the latest
 62 fire season, red dots the values during the latest fire season, and blue dashed lines the
 63 average values across all fire seasons.

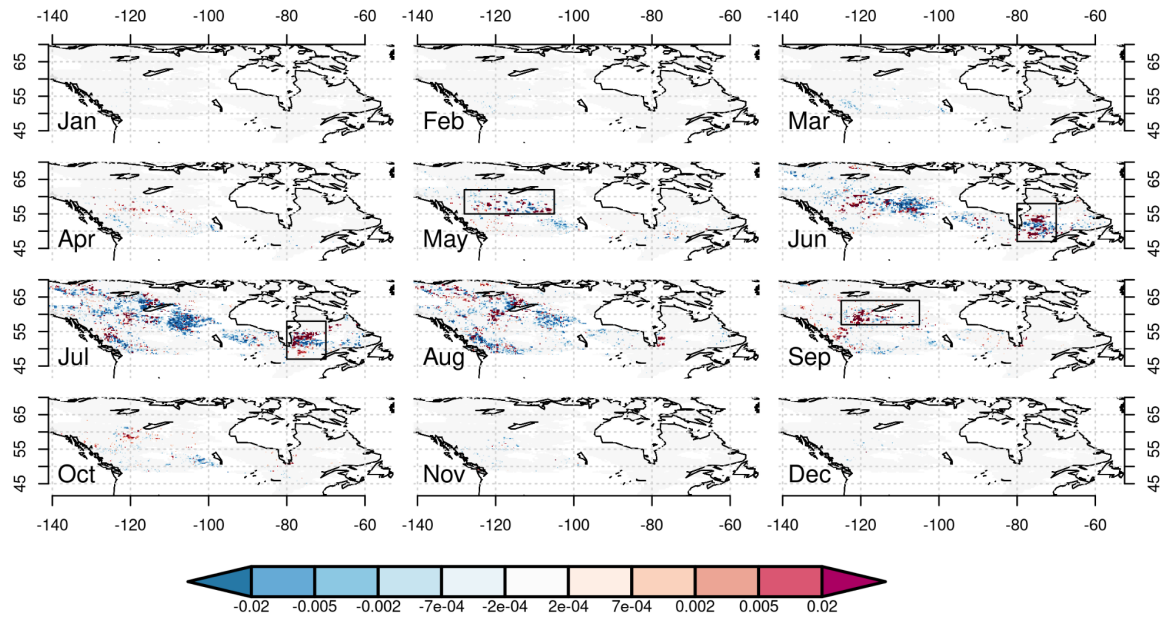


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 65 **Figure S7:** Summary of the 2023-2024 fire season in the state of Western Australia. Time
 66 series of annual fire count, BA, C emissions, PM2.5 emissions, 95th percentile fire size, fastest
 67 daily rate of growth, and 95th percentile fire daily rate of growth. Black dots show annual values
 68 prior to the latest fire season, red dots the values during the latest fire season, and blue dashed
 69 lines the average values across all fire seasons.



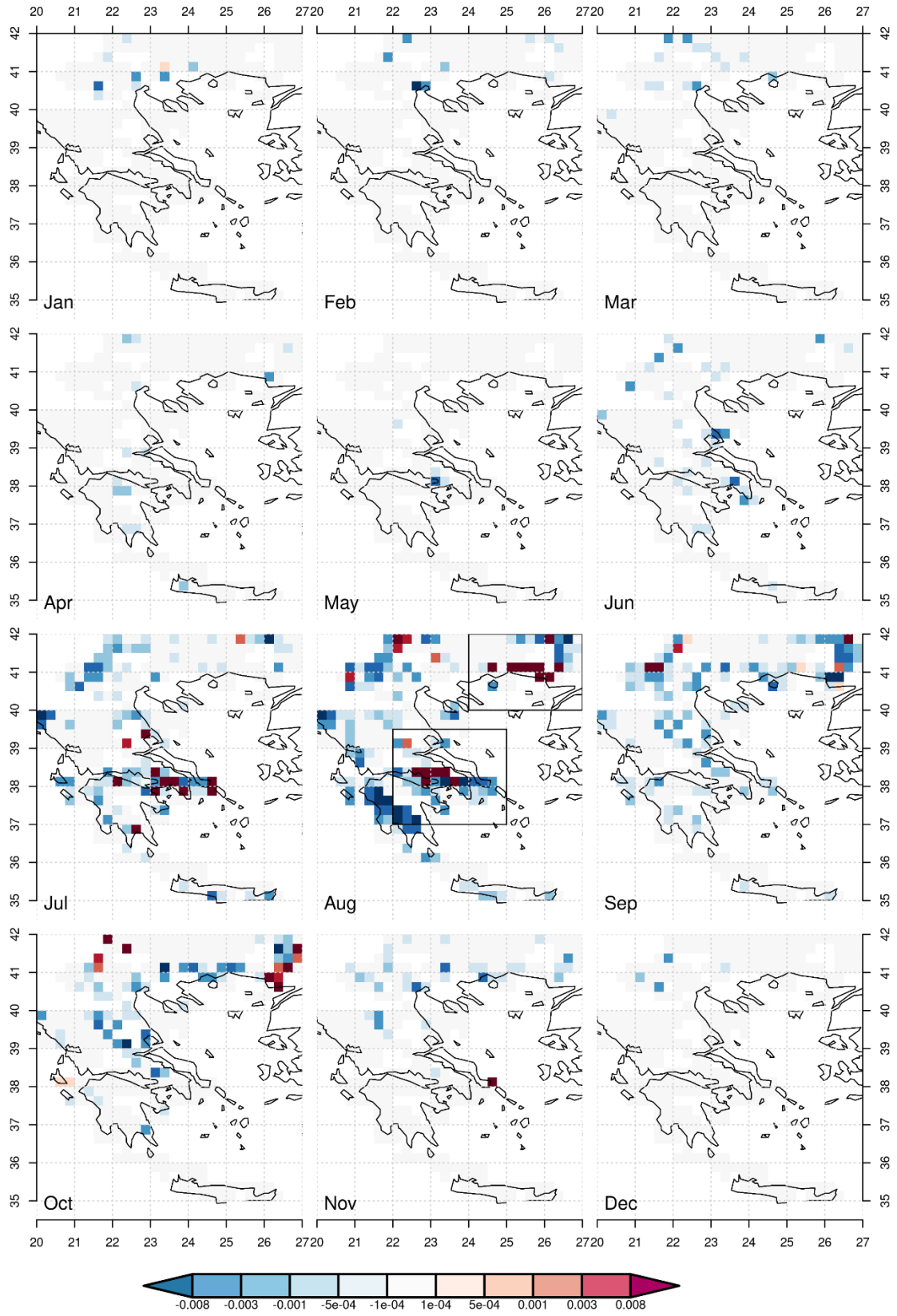
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Figure S8: Summary of the 2023-2024 fire season in Venezuela. Time series of annual fire count, BA, C emissions, PM2.5 emissions, 95th percentile fire size, fastest daily rate of growth, and 95th percentile fire daily rate of growth. Black dots show annual values prior to the latest fire season, red dots the values during the latest fire season, and blue dashed lines the average values across all fire seasons.

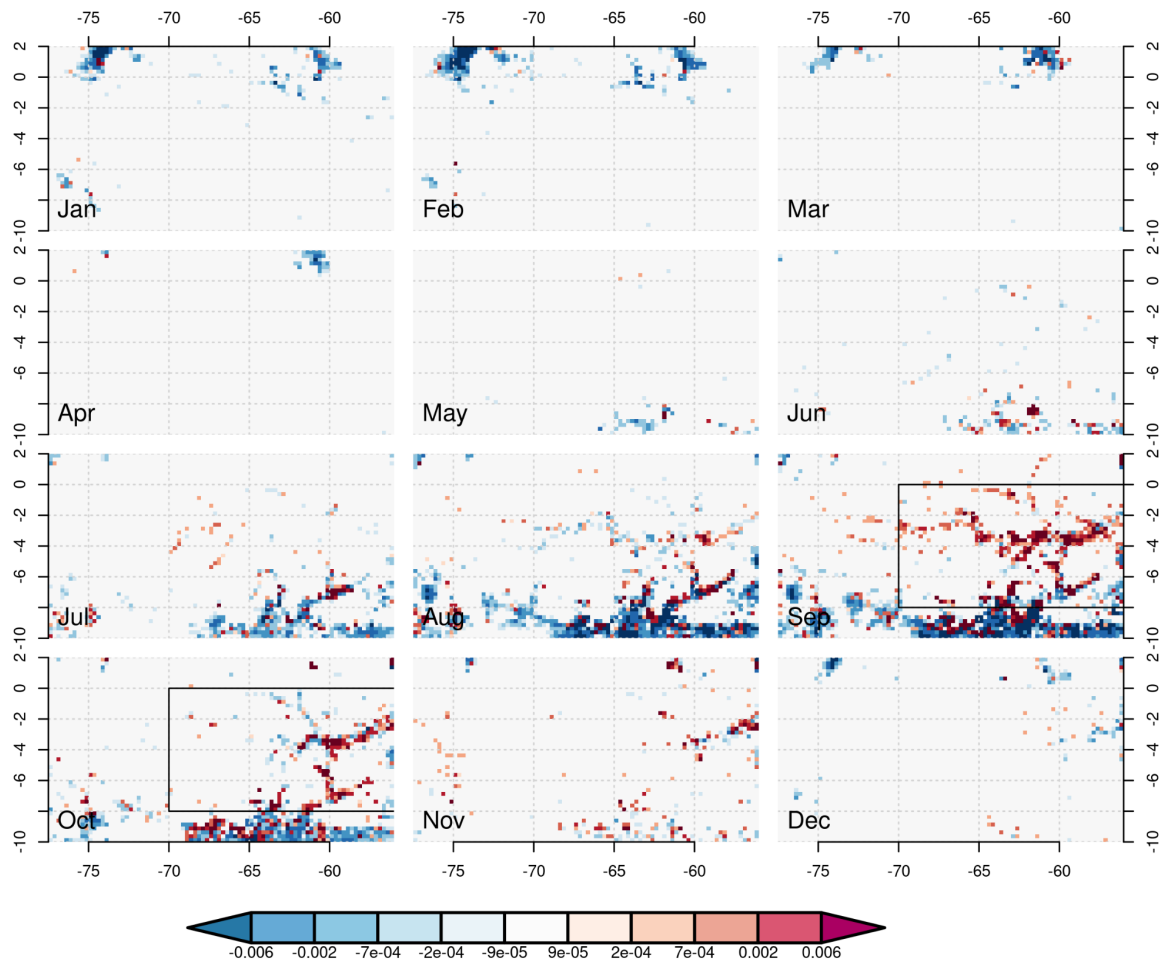


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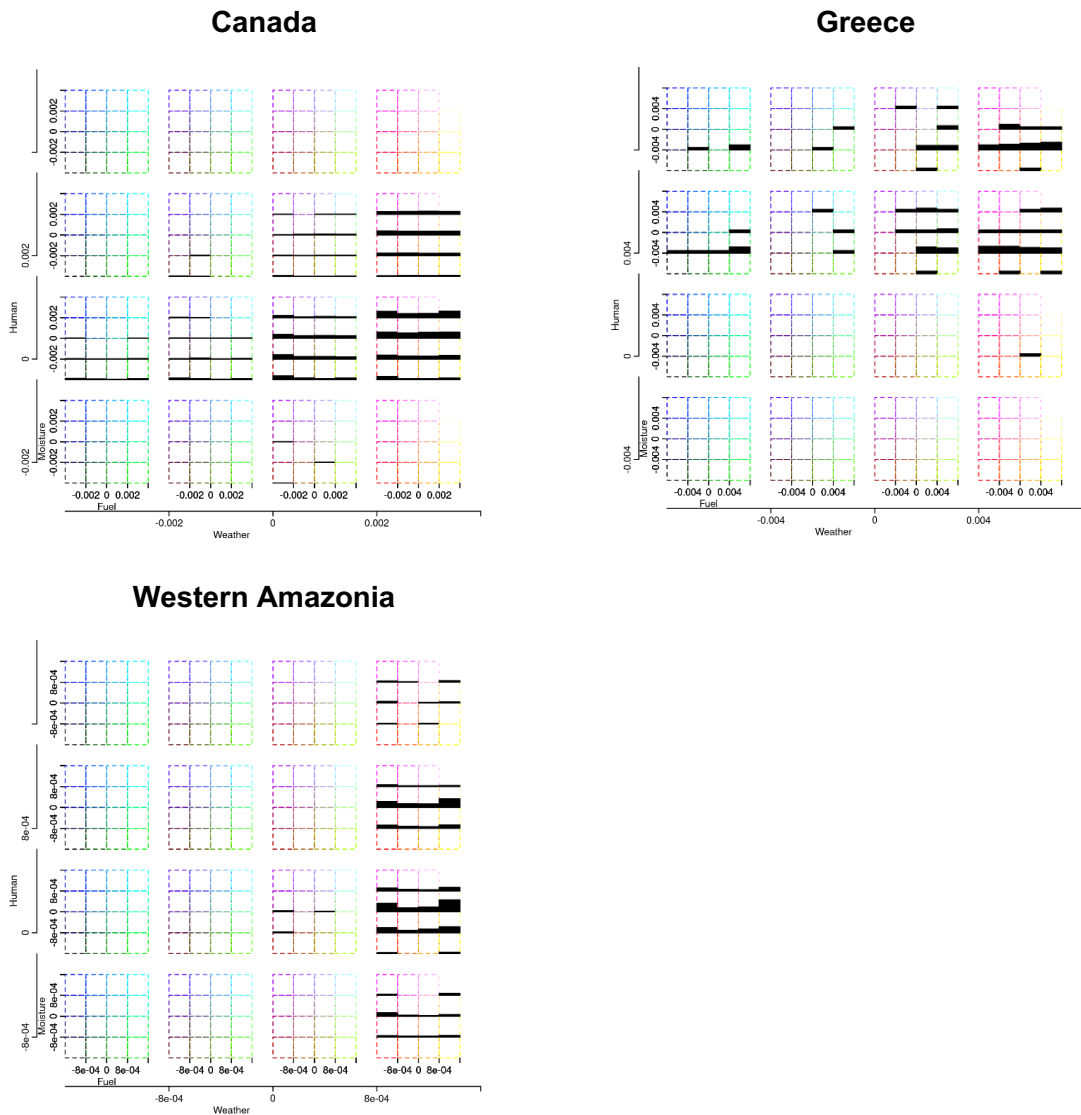
Figure S9: Monthly BA fraction anomaly at 0.25° for Canada for 2023 compare 2002-2023 climatological average. Boxes indicate focal months and regions in driver attribution (Section 3).



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83 **Figure S10: Same as Figure S9 for Greece**

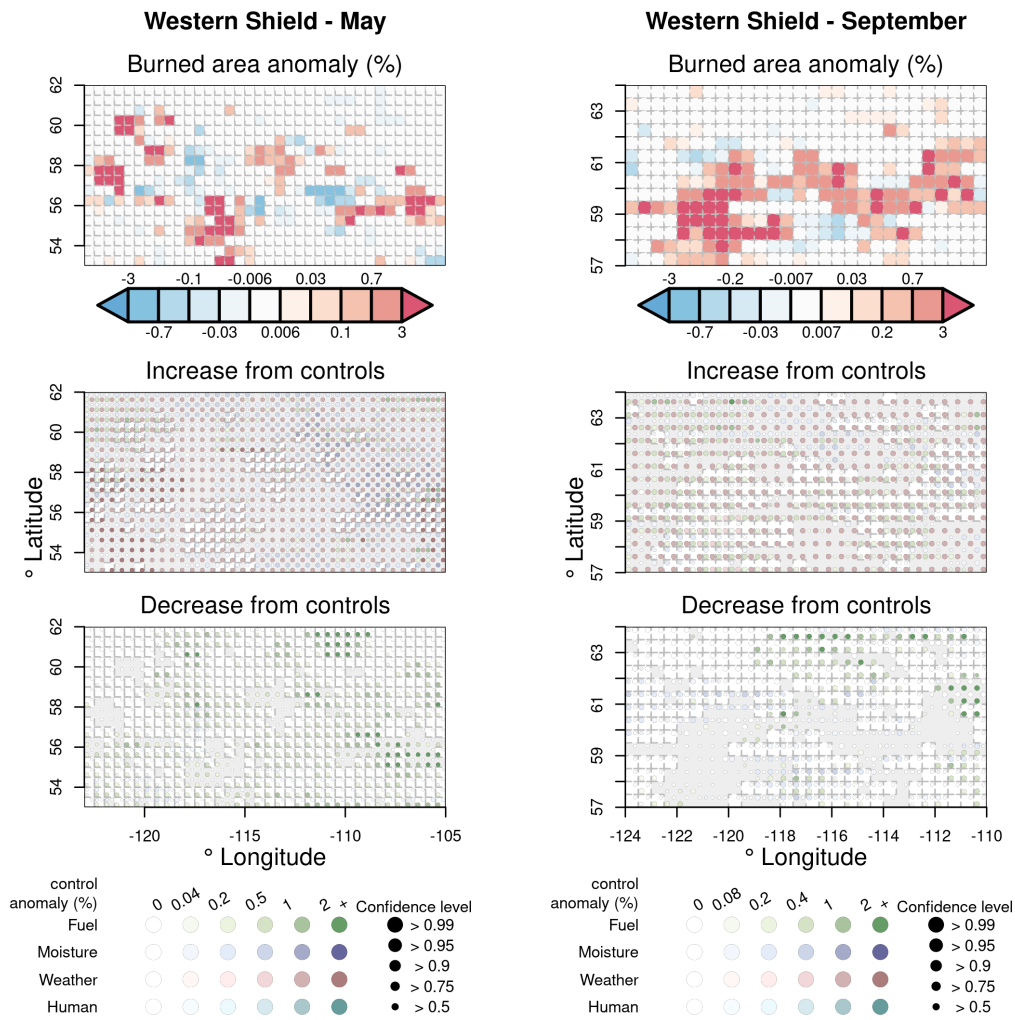


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 85 **Figure S11:** Same as **Figure S9** for Western Amazonia
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Figure S12: shows the co-occurrence of anomalies for 2023 of our four controls in different regions. In each box of 16 bins, the bottom left corresponds to the negative influence of fuel and moisture on fire anomalies, the top is the positive influence of fuel moisture, and the right indicates a positive influence of fuel load. The bottom left box indicates the negative influence of fire weather and humans, while the right boxes indicate the positive influence of fire weather, and the top indicates the positive influence of humans. The shading of each bin for each region indicates how much of that region falls into that bin. The shades themselves represent the uncertainty range, with grey indicating the 10th percentile and black indicating the 90th percentile.

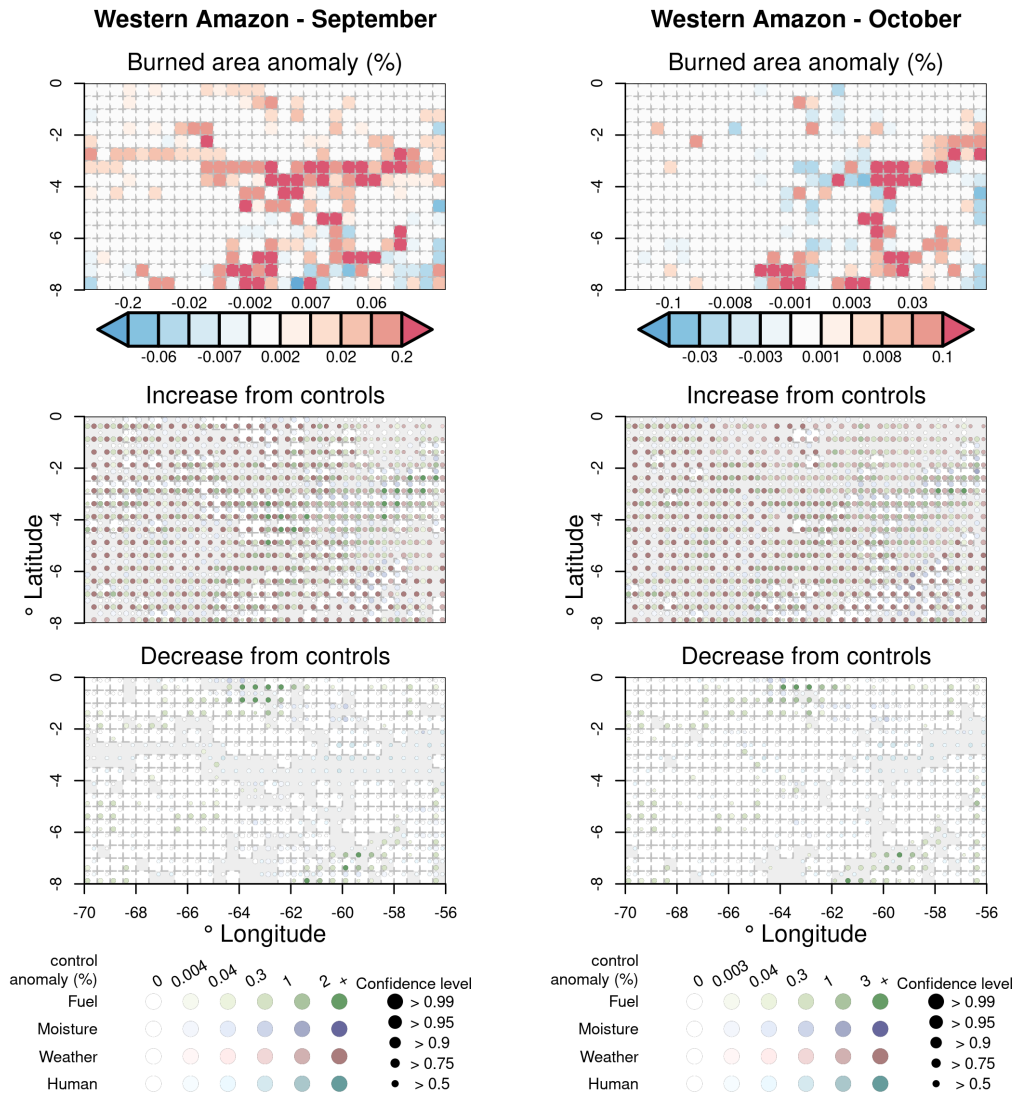


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105 **Figure S13:** Same as **Figure 11** but for the Canadian Western Shield.

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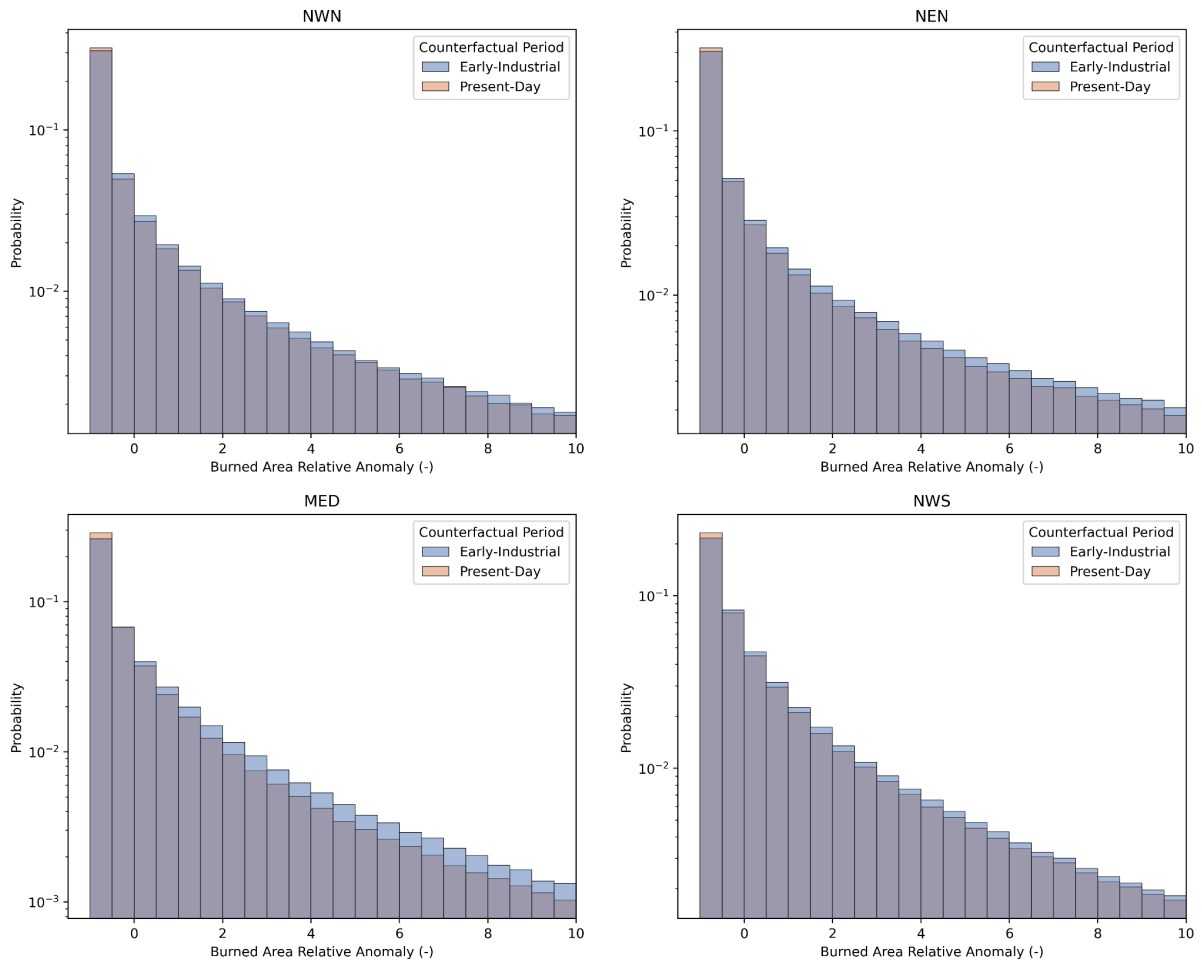
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109 **Figure S14:** Same as **Figure 11** but for Western Amazonia.

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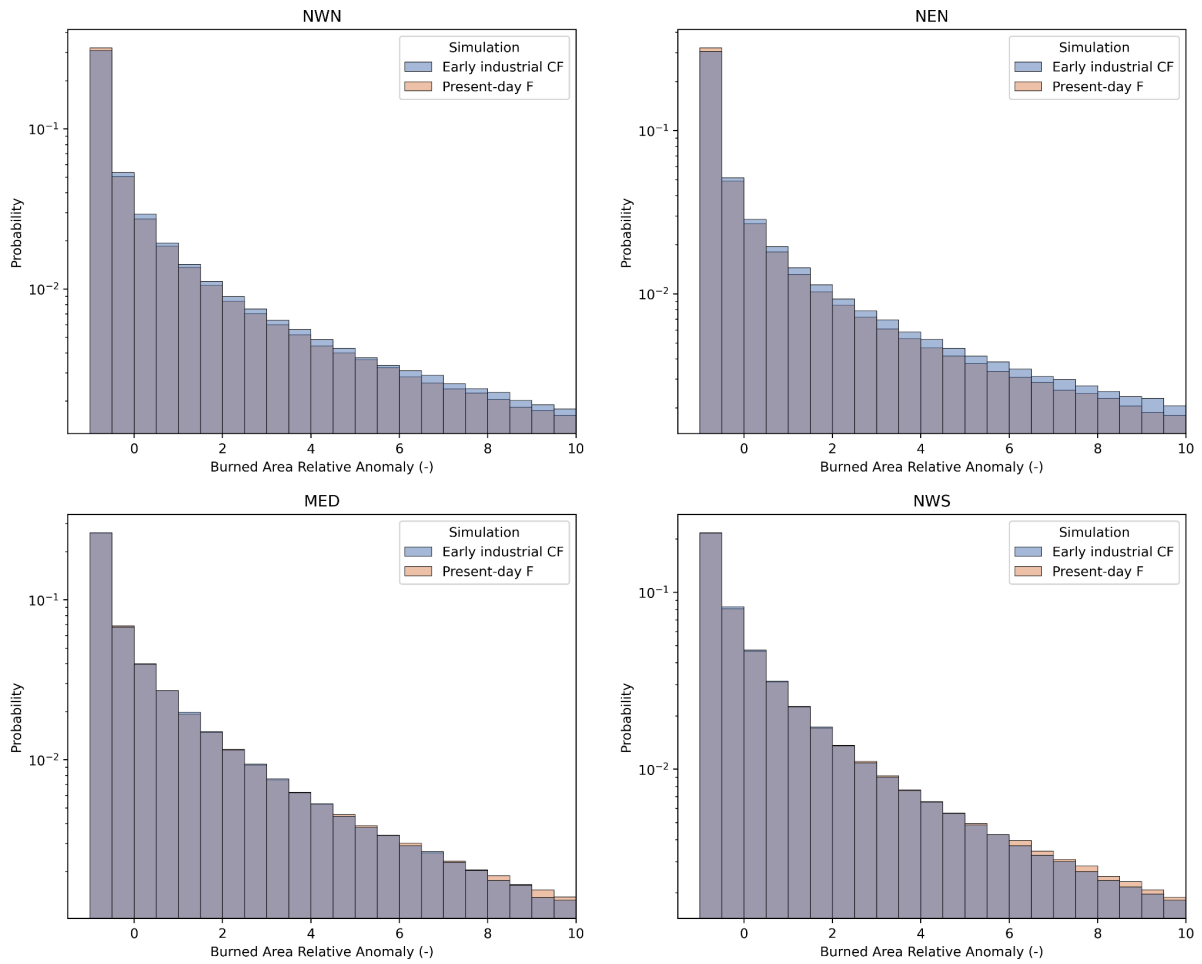
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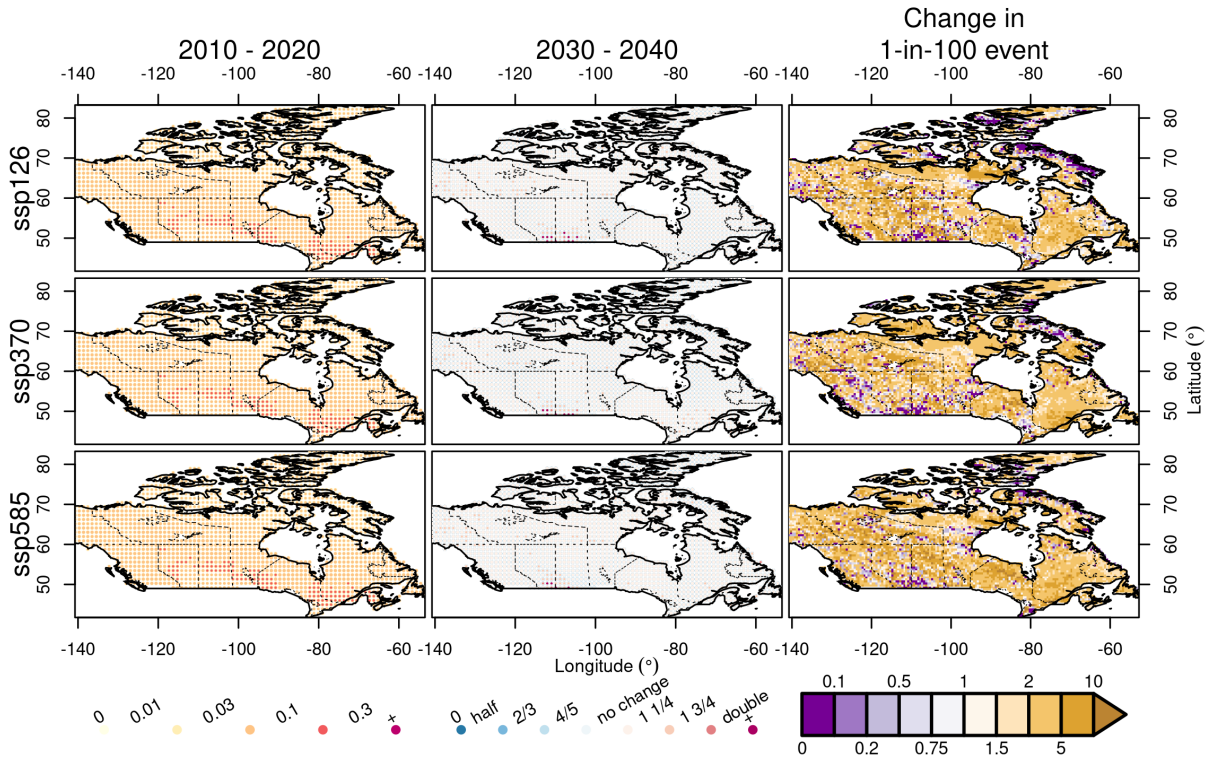
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Figure S15: Change in median BA anomaly due to socioeconomic factors (population and land-use change) from FireMIP. Present day BA (2003-2019) for counterfactual (detrended climate, orange) compared to early-industrial (1901-1917) in the counterfactual (detrended climate, blue), for AR6 regions. Top row: North West North America (NWN, LEFT) and North East North America NEN (RIGHT). Bottom row: Mediterranean (MED, LEFT), and North West South America (NWS, RIGHT). Probability is shown on a log scale.



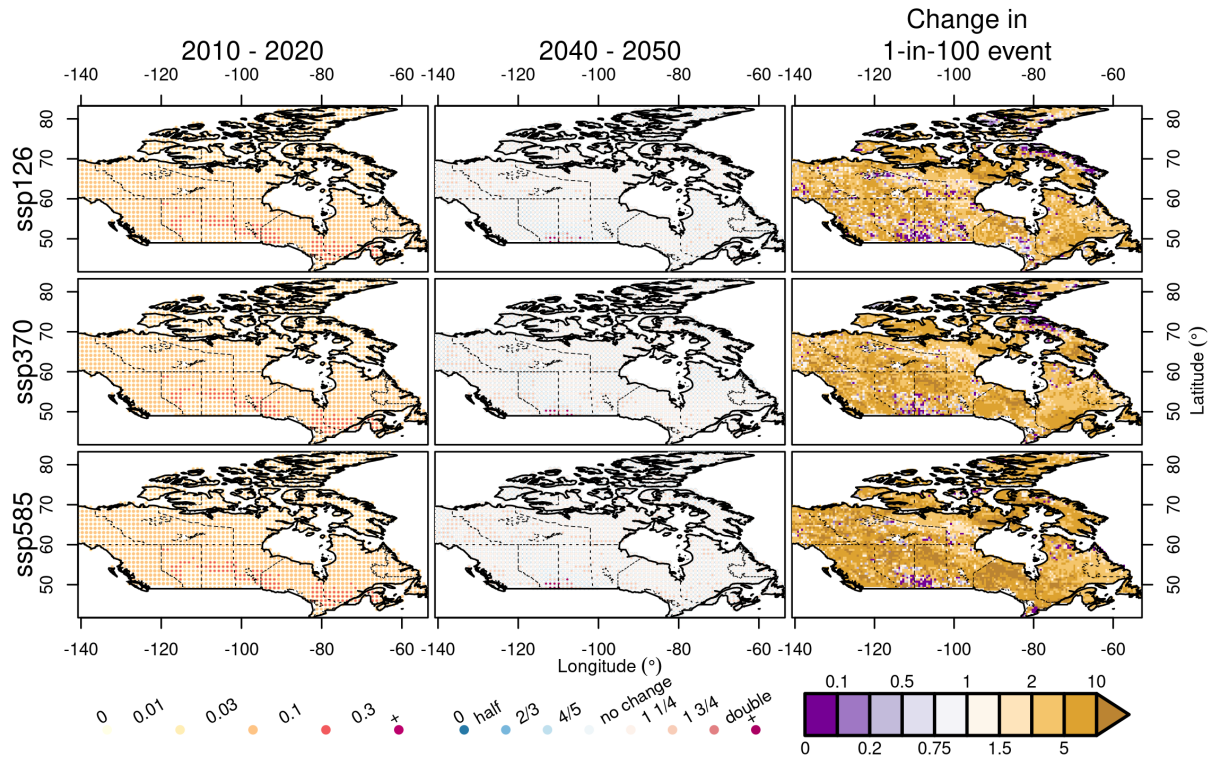
121
 122 **Figure S16:** Change in median BA anomaly due to all forcing (climate change and
 123 socioeconomic factors) from FireMIP. Present day BA (2003-2019) for factual (historical
 124 forcing, orange) compared to early-industrial (1901-1917) in the counterfactual (detrended
 125 climate, blue), for AR6 regions. Top row: North West North America (NWN, LEFT) and North
 126 East North America NEN (RIGHT). Bottom row: Mediterranean (MED, LEFT), and North West
 127 South America (NWS, RIGHT). Probability is shown on a log scale.
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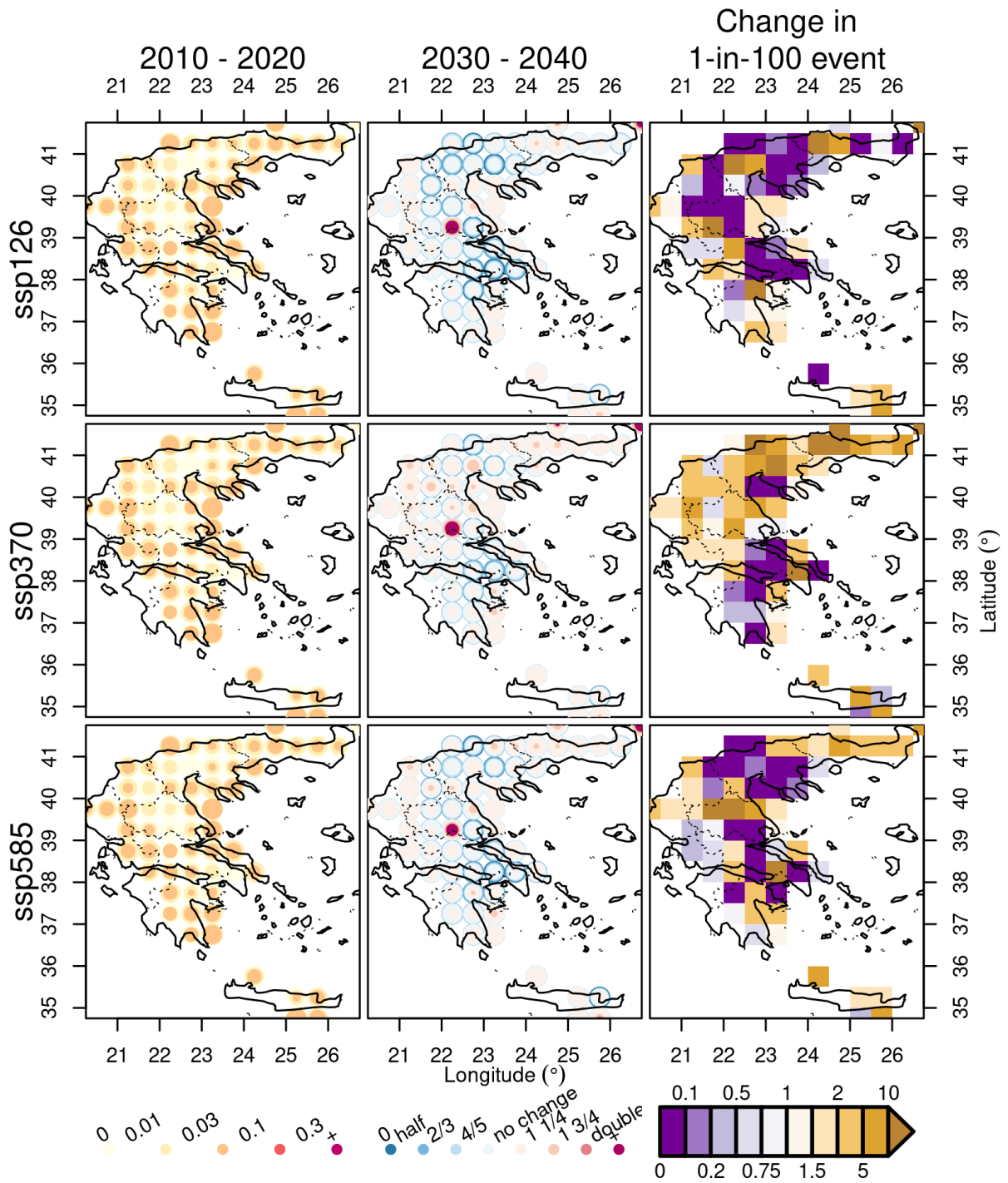
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Figure S17: Same as **Figure 18** but covering 2030-2040



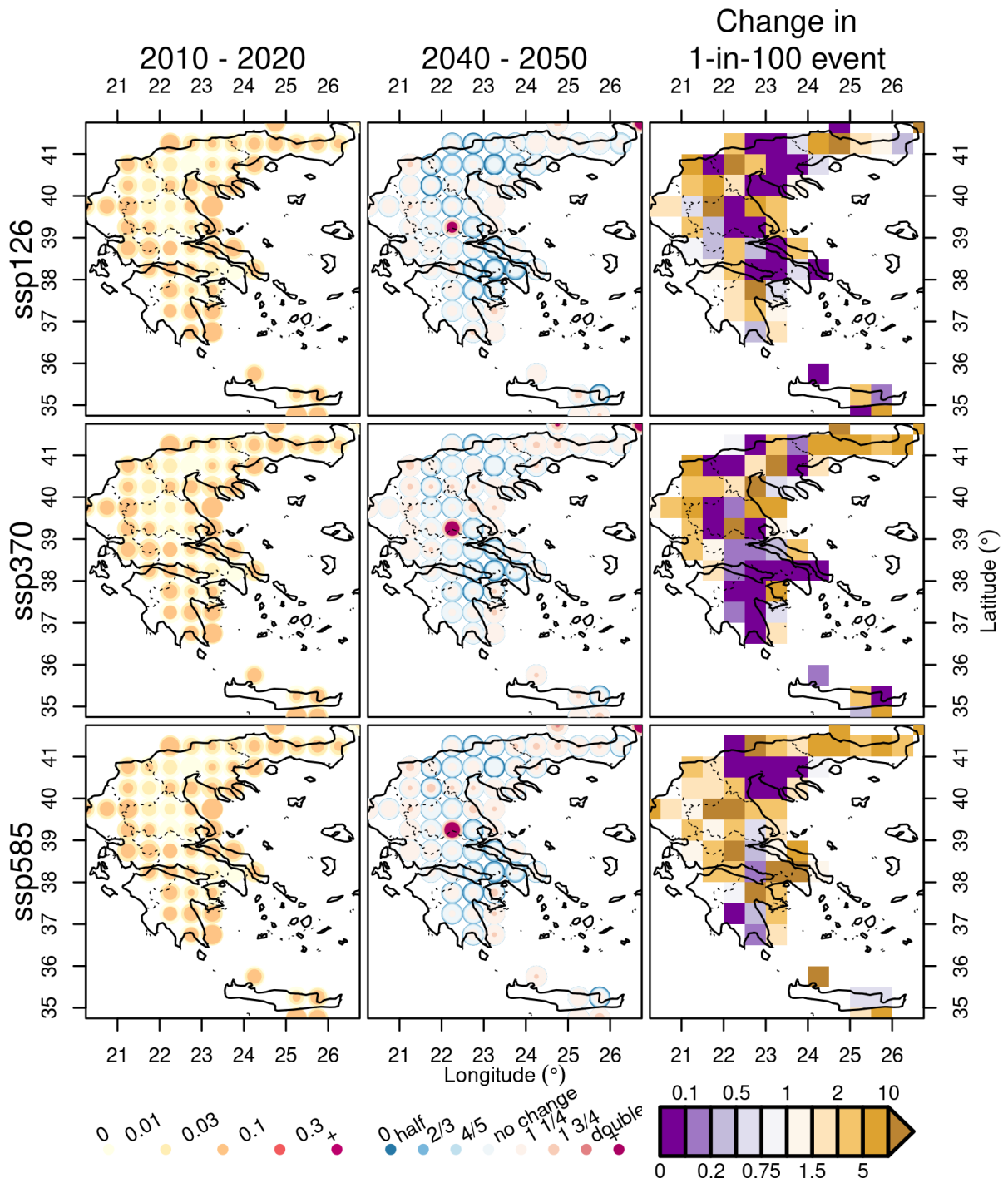
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Figure S18: Same as Figure 18 but covering 2040-2050



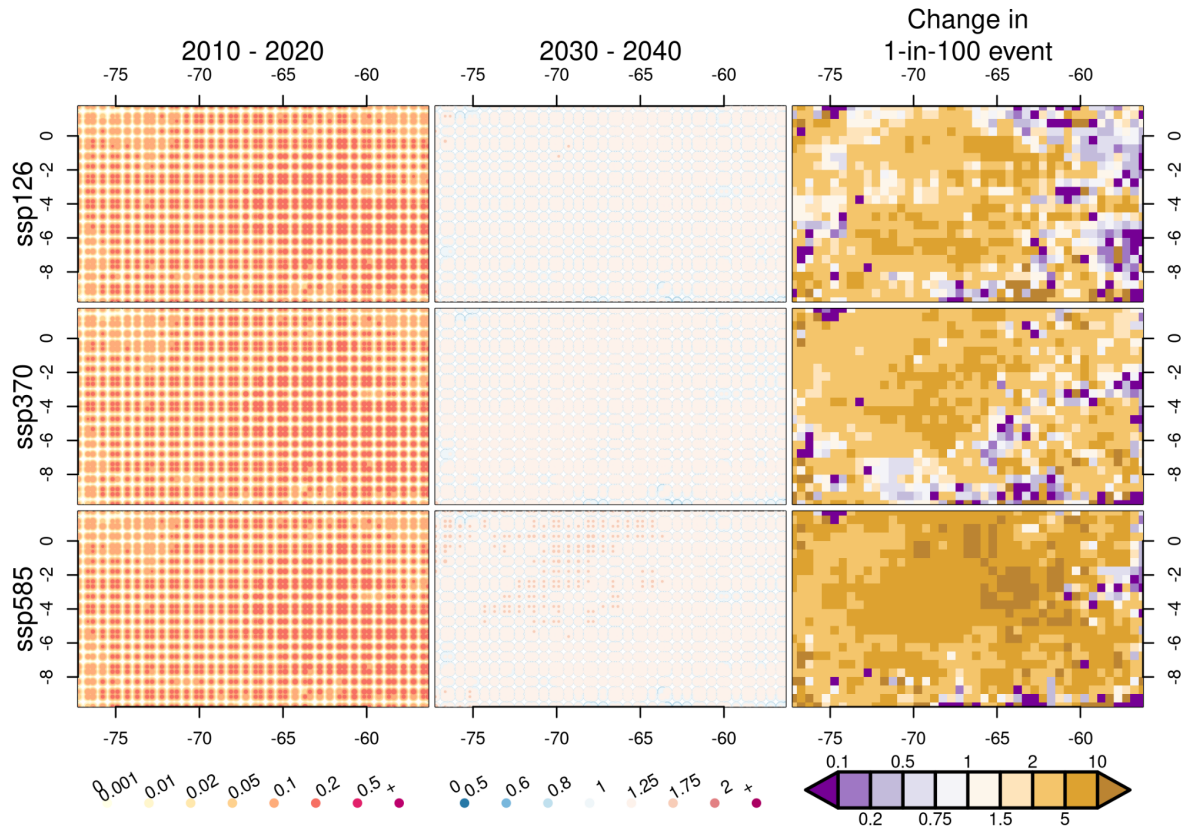
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Figure S19: Same as Figure 19 but covering 2030-2040



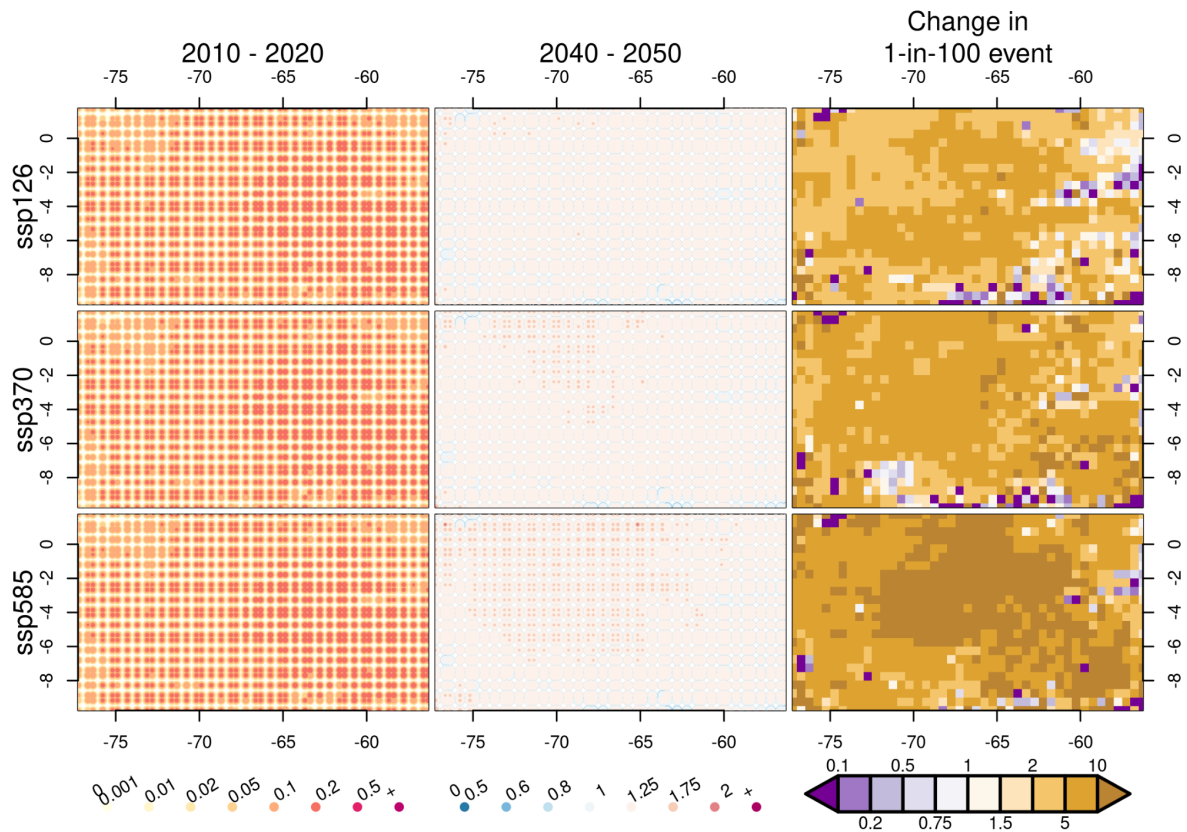
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Figure S20: Same as Figure 19 but covering 2040-2050



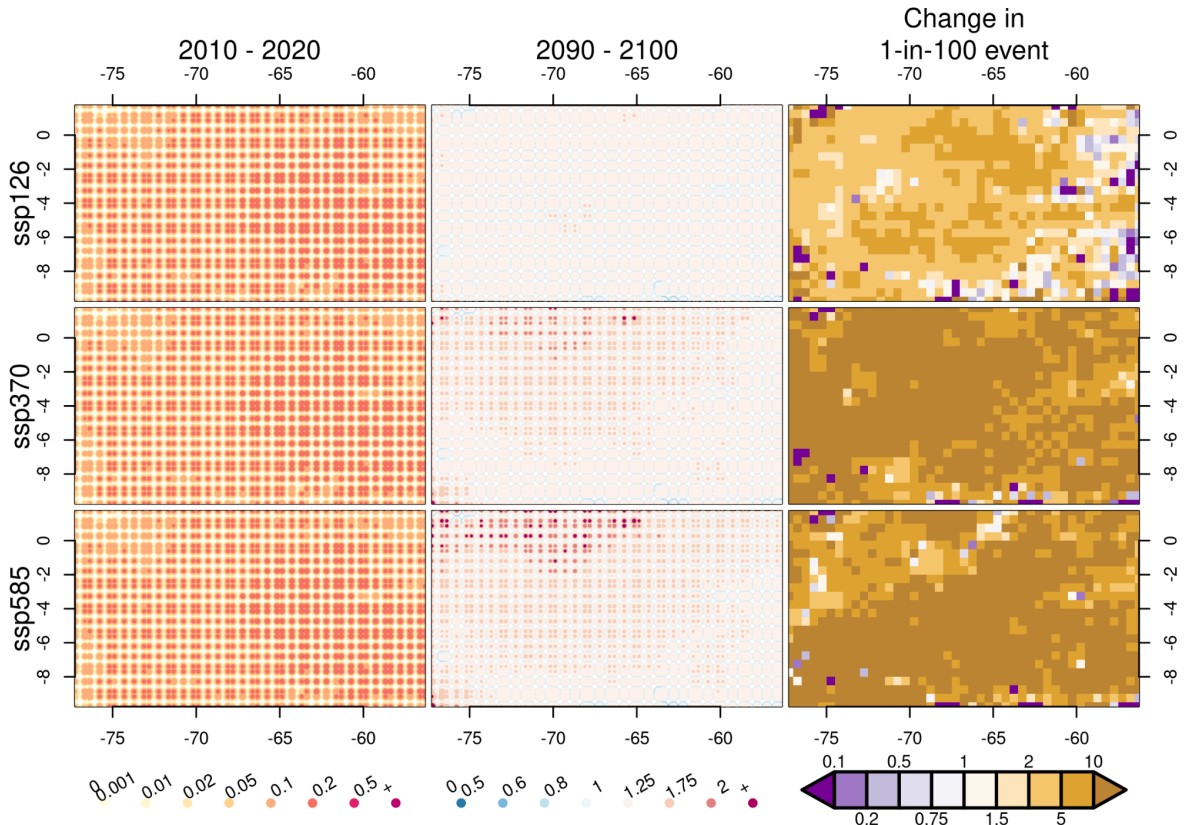
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Figure S21: Same as **Figure 18** but Western Amazonia covering 2030-2040 August-October.



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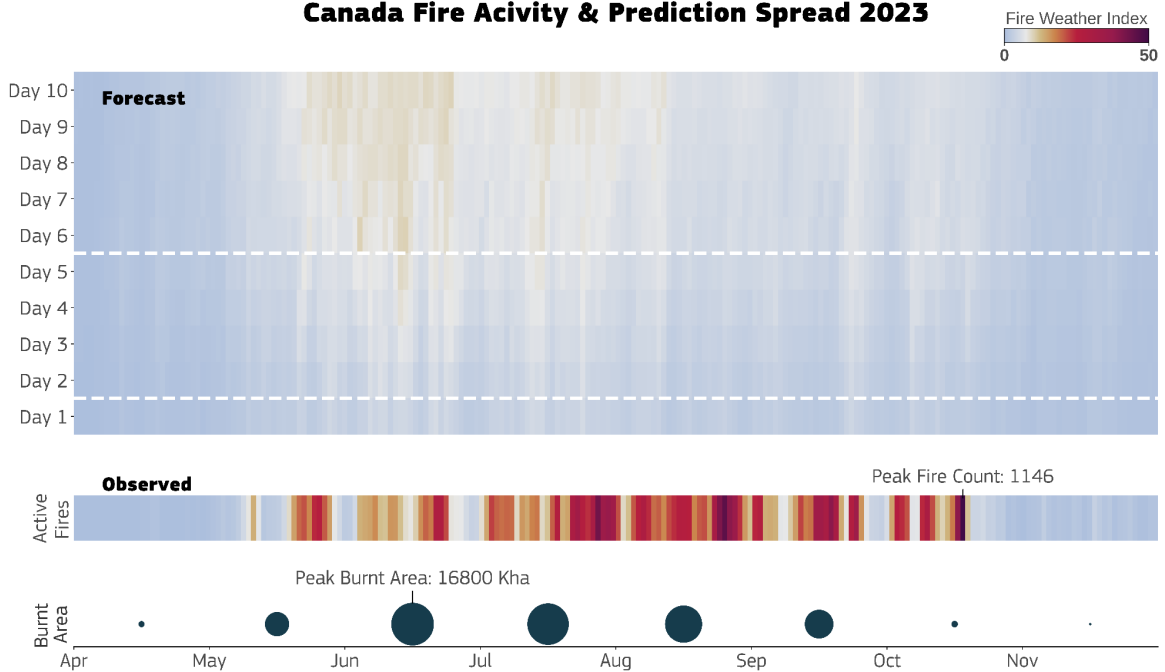
Figure S22: Same as **Figure S21** but covering 2040-2050



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Figure S23: Same as Figure S21 but for 2090-2100.

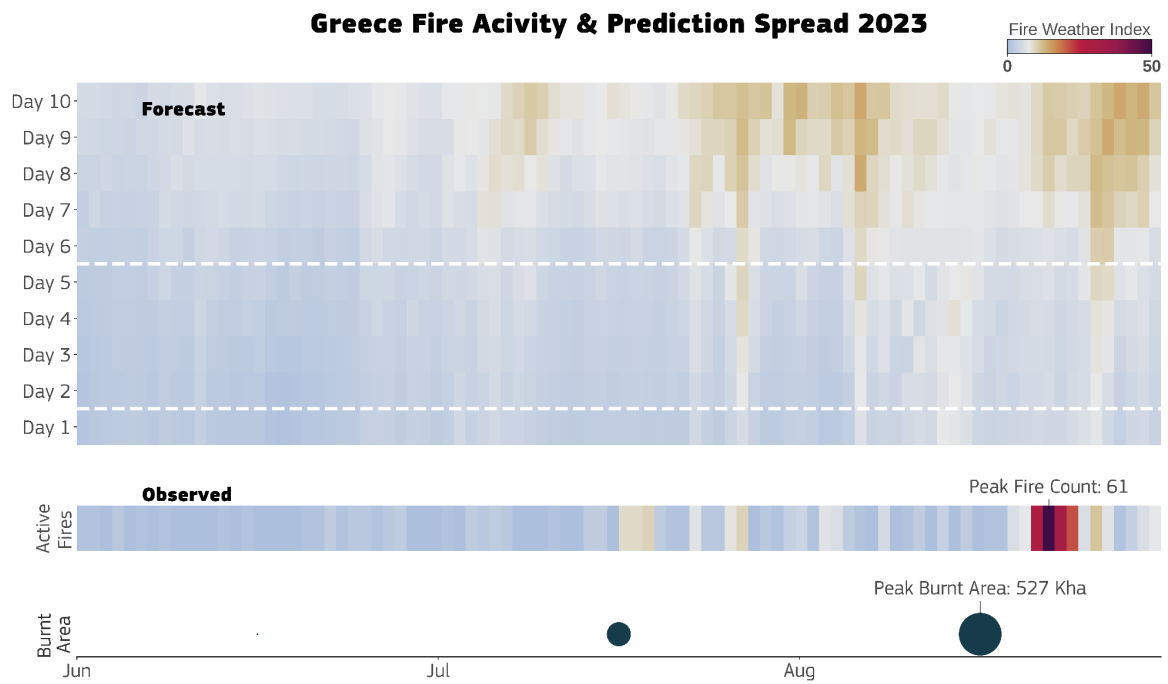
Canada Fire Activity & Prediction Spread 2023



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Figure S24: Ensemble spread in the prediction of the FWI in the first 10 days of forecast for Canada. The horizontal lines indicate changes in temporal resolution in the ECMWF weather forecasting systems. The spread accounts for as much as 10%-15% of the predicted values and, as expected, increases with lead time.

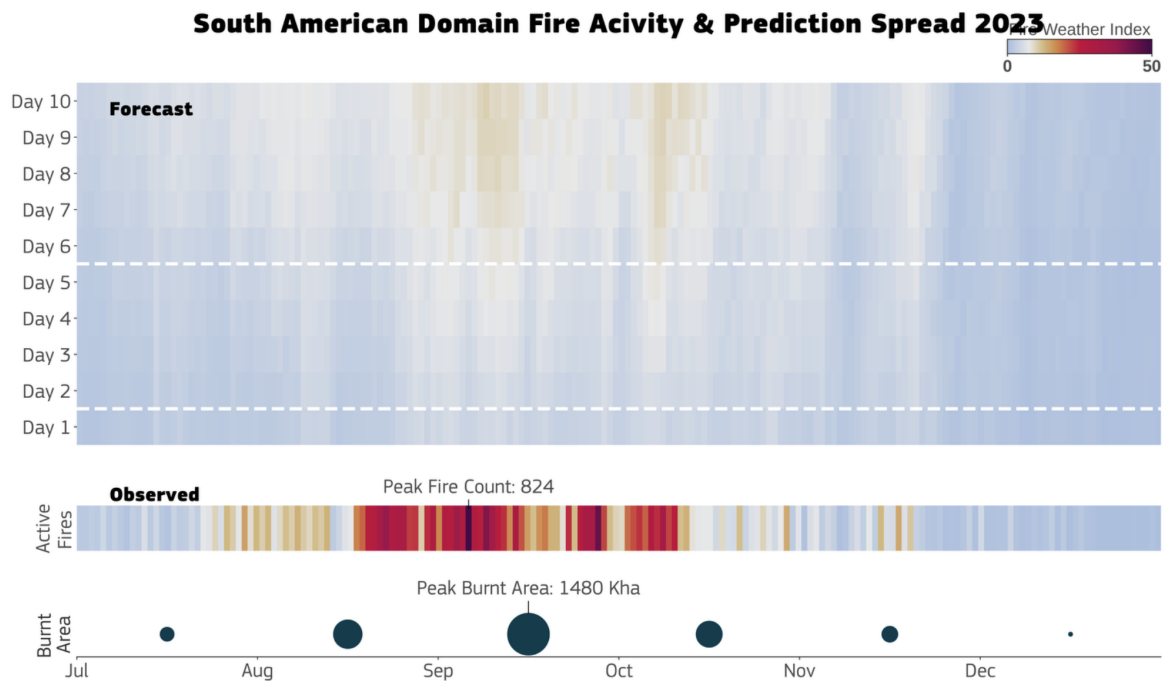
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167 **Figure S25:** Same as **Figure S24** but for Greece.

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170 **Figure S26:** Same as **Figure S24** but for Western Amazonia.

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177 **Table S1:** Overall PoF driver statistics summary from **Figure 9**. Values refer to variance in
 178 active fire counts explained by each control.

	Canada	Greece	Western Amazonia	Target
Weather	43% (Early May) 25% (June-July)	20%(August)	42% (August) 22% (Sept-Oct)	Active fires
Fuel dryness	13% (Early May) 11% (June-July)	9%(August)	26% (August) 10% (Sept-Oct)	
Fuel Abundance	14% (Early May) 12% (June-July)	11%(August)	22% (August) 22% (Sept-Oct)	
Others	29% (Early May) 52% (June-July)	60%(August)	10%(August) 46%(Sept-Oct)	

179 **Table S2:** Overall ConFire driver statistics summary from **Figure 10**. Values refer to the
 180 median and 5th-90th percentile range of anomalies in burned area caused by each control.
 181

	Canada	Greece	Western Amazonia	Target
Weather	19[5-45]% (May) 16[1-100]% (June-July)	24[4-140] % (August)	14[1-140]% (August) 29[2-45]% (Sept- Oct)	Burned Area Anomaly
Fuel dryness	6[-41-65]% (May) 9[-110-88]% (June-July)	-11[-170-21]% (August)	340[79-400]% (August) 100[62-370]% (Sept-Oct)	
Fuel Abundance	-1[-7- 0]% (May) 1[-3.9- 3.3]% (June-July)	7.5[1-72]% (August)	3[-0-16]% (August) 0 [-1-3.6]% (Sept- Oct)	
Humans	5[1-97]% (May) -0[-25-4]% (June-July)	19[-33 - 220]% (August)	8[1-68]% (August) 3[-1-20]% (Sept- Oct)	

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185 S1. Extended Methods

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187 S1.1 Data and Data Processing

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189 S1.1.1 ConFire vegetation fraction driving data

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191 In **Section 2.4.4.1**, we drive ConFire with tree and non-tree vegetated cover from the Joint UK
192 Land Environment Simulator Earth System impacts model (JULES-ES) at version 5.5 (Clark
193 et al., 2011; Mathison et al., 2023) driven with GSWP3-W5E5 forcings provided at a 0.5°
194 spatial resolution by ISIMIP3a. These runs are freely available at
195 <https://www.isimip.org/impactmodels/details/292/>. JULES-ES dynamically models vegetation
196 cover in response to meteorology, hydrology, nitrogen availability, and land use change.
197 JULES-ES has been extensively evaluated against snapshots and site-based measurements
198 of vegetation cover and carbon (Mathison et al., 2023; Burton et al., 2022; Clark et al., 2011;
199 Burton et al., 2019; Sellar et al., 2019). JULES-ES-ISIMIP has previously been used as driving
200 data for ConFire to perform future projections (UNEP et al., 2022), though using a previous
201 round of ISIMIP climate forcing (ISIMIP2b). As per (UNEP et al., 2022), vegetation responses
202 to JULES-ES’s internal fire model were turned off so as not to double-count the effects of
203 burning. However, in (UNEP et al., 2022), residual JULES-ES simulated biases in vegetation
204 cover were allowed to persist, increasing the uncertainty range of local vegetation cover and
205 resultant burned area responses. We therefore correct the bias in JULES-ES’s vegetation
206 cover using a trend-preserving empirical quantile mapping bias adjustment method,
207 implemented using the *ibicus* software package (Spuler et al., 2024). The bias adjustment
208 method calibrates a mapping between the empirical cumulative distribution function of each
209 surface cover type at each grid cell derived from the JULES-ES model output and the
210 corresponding quantiles in the MODIS remote sensed data at this grid cell over the reference
211 period (2003-2019). The method corrects the bias induced by the JULES-ES model rather
212 than the bias of the climate model, assuming that this has been removed by the ISIMIP3BASD
213 method (Lange, 2019).

214

215 The bias adjustment approach maps the empirical cumulative distribution function of each
216 surface cover type at each grid cell derived from the JULES-ES model output to the
217 corresponding quantiles in the MODIS VCF collection 6.1 remote sensed data (DiMiceli et al.,
218 2017) at this grid cell over the reference period (2002-2019). For Canada, where collection
219 6.1 does not extend north of 60DEG, we used collection 6 (Dimiceli and Others, 2015). This
220 mapping is subsequently applied to the surface information output from JULES-ES driven by
221 climate models over the historical (1994-2014) and future (2015-2099) period. To preserve
222 the trend in the vegetation cover over the future periods, additive detrending of the mean is
223 applied:

224

$$225 \quad x_{cm_fut} \rightarrow F_{obs}^{-1}(F_{cm_ref}(x_{cm_fut} + \bar{x}_{cm_ref} + \bar{x}_{cm_fut})) + \bar{x}_{cm_fut} - \bar{x}_{cm_ref} \quad (1)$$

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227 Here F_{cm_ref} is the empirical cumulative distribution of the model over the reference period,
228 F_{obs}^{-1} the inverse cumulative distribution function of the observations, x_{cm_fur} the quantile that
229 is adjusted and \bar{x}_{cm_ref} and \bar{x}_{cm_fur} the means of the model output over the reference and
230 future periods. This mapping is applied over a rolling window of 9 years over the future period.

231

232 The approach ensures that not only the mean but also the shape of the distribution is corrected
233 without assuming a parametric form, whilst also preserving additive trends driven by the future
234 climate model. Furthermore ensures continuity between the historical and future period by
235 using a rolling window over the future period.

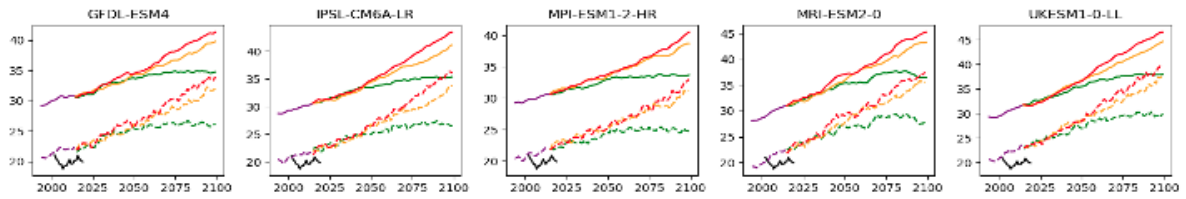
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237 The results were evaluated in terms of the ability of the bias correction method to reduce the
238 model bias over the historical period, as well as preserve the trend between the future and
239 historical period. It was found that the method corrects the bias well over the historical period
240 for most regions, variables and gridcells in both the mean and 80th percentile at each grid cell.
241 The trend between the future and historical period is well preserved in most regions and
242 gridcells, with less than 0.1% of gridcells overall experiencing an absolute trend modification
243 larger than 5%.

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245 To demonstrate the evaluation conducted, **Figure S24** shows the results for treecover over
246 North-Western Canada. The plots for the remaining regions, including tree and no-tree cover,
247 can be found in a notebook <https://github.com/jakobwes/State-of-Wildfires---Bias-Adjustment>.
248 Investigating the timeseries of average treecover over the region, we find that the correction
249 method reduces the bias over the historical period and matches the future period to the
250 historical period (**Figure S24a**). The cumulative distribution functions of average tree cover
251 merged over all spatial locations in observations and model match better after bias adjustment
252 (**Figure S24b**). They do not match perfectly, and we note that this is a non-calibrated aspect
253 that we do not expect to have zero bias but that is important to evaluate. Furthermore, we find
254 that the improvement in both mean and 80th percentile hold across the region (**Figure 19c**).
255 The trend between future and historical period is preserved for the majority of grid-cells, with
256 the absolute change in trend being close to zero for most grid-cells.

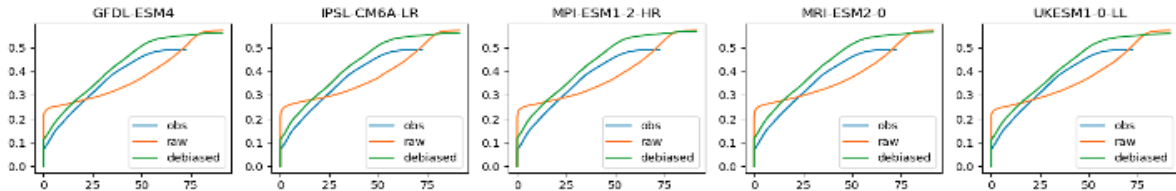
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Evaluation of bias correction results for the JULES vegetation model over North-Western Canada

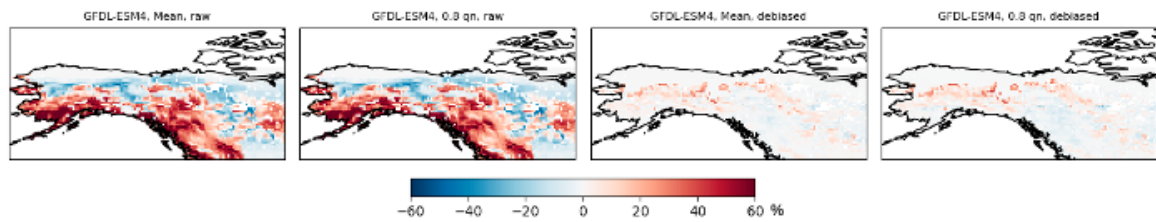
a) Timeseries of total tree cover of the area. MODIS VCF collection 6 in black. Raw model in solid lines, bias corrected model in dashed lines.



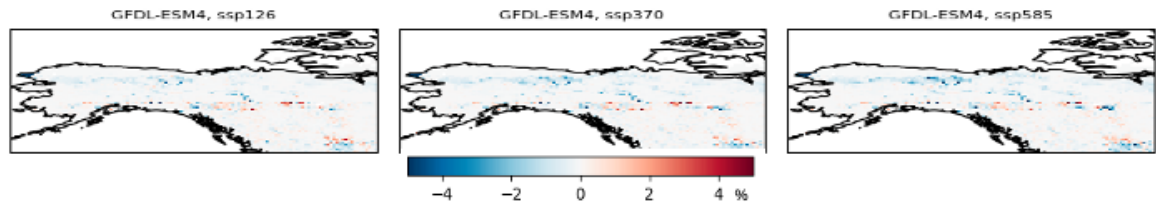
b) Cumulative distribution function across the region and historical time period.



c) Absolute model bias in mean and 80th percentile at each grid cells before (left two plots) and after (right two plots) bias correction.



d) Absolute difference in trend between future and historical period between raw and bias corrected model.



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S1.2 Modelling Frameworks

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S1.2.1 PoF

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The Probability of Fire (PoF) system uses gradient-boosted decision trees to provide a probability forecast of active fire occurrence (McNorton and Di Giuseppe, 2024). The supervised algorithm which trains an ensemble of decision trees uses regularization techniques to prevent overfitting (Chen & Guestrin, 2016). The training, based on 2010-2014 MODIS active fire detections, classifies a positive fire event as any detection within a 9 km grid cell.

282 The relative contribution of each input control to the model prediction is evaluated using
 283 Shapley values, computed using the Shapley Additive exPlanations python library (Lundberg
 284 & Lee, 2017). The SHAP value indicates the importance of each feature in a model, where a
 285 positive SHAP value reflects a positive impact on the model prediction and a negative SHAP
 286 value reflects a negative impact. Specifically for this study we use the TreeExplainer, which
 287 computes the SHAP values by interrogating the structure of the decision trees within the model
 288 based on the input feature values. The probability controls are then normalised and grouped
 289 into the four categories given in **Table 3** of the main text. By combining these with the total
 290 amount of fires predicted for a given area we can attribute those fires into one of the four
 291 controls. The ‘Other’ control also includes fire occurrences not predicted by the model. This is
 292 computed given by:

$$294 \quad \text{Other} = \text{SHAP}[\text{Other}] + \max(0, \text{Area_Total_Fires_Observed} \\ 295 \quad \quad \quad - \text{Area_Total_Fires_Predicted}) \quad (2)$$

298 Where, $\text{SHAP}[\text{Other}]$, is the contribution of the ‘Other’ control to the total predicted fires for a
 299 given region and, $\text{Area_Total_Fires_Observed}$ and $\text{Area_Total_Fires_Predicted}$ are the total
 300 number of observed and predicted fires for the same region.

302 S1.2.2 ConFire

304 ConFire is a burned area attribution tool, used for trend detection and attribution (Kelley et al.,
 305 2019), event attribution (Kelley et al., 2021) and future projections (UNEP et al., 2022).
 306 ConFire finds the likelihood of causes of or changes in BA by optimising a simple, semi-
 307 empirical process representation model by applying Bayes Theorem. In our case, Bayes
 308 Theorem states that the likelihood of a model configuration described by a parameter set $\{\beta\}$
 309 and monthly explanatory variables (i.e model driving data) $\{X_{iv}\}$ given some training
 310 observation of monthly burned area fraction $\{Obs_i\}$ from MODIS MCD64A1, for cells i , is
 311 proportional to the prior probability of $\{\beta\}$ ($P(\{\beta\})$) multiplied by the probability of the
 312 observations given that model configuration:

$$313 \quad P(\{\beta\}|\{Obs_i\}, \{X_{iv}\}) \propto P(\{\beta\}) \times P(\{Obs_i\} | \{X_v\}, \{\beta\}) \quad (3)$$

315 We use the zero-inflated logit distribution introduced by (Kelley et al., 2021) as our update
 316 distribution, as this is specifically designed to better represent the tails of the distribution during
 317 fire events:

$$318 \quad P(\{Obs_i\} | \{X_{iv}\}, \{\beta\}) = \prod_i^N P(Obs_i | \{X_v\}_i, \{\beta\})$$

$$319 \quad P(Obs_i = 0 | \{X_v\}_i, \{\beta\}) = (1 - M(\{X_v\}_i, \{\beta_M\})^{P_1}) \times (1 - P_0) \\ 320 \quad P(Obs_i > 0 | \{X_v\}_i, \{\beta\}) \\ 321 \quad = (1 - P(Obs_i = 0 | \{\beta\})) \times \mathfrak{N}(\text{logit}(Obs_i) - \text{logit}(M(\{X_v\}_i, \{\beta_M\})), \sigma) \quad (4)$$

322 where $\{\beta_M\}$ is the set of parameters related solely to the underlying model, M , $\text{logit}(x) =$
 323 $\text{log}\left(\frac{x}{1-x}\right)$, P_0 , P_1 and σ are parameters within the full set $\{\beta\}$ which describe the model error
 324 and $\mathfrak{N}(\mu, sd)$ is a normal distribution with mean of μ and standard deviation of sd .

328 The model, M , simulates fractional BA (fraction) via a number of controls. For attribution and
 329 outlook, these controls follow (Kelley et al., 2021; Burton et al., 2019): Fuel load, fuel moisture,
 330 ignitions and suppressions. This follows the general model structure of global fire models
 331 (Hantson et al., 2016; Rabin et al., 2017) and is most appropriate for looking at long term,
 332 coarse fire drivers (Moritz et al., 2005). For driver assessment, we separate out an additional

333 control for “fire weather” and introduce a “snow cover” control. Model BA is the product of
 334 these controls, c :

$$335$$

$$336 M(\{X_v\}, \{\beta_M\}) = F_{max} \times \prod_c f(\{X_c\}, \{\beta_c\}) \quad (5)$$

$$337$$

338 Where F_{max} describes maximum monthly burned area fraction and is an optimizable
 339 parameter in set $\{\beta_M\}$, $\{X_c\}$ are the BA driving variables, $\{\beta_c\}$ the parameters related to control
 340 c and f is the function that describes the control influence on BA. Each control describes the
 341 expected BA if all other controls imposed no limitation on burning - for example, when c is fuel,
 342 $f(\{X_c\}, \{\beta_c\})$ describes the BA in perfectly dry conditions with saturated ignitions and no
 343 suppression. To achieve this, f is the logical function:

$$344$$

$$345 f(\{X_c\}, \{\beta_c\}) = 1 / \left(1 - \exp(-\beta_{c,0} - \sum_j \beta_{c,j} \times X_j) \right) \quad (6)$$

$$346$$

347
 348 where $\beta_{c,j}$ is the contribution of driving variable X_j to the control and $-\beta_{c,0}$ is a parameter that
 349 can shift the midpoint of the sigmoid curve.

350
 351 All variables X_v where normalised to be between $[0, 1]$ based on the training data to aid priors
 352 selection and optimization - though analytically this should have no impact on our results. Our
 353 priors fix the direction each drive can influence a control (drivers and direction are listed in
 354 Table 3 and 5) but beyond this relatively uninformed. Priors for $\beta_{c,j}$ where described by a log-
 355 normal distribution with a μ of 0 and σ of 10, and set to be positive for liberative drivers (one
 356 that increases the strength of a control) and negative for suppressive (ones that reduce the
 357 strength of a control). $\beta_{c,0}$ priors were set to a normal distribution with a mean of 0.5 and a
 358 standard deviation of 1. F_{max} and P_0 priors were set as a uniform distribution between 0 and
 359 1 σ was set to a half-normal with mean of 0 and standard deviation of 10.

360
 361 We sampled the posterior distribution using Bayesian inference following a similar protocol
 362 to (Barbosa, 2024) with the pymc python package version 5 (Abril-Pla et al., 2023),
 363 employing 100 chains each over 1000 warm-up iterations (that were not subsequently used)
 364 and 100 sample iterations using the No-U-Turns Hamilton Monte Carlo sampler (Hoffman
 365 and Gelman, 2011) while utilising 50 % of the data or a minimum of 6000 grid cells. To
 366 sample the posterior distribution, we then randomly sample 50 iterations from each chain,
 367 thereby approximating the posterior with 1000 ensemble members. As per Barbosa (2024),
 368 for evaluation (**Figure S28-S39**) we trained the first half of the period and tested on the
 369 second half. For the rest of the results, we trained on the full period.

370
 371 We obtaining probability distributions from the model posterior for our results, ConFire offers
 372 two probability, which we have adapted slightly from (2021) :

- 373
- 374 1. The likelihood of different levels of burning for a specific event (i.e a grid cell in a
 375 given timestep) which considers uncertainty explained by the model and residual
 376 uncertainty described by our error parameter, σ . We use this when we are
 377 comparing a single grid of cells and months, such as for evaluation, and for
 378 assessing the un. The likelihood of a Burned Area, BA , under drivers, X , which can
 379 be out-of-training sample, is:

$$380$$

$$381 P(BA | (X_v, \beta | \{Obs_i\}, \{X_{iv}\})) = \int_{\beta} P(\beta | \{Obs_i\}, \{X_{iv}\}) \times P(BA | \beta) d\beta \quad (7)$$

$$382$$

383 Where $P(BA | \beta)$ is take from equation 4.

384

385 When building distributions for multiple grid cells or time periods, as with building a
 386 climatology in **Section 3**, we convolute the probability distributions of individual time
 387 periods and cells following equations in (Kelley et al., 2021). Converting probabilities
 388 over a large number of cells gives us the second measure.
 389

390 The emergent probability of different mean levels of BA over many events explained directly
 391 by the model and its driving variables. We use this when assessing the emergent likelihood
 392 of burning in **Section 4** and **Section 5**. This is the same as taking the mean of n simulations
 393 in equation 7 as n tends to infinity. Doing this, $P(\{Obs_i\} | \{X_{iv}\}, \{\beta\})$ from equation 4 will tend
 394 towards a BA of model M output weighted by the likelihood of a zero BA:
 395

$$396 \quad D(BA) = \lim_{n \rightarrow \infty} \left[\sum_{i=1}^n P(BA | (X_v, \beta | \{Obs_i\}, \{X_{iv}\})) / n \right]$$

397 2.

$$398 \quad = \int_{\beta} M(\{X_v\}_i, \{\beta_M\}) \times (1 - M(\{X_v\}_i, \{\beta_M\})^2) \times (1 - P_0) d\beta \quad (8)$$

399 For attribution and future projections, ConFire produces correctly ranked by consistently
 400 biased probability distributions (**Supplement Section “Change in Likelihood of High
 401 Burned Area in 2023 due to Total Climate Forcing and Socioeconomic factors”**). The
 402 final step is therefore to introduce a correction factor. As this distribution bias is constant
 403 across the observed BA distribution, a simple scaling factor is all that's needed. To do this,
 404 we assign the likelihood associated with the BA in equation 8 with a scaled burned area
 405 (BA^*) so that the mean of the sample distribution matches the mean of the observation for
 406 the period 2003-2019.
 407

$$408 \quad BA^* = BA \times \Sigma(\{Obs_i\}) / \int_0^1 D(BA) \times BA dBA \quad (9)$$

409
 410 BA^* is then used in equation 8.
 411

412 **S1.2.3 Attributing Fire Weather**

413

414 Bias Correction

415 We evaluated the individual variables in the FWI (see evaluation), and found that each variable
 416 was slightly biased compared to ERA5 reanalysis. We therefore applied a bias correction to
 417 the final FWI, rather than bias-correcting each individual variable.
 418

419 We bias-corrected the HadGEM3 2023 large ensemble based on a bias assessment of the 15
 420 historical members from 1960-2013 vs. ERA5 observation-driven FWI, using a simple linear
 421 regression on fwi transformed using:
 422

$$423 \quad fwi_* = \log(\exp(fwi) - 1) \quad (10)$$

424 to remove the physical bound at 0. We use this instead of using a straight \log transformation
 425 as it ensures numerical stability at higher values, crucial when dealing with extreme FWI
 426 values, thereby avoiding blow-up effects. It also preserves the extreme tail of the FWI
 427 distribution, allowing us to accurately capture and analyse critical events associated with high
 428 fire risk..
 429

430 We perform a simple linear regression on ERA5 and on each historical member to obtain the
 431 basic regression parameters:

$$432 \quad fwi_* \sim fwi_{*,0} + \Delta_{fwi} \times t \quad (11)$$

433 Where t is time, and $t = 0$ is set to 2023, Δ_{fwi} is the rate of change, or trend, of fwi_* and $fwi_{*,0}$
 434 is the estimated fwi_* for 2023. Our bias correction is therefore based on present-day levels of
 435 warming, taking account of the additional warming from 2013-2023 (assuming the trend from
 436 1960-2013 continues to 2023 linearly). If anything this is likely conservative given that warming
 437 rates may have increased more rapidly in the last 10 years.

438
 439 We generate the bias-corrected 2023 ensemble by correcting each of the 525 present-day
 440 ensemble members against each of the 15 historical members (creating an ensemble of 7875
 441 members). Due to the perturbation procedure used to generate the 2023 ensemble from the
 442 historic (Ciavarella et al., 2018), we can not assume that present-day members pair to
 443 historical members. We therefore iterate over all possible pairs:
 444

$$445 \quad fwi_{corrected} = (fwi_{*i} - fwi_{*,0,j}) \times \sigma(fwi_{*era5}) / \sigma(fwi_{*j}) + fwi_{0,*era5} \quad (12)$$

$$446 \quad \sigma_{\Delta}(fwi_{*}) = sdev(fwi_{*} - \Delta_{fwi} \times t)$$

448 Where i is a present-day ensemble member, and j is a historical member.

449
 450 We finish by applying the inverse of the transformation from equation 10 :

$$451 \quad fwi_{corrected} = \log(\exp(fwi_{*,corrected}) + 1) \quad (13)$$

454 Probability Ratio

455 We use the ERA5 2023 FWI for our event threshold in each region, using the month of peak
 456 anomaly from **Figure S2** in each region. We use this threshold to calculate the probability ratio
 457 (PR) of the event occurring with and without climate change. To calculate the PR, we find the
 458 number of ensemble members that exceed the 2023 ERA5 FWI value in the bias-corrected
 459 ALL simulation, and divide this by the number of members that exceed the same value in the
 460 bias-corrected NAT simulation, bootstrapping 10,000 times to giving the probability of
 461 exceeding the observed 2023 FWI value in a world with and without climate change plus
 462 uncertainty bound for the 5-95th percentile.

$$463 \quad PR = p(ALL) / p(NAT)$$

465 S1.2.4 FireMIP

466
 467 For the multi-model ensemble we use simulations from the ISIMIP3a fire sector, as published
 468 in (Burton & Lampe et al. 2023). The 7 models reporting BA for ISIMIP3a are shown in the
 469 table below. The methodology follows the ISIMIP3a Impacts Attribution protocol, as outlined
 470 in (Mengel et al., 2021), where the factual historical simulations are driven with GSWP3-W5E5
 471 reanalysis data, and the counterfactual simulations are the same historical data which has
 472 been detrended via quantile mapping (Mengel et al., 2021).

473
 474 As outlined in (Hantson et al., 2016), the spread in the absolute BA is large amongst the
 475 observations, models and regions and therefore a normalised relative anomaly (RA) rather
 476 than absolute BA is used for the analysis. To calculate the RA in present day BA, we subtract
 477 the counterfactual mean, and divide by the counterfactual mean. By comparing both factual
 478 and counterfactual experiments to the counterfactual mean, we are looking at the fractional
 479 increase in BA driven by climate change compared to a baseline without climate change.
 480 Based on model performance by AR6 region, a region-specific weighting is also applied
 481 following (Knutti et al., 2017). The weighting is based on the model's distance to the observed
 482 BA temporal RA using both FireCCI5.1 and GFED5, measured using NME as per (Kelley et

483 al., 2013). To measure the uncertainty, random noise is generated and scaled by the
484 climatological RMSE of each model. This noise is then added to the modelled relative
485 anomaly, this process is repeated 1000 times. This performs the same function and the
486 uncertainty quantification from model error as Equation 4 does for ConFire. Then,
487 bootstrapping is applied to the monthly regional BA RA (now with noise added in) according
488 to the weight for each model. Uncertainty is calculated by taking the 2.5-97.5th percentile of
489 the resultant histogram. All results are reported as P50 [P2.5, P97.5]. The methods are
490 explained in full in (Burton & Lampe et al. 2023).

Table S3: FireMIP Models used for attributing median burned area. Table reproduced from (Burton & Lampe et al. 2023)

Model		CLASSIC	INFERNO	LPJ-GUESS-SIMFIRE-BLAZE	LPJ-GUESS-SPITFIRE	ORCHIDEE-MICT-SPITFIRE	SSiB4/TRIFFID	VISIT
Fire Model		CLASSIC	INFERNO	SIMFIRE	SPITFIRE	SPITFIRE	Li	After (Thonicke et al., 2001)
Land / Vegetation		CLASSIC	JULES	LPJ-GUESS	LPJ-GUESS	ORCHIDEE	SSiB	VISIT
Dynamic Veg	Physiology	Yes	Yes, via TRIFFID	Yes	Yes	Yes	Yes, via TRIFFID	Yes
	LAI	Yes	Yes, via TRIFFID	Yes	Yes	Yes	Yes	Yes
	Bio-geography	No	Yes, via TRIFFID	Yes	Yes	Yes	Yes	No
Nitrogen Cycle		Yes	Yes	Yes	Yes	No	Yes	Yes, but C-N coupling is limited
No. PFTs		9	13	17	17	19	7	33 (biome types)
No. Soil Layers		20	4	2	2	11	3	2
Fuel		Vegetation and litter	Vegetation & top soil layer as proxy for litter	Vegetation, litter	Litter	Vegetation and litter	Vegetation and litter	Litter
Ignitions	Natural	Prescribed lightning	Prescribed lightning	SIMFIRE describes annual BA + fire-climatology -> daily BA used as Fire-Probability	Prescribed lightning	Prescribed lightning	Prescribed lightning	Probabilistic based on fuel wetness

	Anthropogenic	Prescribed population density	Prescribed Population density	SIMFIRE includes suppression by humans	Prescribed population density	Prescribed population density	Prescribed population density	No
Suppression		Prescribed population density	Crops, population density	Crops (100%), prescribed population density (Hyde3.1)	Crops, population density	Prescribed population density, crops	Prescribed population density and GDP	Low fuel load
Spread		Wind speed and soil moisture	None	Daily BA (no explicit spread)	Rothermel equations including wind speed, tree fraction, grass fraction, fuel moisture, fuel load and characteristics	wind speed, tree fraction, grass fraction, fuel moisture, fuel load	Wind speed and soil moisture	None
Model inputs		SW & LW radiation, precipitation, air temperature, specific humidity, wind speed, atmospheric pressure, population density, lightning	SW & LW radiation, precipitation, air temperature, specific humidity, wind speed, population density, lightning	SW radiation, precipitation, air temperature (mean, min, max), relative humidity, wind speed	SW radiation, precipitation, air temperature, specific humidity, wind speed, atmospheric pressure, population density, lightning	SW & LW radiation, precipitation, air temperature, specific humidity, wind speed, atmospheric pressure, PFT map, population density	SW & LW radiation, precipitation, air temperature, specific humidity, wind speed, atmospheric pressure, population density, and GDP, peat map, land cover change	Air temperature, precipitation, air vapor pressure, cloudiness, wind
Resolution		1 deg	0.5 deg	0.5 deg	0.5 deg	0.5 deg	0.5 deg	0.5 deg
References		(Melton et al., 2020)	(Burton et al., 2019, 2020; Mangeon et al., 2016)	(Rabin et al., 2017; Smith et al., 2014; Knorr et al., 2014)	(Rabin et al., 2017; Smith et al., 2014; Thonicke et al., 2010; Lehsten et al., 2009)	(Yue et al., 2014, 2015)	(Huang et al., 2021, 2020; Li et al., 2012; Hugelius et al., 2013; Li et al., 2013)	(Ito, 2019)

493

494

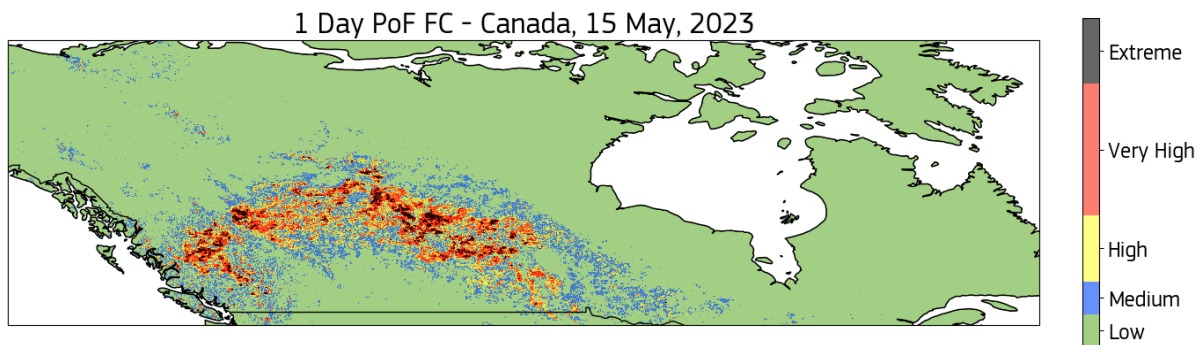
495 S2. Evaluation

496 S2.1 POF

497

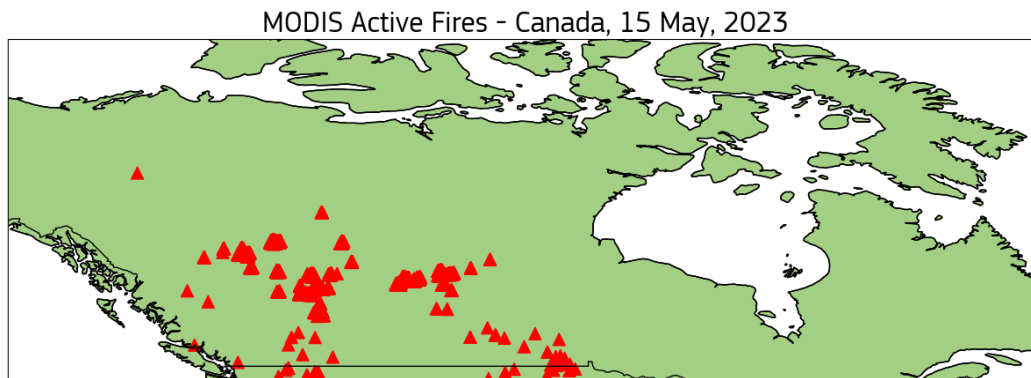
498 The PoF model, trained on observed fire activity, provides a daily probability of fire occurrence
499 based on the input variables described in **Table 3** of the main text. The three cases explored
500 in the main study can be visualised as fire risk maps at a 1 km resolution, higher than the 9
501 km used for attribution. The 1km predictions show that whilst PoF often fails to capture the
502 true total number of active fires, the relative attribution is likely to be accurate given by the
503 models ability to capture the spatiotemporal pattern of fire activity reflected by the forecast
504 danger shown in the figures below. Of the three case studies the model accurately reflects fire
505 activity for Canada and Western Amazonia, and whilst high fire danger is modelled over
506 Alexandroupolis, Greece, it fails to capture the severity of the event.

507



508

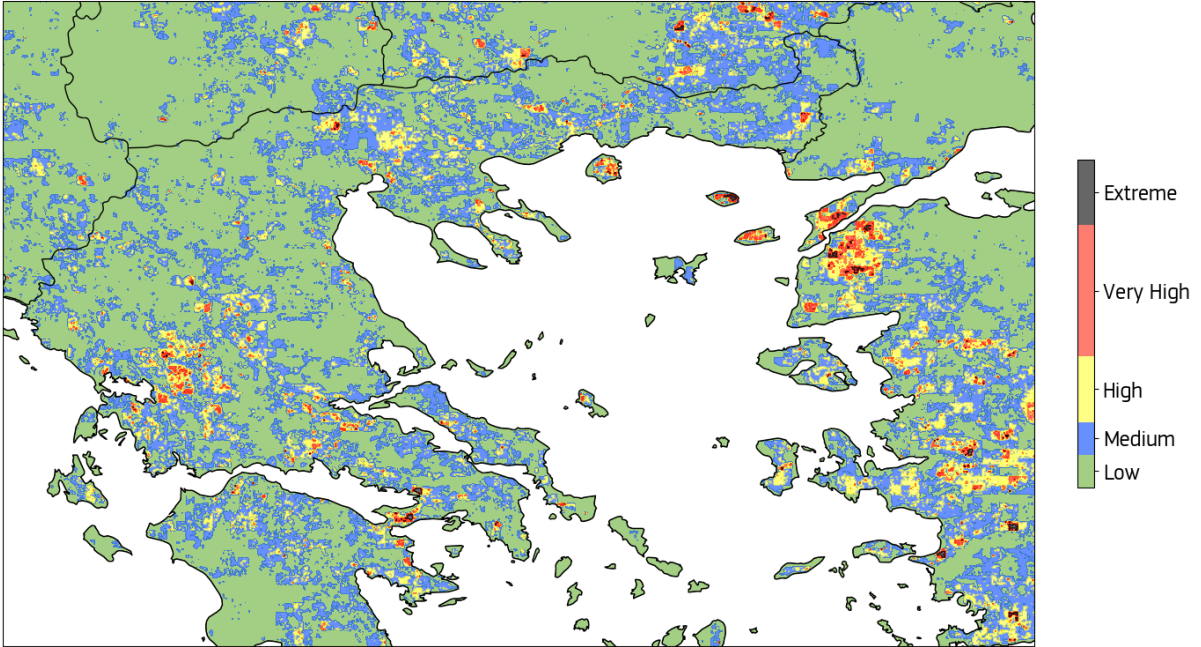
509



510

511 **Figure S25:** Spatial representation of the day 1 PoF forecast at ~1km resolution expressed as a danger rating for the 15th May over Canada (top). MODIS active fire detections for the same day and domain (bottom).
512
513

1 Day PoF FC - Greece, 21 August, 2023



514

MODIS Active Fires - Greece, 21 August, 2023



515

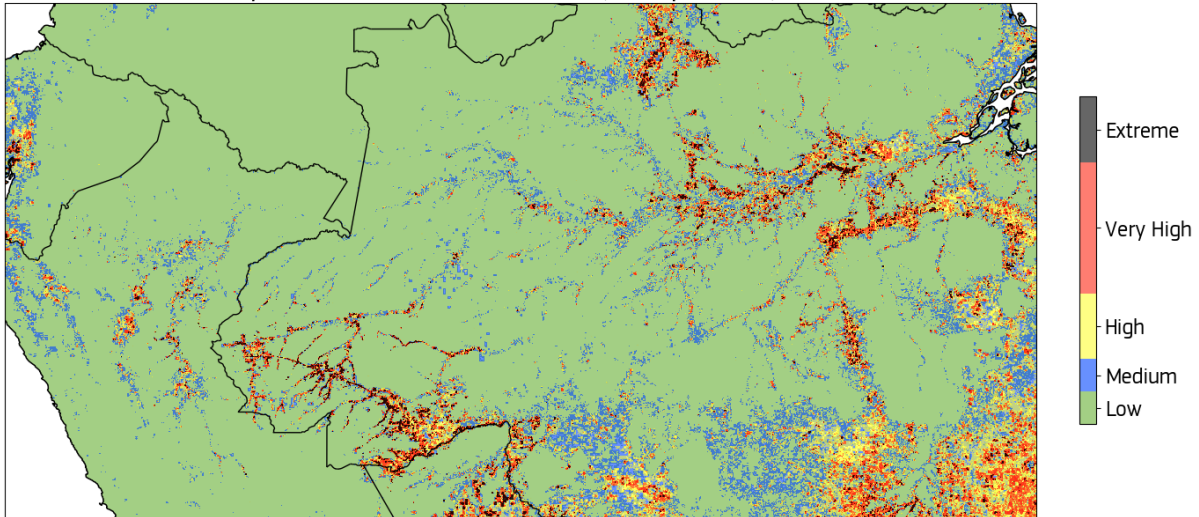
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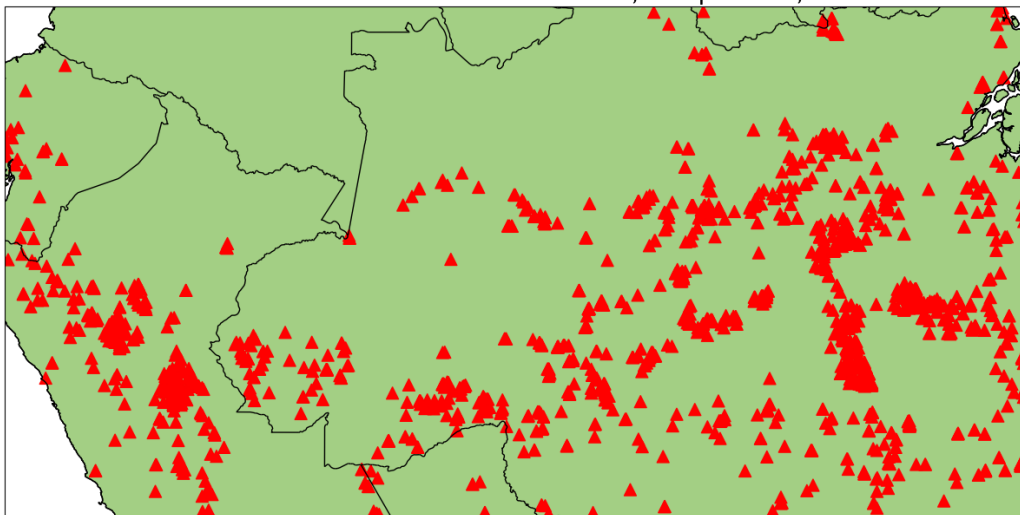
Figure S26: Spatial representation of the day 1 PoF forecast at ~1km resolution expressed as a danger rating for the 21st August over Northern Greece (top). MODIS active fire detections for the same day and domain (bottom).

1 Day PoF FC - Western Amazonia, 9 September, 2023



519

MODIS Active Fires - Western Amazonia, 9 September, 2023



520

521 **Figure S27:** Spatial representation of the day 1 PoF forecast at ~1km resolution expressed
522 as a danger rating for the 9th September over Western Amazonia (top). MODIS active fire
523 detections for the same day and domain (bottom).

524

525

526

S2.2 ConFire

527

528 The ConFire model simulates a probability distribution of BA which, unlike most numerical or
529 ensemble-based models, requires a probabilistic technique for evaluation.

530 The uncertainty range of the ConFire is crucial for the analysis in this study. We obtain
531 confidence in our results by seeing if the shift of the model's probability distributions is
532 significant compared to the size of the uncertainty of that distribution. Suppose the uncertainty
533 range is larger than any change when testing for i.e, attributing with or without climate change,
534 future changes, or seasonal anomaly. In that case, the framework will tell us, and our results
535 will show that these are unlikely/not significant. Conversely, if the change in distribution is
536 larger than the model's uncertainty range, we can make a confident statement even if that
537 model is uncertain.

538

539 As the precision of the modeling framework is inherent in the analysis itself, the main aspect
540 to evaluate is the ability of the model's probability distribution to represent the range of
541 uncertainties when tested against observations accurately. To do this, we followed the
542 evaluation procedure outlined in (Barbosa, 2024), which we summarise here.

543

544 We trained the model during the first half of each period used in the analysis and performed
545 subsequent evaluations on the second half. The training period for near-real-time driver
546 assessment was 2014-2018, and for the attribution/future projections run, 2003-2011. The
547 evaluation period was 2019-2023 for driver assessment and 2012-2019 for attribution/future
548 projections. Using a different period from the optimization ensures an independent model
549 evaluation and provides an indication of how well the framework captures uncertainty in out-
550 of-temporal sample observations.

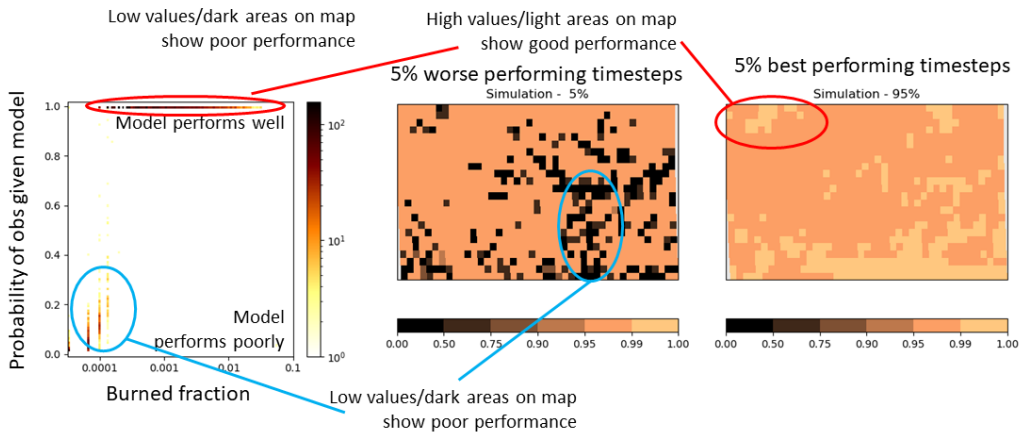
551

552 The FLAME system (Barbosa, 2024) that we merge with ConFire automatically generates a
553 series of evaluations which we show for region region in turn in the subsequent two sections.
554 While the techniques are outlined in (Barbosa, 2024), these automated figures have not
555 previously been published. So alongside the evaluation procedure below is a guide to interpret
556 if these plots show a good model performance.

557

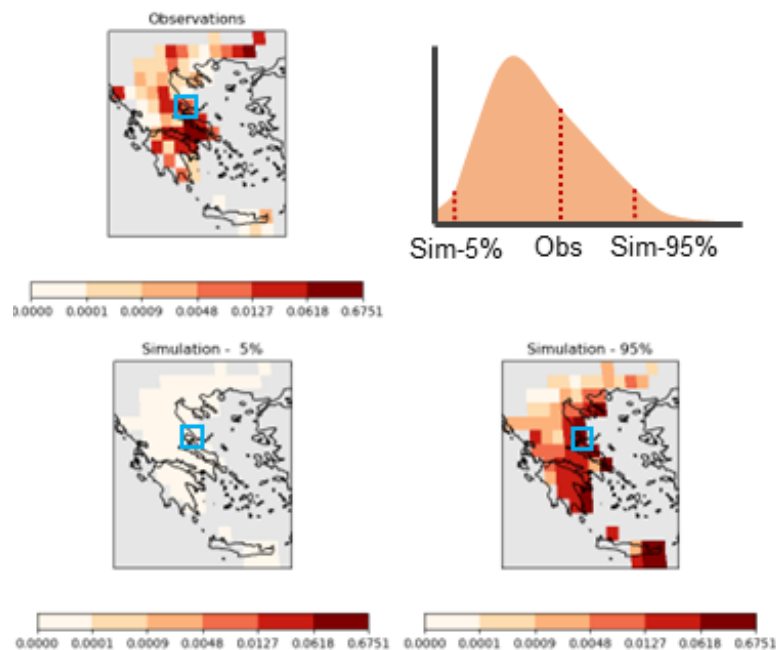
558 For the evaluation period, we assess how well the model predicts new observations by testing
559 how likely the observations are given the optimized model (equation 7). While this sounds
560 counter-intuitive, we do this rather than test the model given the observations because the
561 model doesn't yield a single answer or a set of numbers, but rather a distribution of model
562 parameters and output, reflecting inherent uncertainty in the processes. This approach allows
563 for comprehensive testing of the entire model's posterior probability distribution at once and
564 provides insight into the model's ability to generate the observed distribution and capture the
565 uncertainty in the modeled process. We approximate the probability of an observation given
566 our model by sampling 10 parameter ensemble members from each of our 100 chains,
567 providing us with 1000 ensemble members, and sample the likelihood as per (Kelley et al.,
568 2021). The example below, taken from **Figure S32**, shows how we summarise this for each
569 observation (scatter plot left) and all observations in a time series for each cell (middle and
570 right). If the model performed perfectly, the probability of the observations given the model will
571 all be close to 1, as the scatter plot indicated for BA fractions above ~ 0.0003 . The model won't
572 always capture the uncertainty required to generate the observations. This generally happens
573 at specific burned areas (like low ones in this example). Areas where this happens often are
574 highlighted on the map with the map in the middle showing the performance at the 5th
575 percentile of the time series.

576



577
578
579

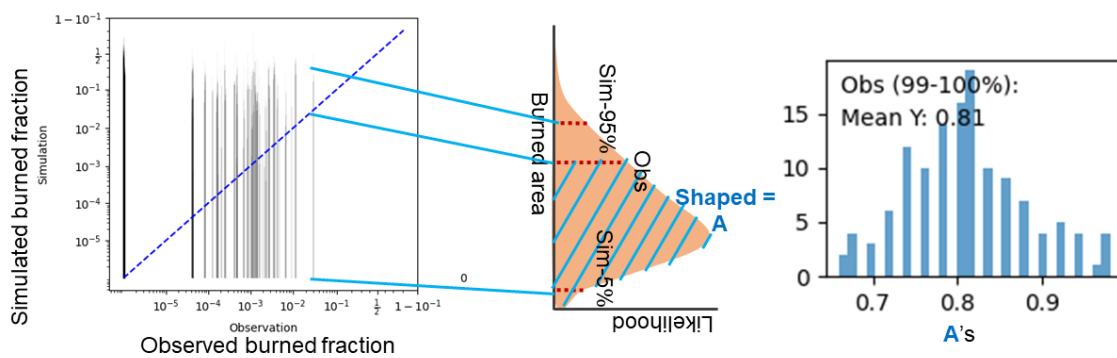
580 We also determine the percentile of our observations within the model's posterior probability
581 distribution. In an unbiased model, we expect the observation's position to be random. We can
582 start by doing this visually, as shown in the example from **Figure S30**: Observational BA (top
583 left) should generally fall between the two simulation maps (bottom) that span the 5-95
584 percentile of the model distribution. Taking the cell highlighted in blue for example - the lower
585 model estimate is close to zero and the upper is higher than the observations, indicating a
586 good performance at capturing the observations. Evaluating include parameters representing
587 noise or stochasticity in the system, that is not always included in the main analysis. Given the
588 inherent randomness in fire in our study regions, this does result in very broad BA distributions
589 in the model so a larger difference between the maps showing the BA in the model's tails
590 ("simulation - 5%" and "simulation - 95%") is to be expected.



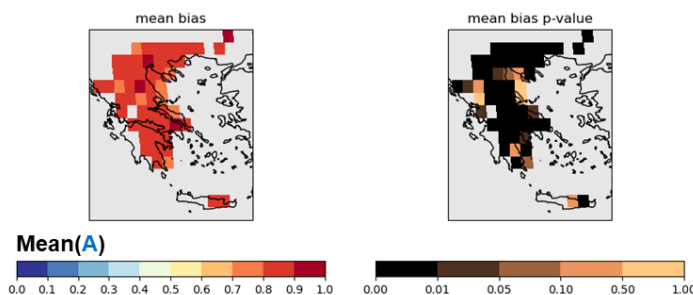
591
592
593
594

We compare the observations (x-axis) likely range (5-95 percentile) of the model's probability distribution. Similarly to the maps above, if the model captures the uncertainties, the

595 observations should fall within this range - i.e the 1:1 line should fall inside the span of the
 596 model, as seen in this example from **Figure S36**. We also calculate the mean position of the
 597 observations. This is simply the probability of BA greater than the observation, calculated by
 598 integrating equation 7 for BAs in the range $[0, BA]$. For simulations used in attribution, we also
 599 build histograms (right, taken from **Figure S39**) of this bias across different percentiles of the
 600 observations. This shows us if there is any part of the distribution that has a substantially
 601 different bias. In an unbiased model, these observational positions in the framework's
 602 probability distribution should average ("Mean Y:" in the histogram) to 0.5. Numbers close to
 603 1 indicate the observations on average tend towards the higher BA in the distribution, and the
 604 model generally underestimates BA. This alone does not show if the model performs poorly,
 605 and a consistent bias across all parts of the BA distribution indicates correct ranking, though
 606 the need for scaling for attribution analysis (see **Supplement Section "Modelling
 607 Frameworks > Confire"**).
 608



609 We also map out the mean position of the observations of the times series. Again, in an
 610 unbiased model, given enough timestep, this should average out to 0.5 for each gridcell.
 611 However, given the small number of timesteps, we map is the observational position in the
 612 frameworks posterior tends to be significantly different to 0.5 using a t-test to calculate a p-
 613 value for if the mean of the posterior position of the monthly observations for a given grid cell
 614 is significantly different from 0.5. Low p-values indicate where the model is biased, which tends
 615 to suggest too low or high burning.
 616
 617



618
 619
 620
 621
 622 **S2.2.1 Drivers of Regional Burned Area Extremes**
 623
 624 The model has shown a consistent ability to capture observations within its uncertainty range
 625 across all regions, indicating a robust representation of uncertainty. It also demonstrates a

626 high likelihood of aligning with actual observations, indicating strong alignment between
627 model outputs and real-world data. It effectively represents BA anomalies based on the
628 driving variables, demonstrating strong explanatory power across different regions.

629

630 However, the model consistently exhibits a low bias in estimating BA across regions, often
631 underestimating the BA, particularly in specific high-burn regions such as deforestation
632 areas in Western Amazonia and patches of high BA in northern Canada. This highlights the
633 common need across regions for better integration of data on human influences and
634 interactions with fire. It may also hint at the need for better representation of non-linearity
635 between drivers and BA.

636

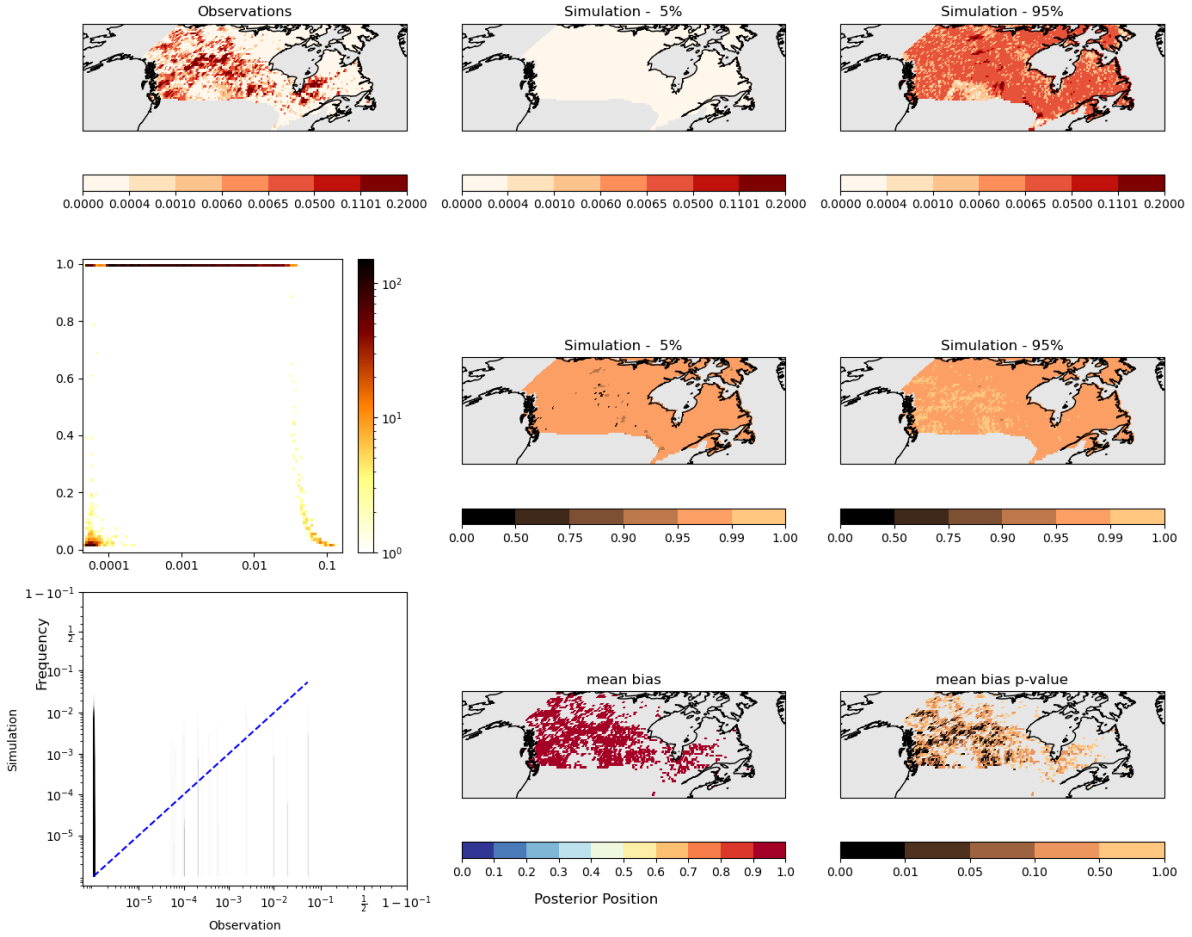
637

638 **Canada**

639 Our evaluation indicates that the model's assessment of uncertainty does a reasonable job
640 of capturing the observational range, particularly for high BA. The top row of **Figure S28**
641 demonstrates this - the observed (on the left) falls between the 5-95% range of the model.
642 However, there are patches of high BA that are slightly farther north than in the
643 observations. The model accurately identifies low burning in agricultural regions in southern
644 Saskatchewan and Alberta, and it suggests lower burning in the north while still
645 acknowledging the possibility of some burning.

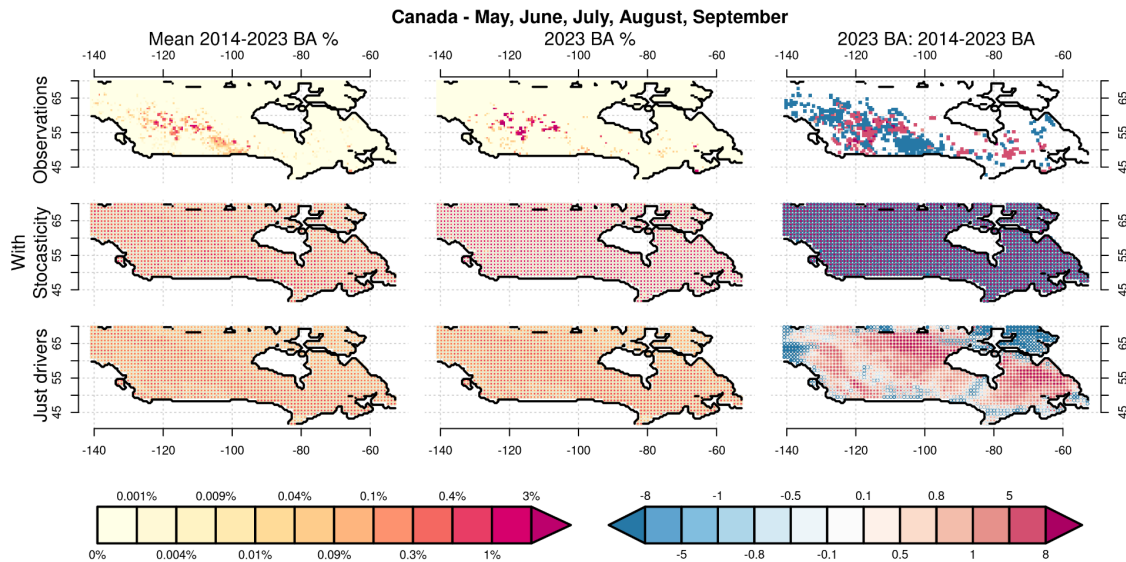
646

647 The probability of observations given the model is very high, especially for medium to high
648 BA ranging from 0.03% to 3% (**Figure S28** left middle row). This demonstrates that the
649 modelling framework does a reasonable job of generating observations within this
650 range. Except for a few locations, even the worst performances tend to show a likelihood of
651 observations given a model of > 0.95 . For very high BA, the probability decreases but
652 generally falls within the model's uncertainty range (bottom left), and ranked indicating that
653 the model can effectively identify high burning anomalies. However, it tends to
654 underestimate the increase in BA during such anomalies (i.e in **Figure 10**). In fact, the model
655 tends to be biased towards slightly lower BA in many regions of Canada (bottom middle
656 map), though only significantly so in the South and West of the country (bottom left map).
657 Interestingly, the BA picked up by the driving variables alone also effectively reproduces
658 spatial patterns of BA and regions of high anomalies in 2023, suggesting that the driving
659 variables used are good at explaining the observed patterns in BA.



660

661 **Figure S28:** Evaluation plot for driver attribution configuration in **Section 3** over Canada. (top
 662 observed and simulated BA fraction (%). (Middle row) the likelihood of the out-of-sample
 663 observations given the models probability distribution and (bottom row) observations position
 664 in the model distribution. See top of this section for interpretation guide.
 665



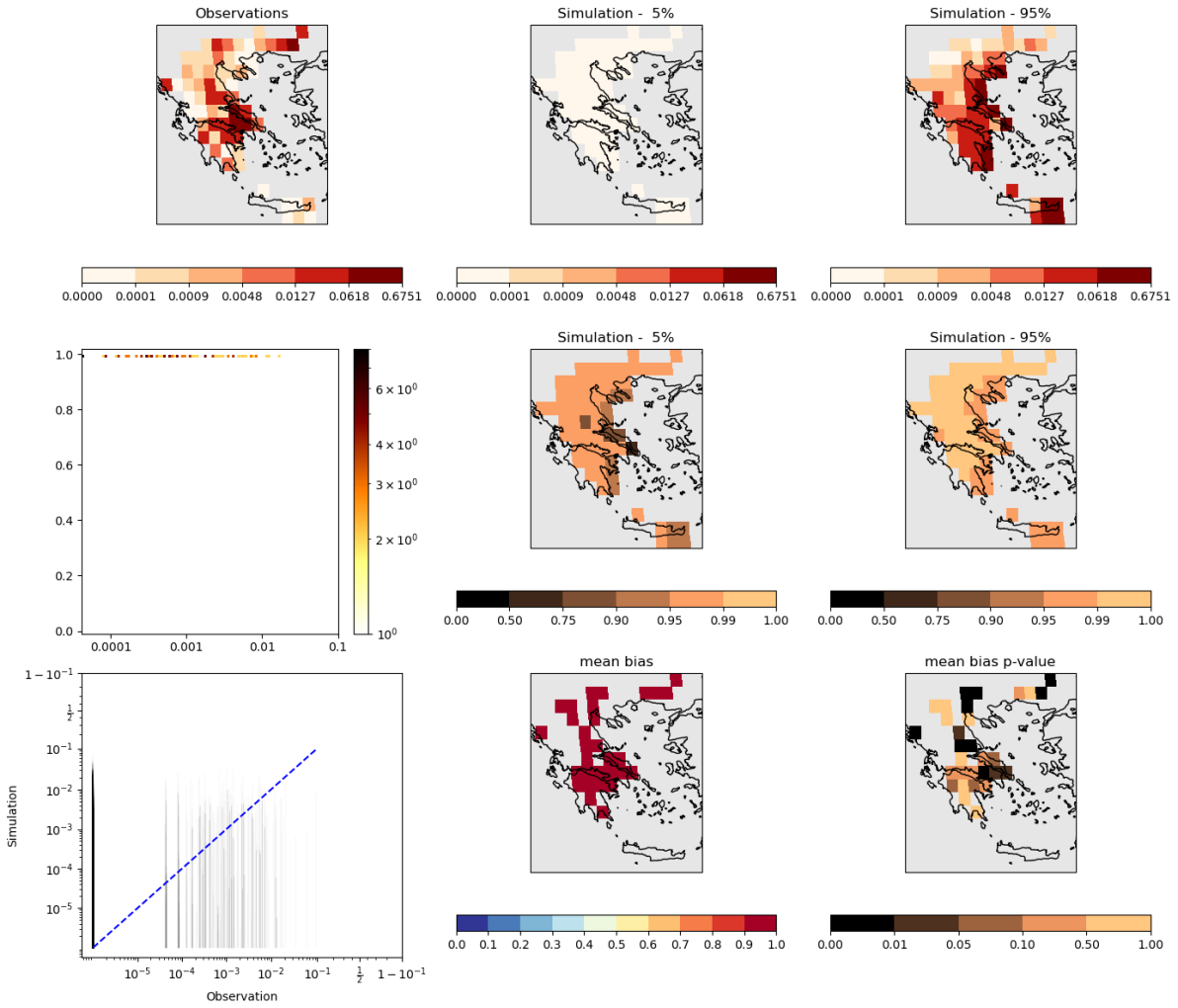
666

667 **Figure S29:** BA % over Canada for May-September for driver attribution configuration in
 668 **section 3** over (left) 2014-2023 (middle) 2023 and (right) for 2023 anomaly compared to 2014-
 669 2023, expressed as a factor of increase (red) or fractional decrease (blue). The top row is
 670 observations, the middle row in ConFire includes stochasticity (equation 7) and the bottom,
 671 just considers the influence of drivers (equation 8). For ConFire, the size of the dot in each
 672 grid cell shows the likelihood (larger = higher likelihood) of a BA fraction (or BA change) being
 673 greater than a given threshold (where the threshold is represented as a coloured dot, see
 674 legend at the base). High BA overlap smaller. i.e on the left, a large pale orange dot indicates
 675 a high likelihood of annual average BA exceeding 0.1%, with a small dark red dot indicating a
 676 small but non-zero likelihood of exceeding 3%

677

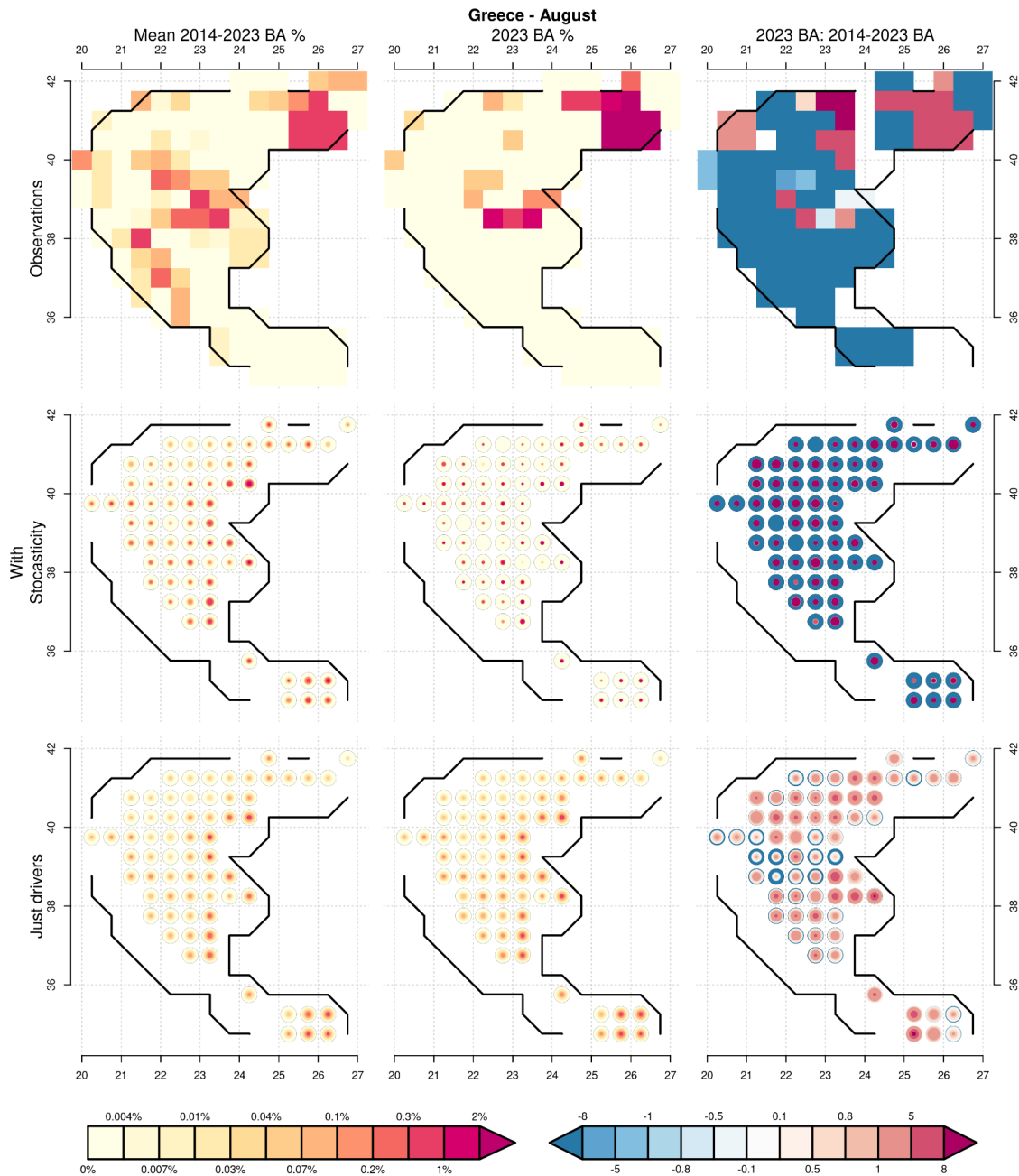
678 **Greece**

679 The model effectively represents uncertainties surrounding observed BA and accurately
 680 captures the gradient between low burning in the northwest of Greece and high burning
 681 around the southeastern coast. The model's observations show extremely high likelihood
 682 across all BA, with only a slight dip to around 0.75 likelihood in a few months in coastal
 683 Thessaly. Additionally, there is a consistent pattern of underestimating BA across all areas of
 684 Greece, although this is only significant in a few places.



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Figure S30: same as **Figure S28** for Greece



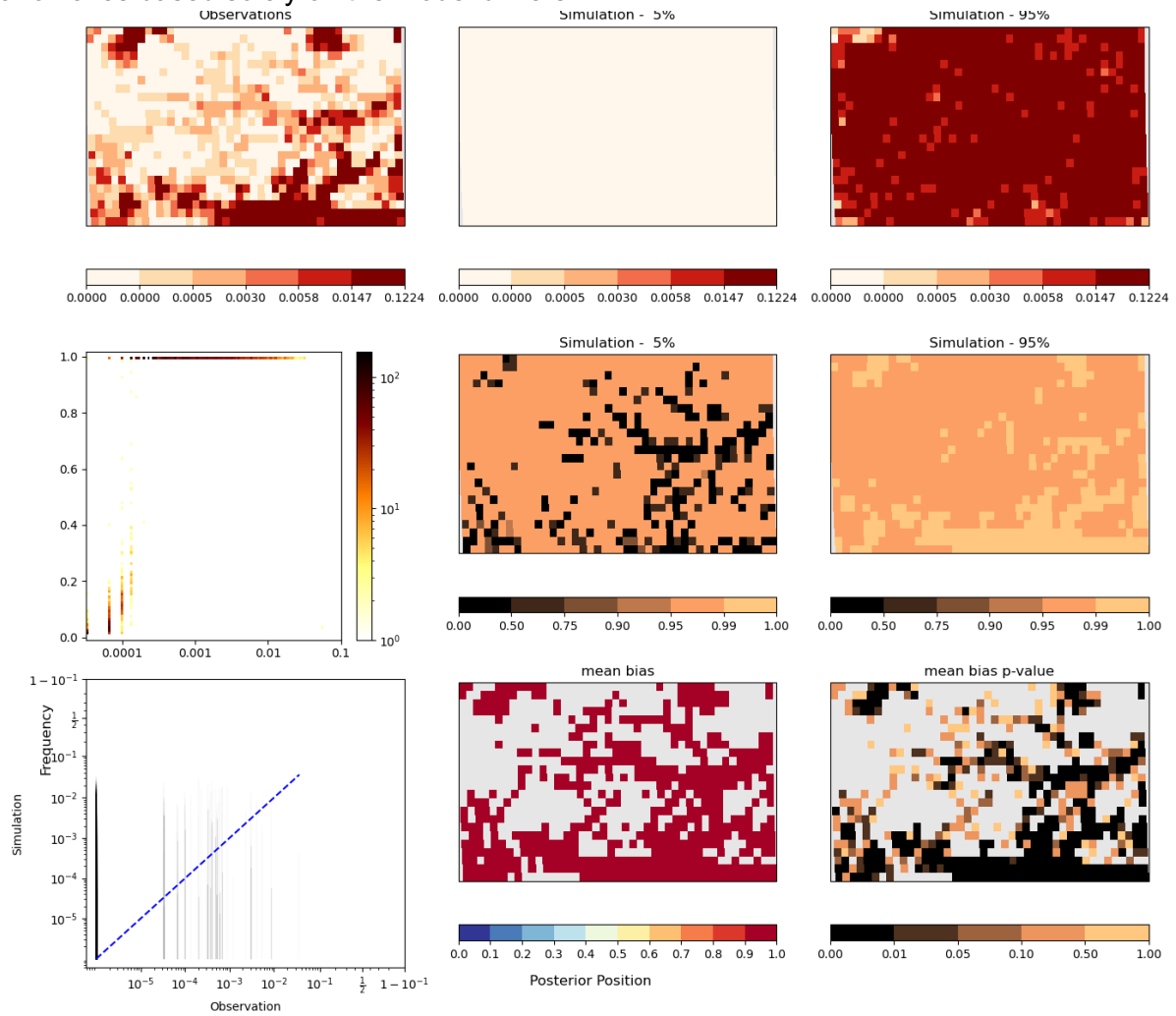
688
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Figure S31: same as **Figure S29** but for Greece in August

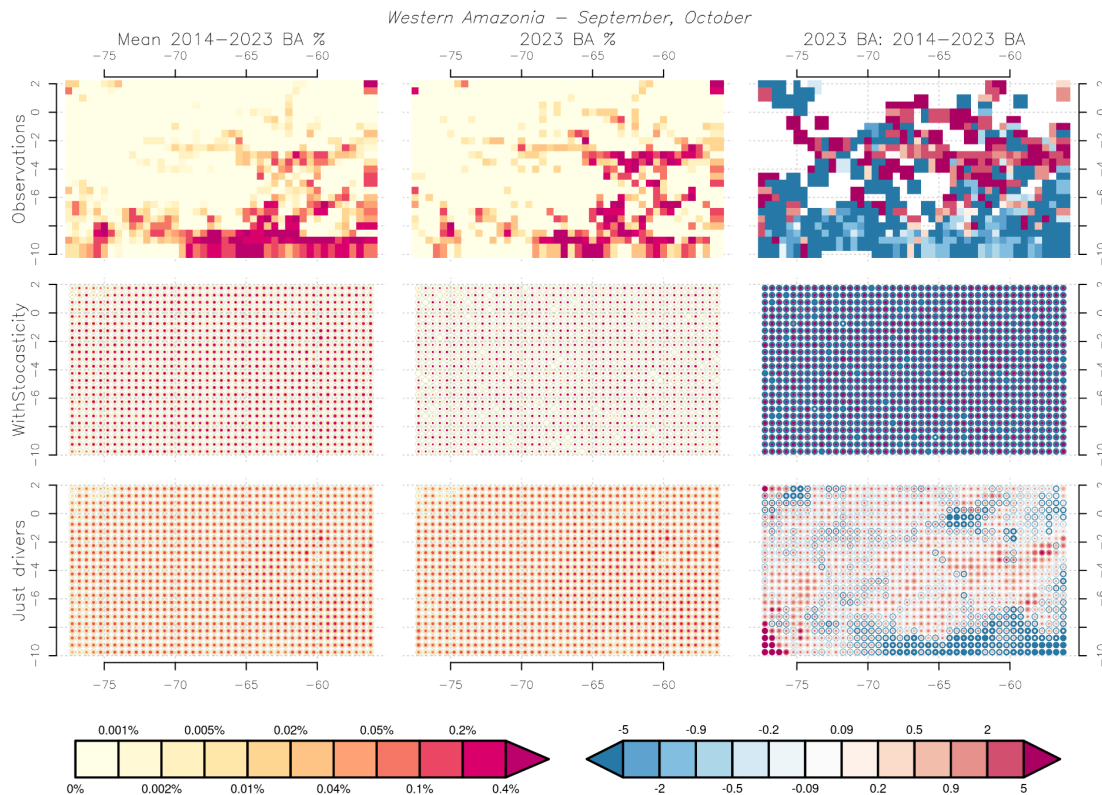
692 **Western Amazonia**

693 The model captures observations within its uncertainty range, but it fails to differentiate
 694 between high burning in deforestation regions in the south and north of the country. This
 695 suggests that vital data on deforestation and its interaction with fire may have been missed.
 696 The model is able to generate observations out of its sample, indicated by a high likelihood
 697 given observations. However, it does not generate very low BA, particularly in places where

698 high BA are also commonly observed in regions of deforestation. This suggests that the
 699 model may fail to capture variations in BA in these human-dominated areas. Similar to the
 700 other two regions, the model demonstrated a low bias. However it can accurately capture BA
 701 anomalies based solely on the model drivers.



702
 703 **Figure S32: same as Figure S28 for Western Amazonia**
 704
 705



706
707 **Figure S33:** same as **Figure S29** but for Western Amazonia in September and October.

708
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710 **S2.2.2 Change in Likelihood of High Burned Area in 2023 due to Total Climate Forcing**
711 **and Socioeconomic factors**

712
713 The framework utilising ISIMIP3a reanalysis data has been found to outperform its near-real-
714 time counterpart in simulating BA. It effectively represents high BA and extremes across all
715 regions. Furthermore, the probability of observations given the model is generally higher in
716 areas with extreme fires or high BA, indicating the model's reliability for attribution analysis.

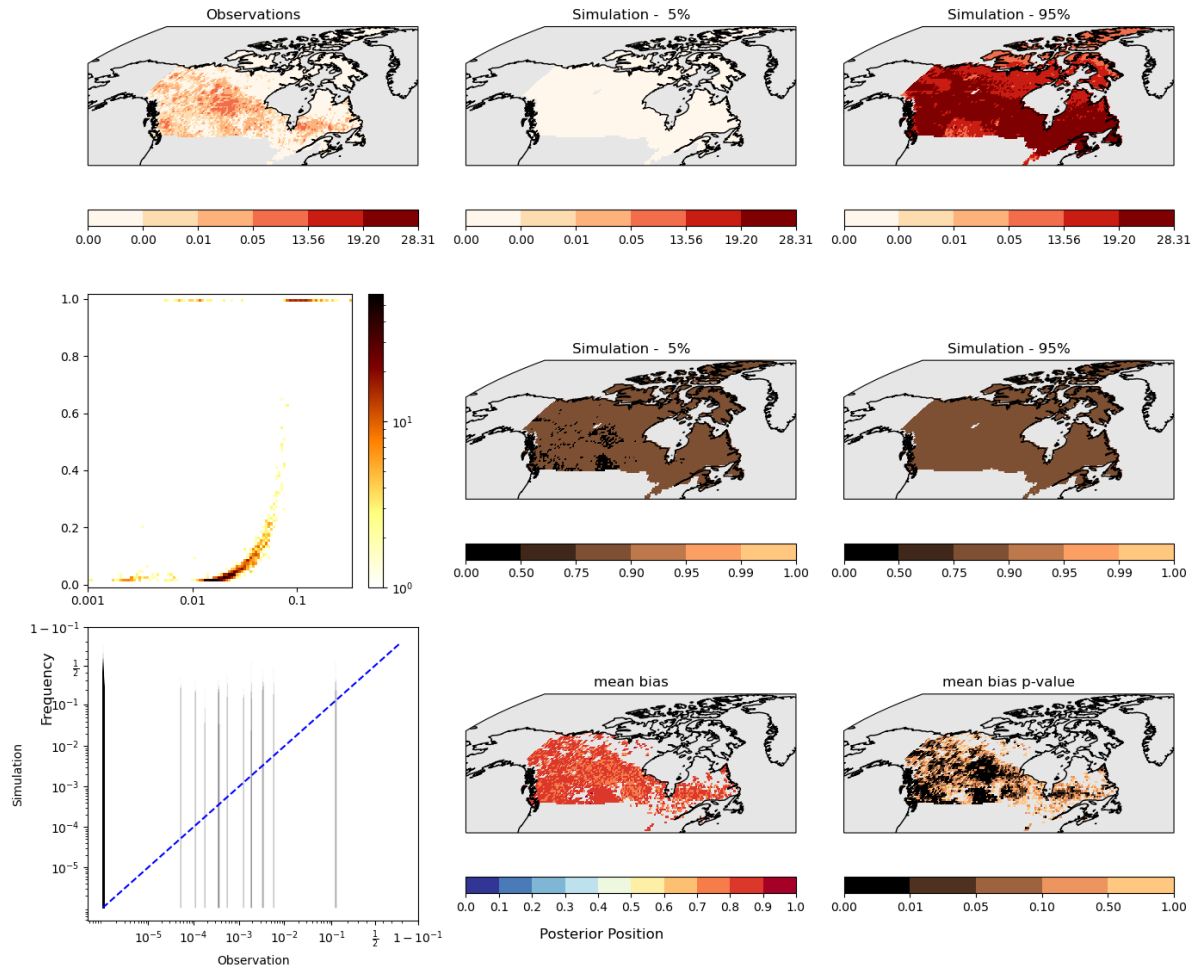
717
718 However, in regions of significant land use change, such as Western Amazonia, the model
719 struggles with reproducing higher BA, indicating a common challenge across regions in
720 capturing detailed land use interactions.

721
722 While observations consistently fall within range of the model distribution, the model
723 demonstrates consistent low bias. This simple scaling is suggested to align the model with
724 observations, highlighting a need for calibration to improve accuracy across regions.

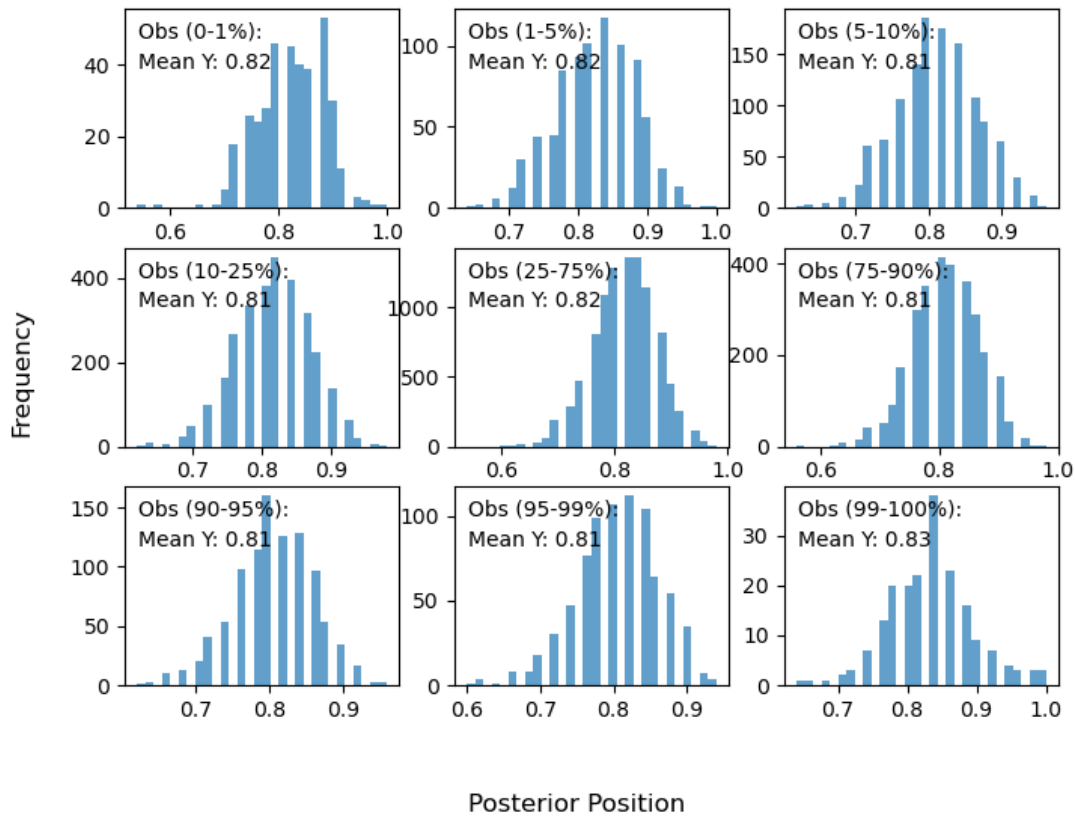
725
726
727 **Canada**

728 The analysis using isimip reanalysis data shows that the framework performs much better
729 than its near-real-time counterpart in assessing the drivers of BA (**Figure S34**). Although
730 there are still large uncertainty ranges, the highest BA in the distribution closely match the

731 high BA observed. In Canada, the model generally performs slightly worse in generating
 732 observations, but it still tends towards a probability of observations given the model of
 733 greater than 0.75. However, the model shows that high BA are very likely, indicating that the
 734 model is useful in representing extremes - critical for attribution analysis. Overall, the model
 735 exhibits less bias than its near real-time version, with observations falling on average around
 736 0.6-0.9 throughout the model's distribution. This consistent pattern across the observed
 737 distribution suggests that a simple scaling is required for attribution application (**Figure S35**).
 738



739
 740 **Figure S34:** same as **Figure S28** for attribution and future projections configuration used in
 741 **section 4** and **section 5**.

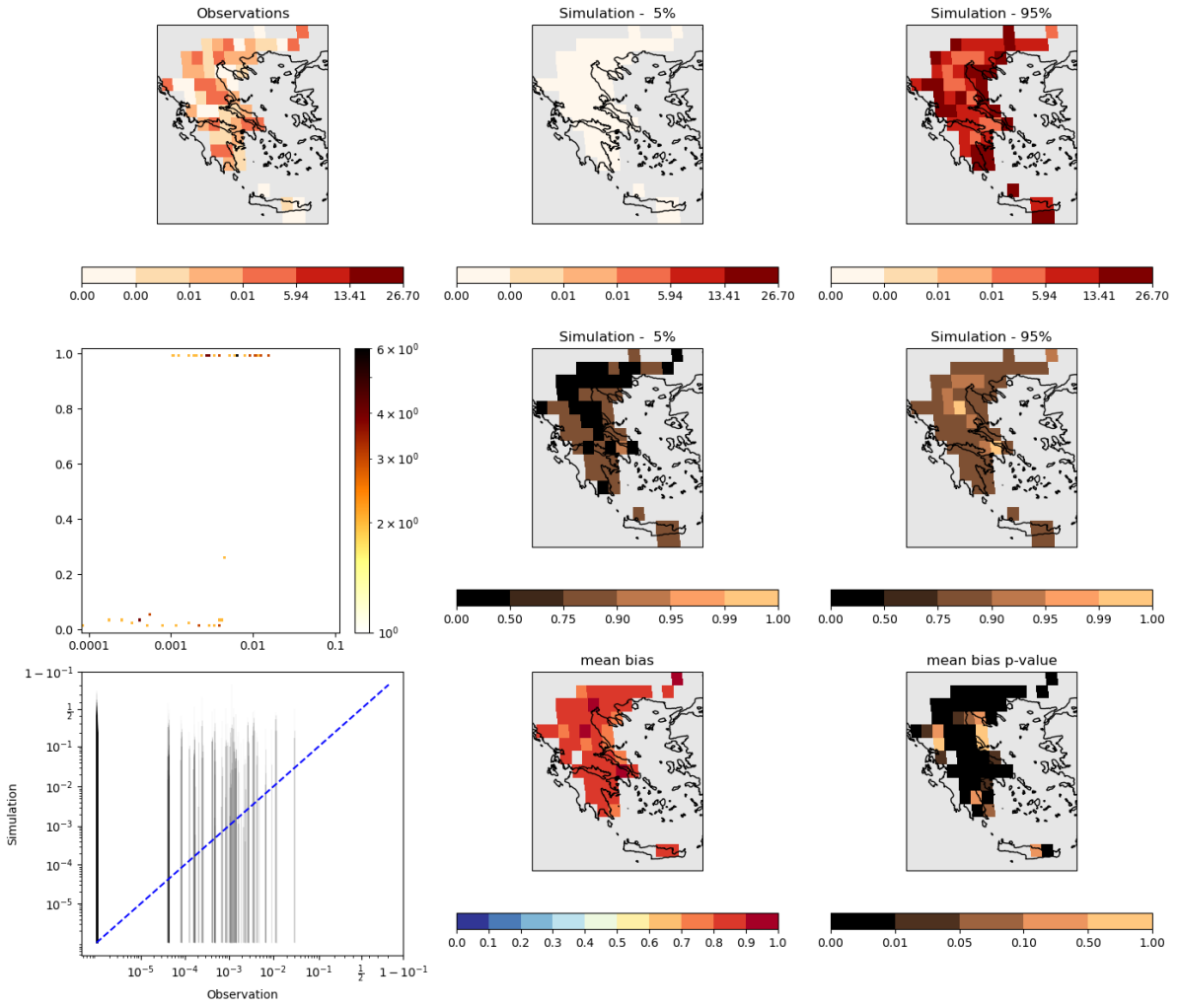


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Figure S35: The position of the observed BA in the model's probability distribution over the evaluation period using attribution and future projections configuration from **Section 4** and **Section 5** over Canada. Histograms are for observed percentiles indicated in the top left corner. See start of section for interpretation guide.

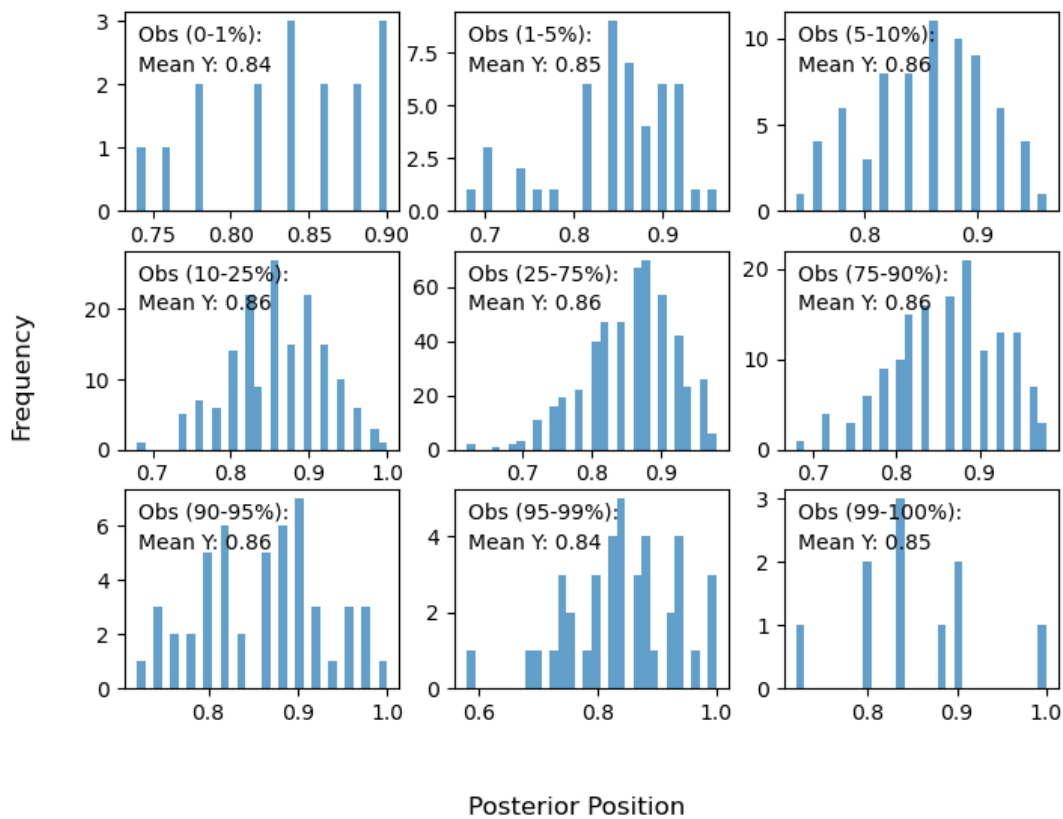
748 **Greece**

749 Over the longer evaluation periods, observations tend to be much noisier across Greece
750 than in the near real-time driver analysis (**Figure S36**). However, there is still a noticeable
751 trend towards more burning in the Southeast. This trend is well captured by the model,
752 including the more noisy spatial distribution in the observations. The probability of observing
753 a given model can be quite low, but it tends to be higher in areas where extreme fires were
754 observed and in areas with high burn areas, making it useful for attribution applications.
755 Additionally, while the model is biased low, similar to Canada, this bias is consistent across
756 the observed BA distribution (**Figure S37**).
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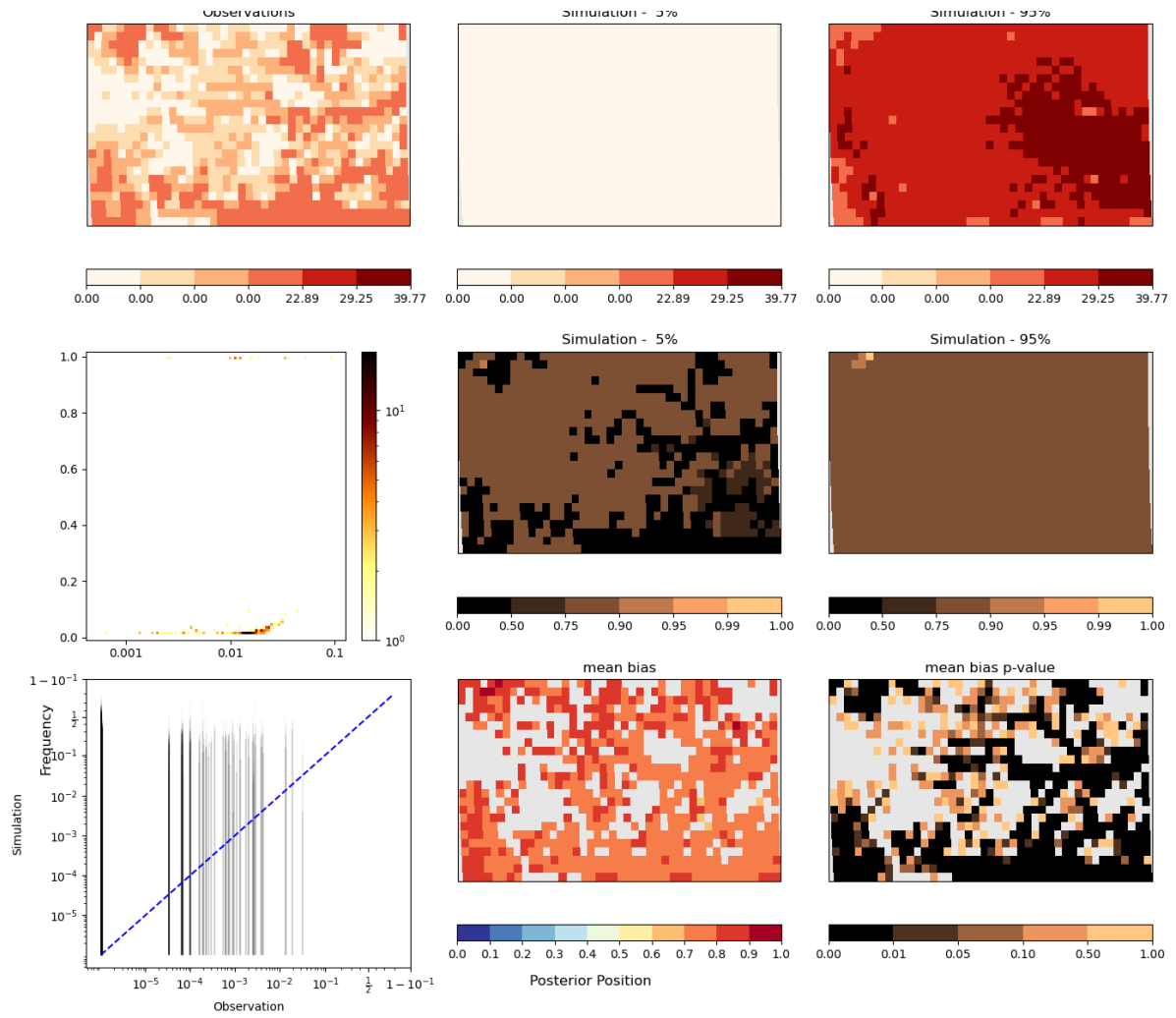
Figure S36: same as **Figure S34** for Greece



760
 761 **Figure S37:** same as **Figure S35** for Greece
 762

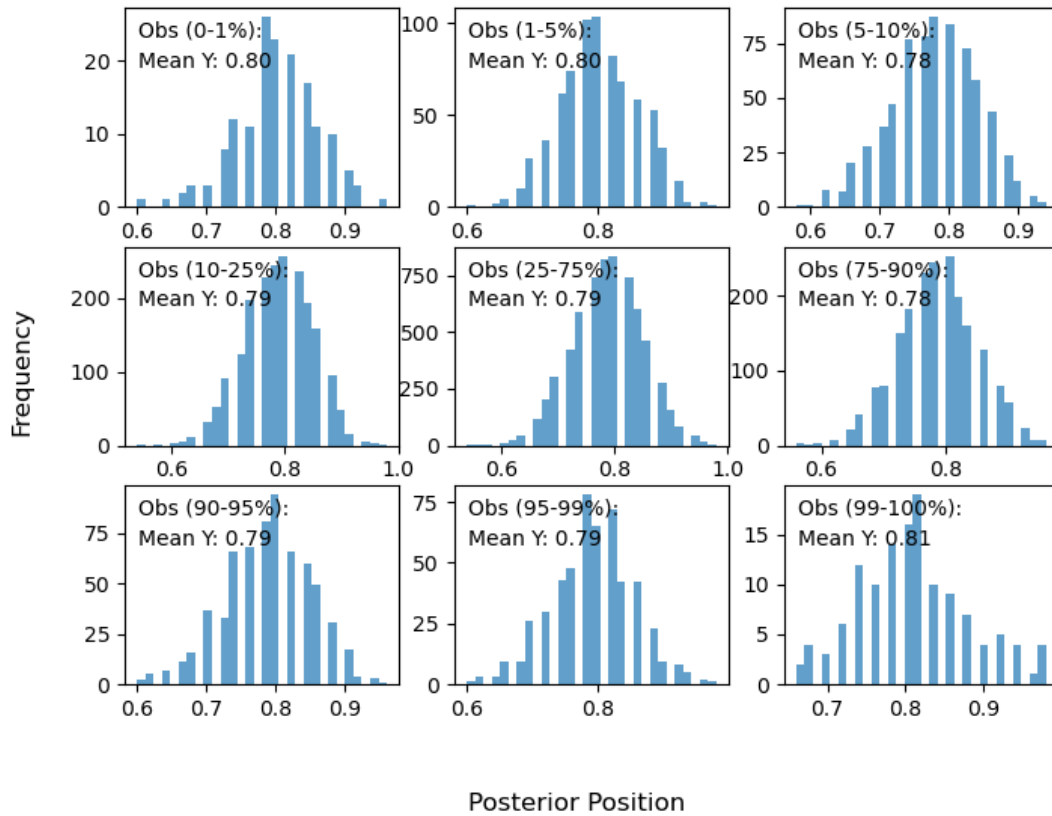
763 **Western Amazonia**

764 The framework outperforms its near real time counterpart in simulating higher BA around
 765 Manaus, although it still struggles to reproduce higher BA in regions of land use change
 766 (**Figure S38**). Observations fall within the model range and, like the other two regions, the
 767 observations indicate that the model tends to perform better at generating observed BA at
 768 higher levels of burning. Overall, this is the least biased region out of the three, although the
 769 model still tends to underestimate BA, with the observations falling at around 0.7-0.8 of the
 770 model distribution (**Figure S39**). This pattern is consistent across the distributions .



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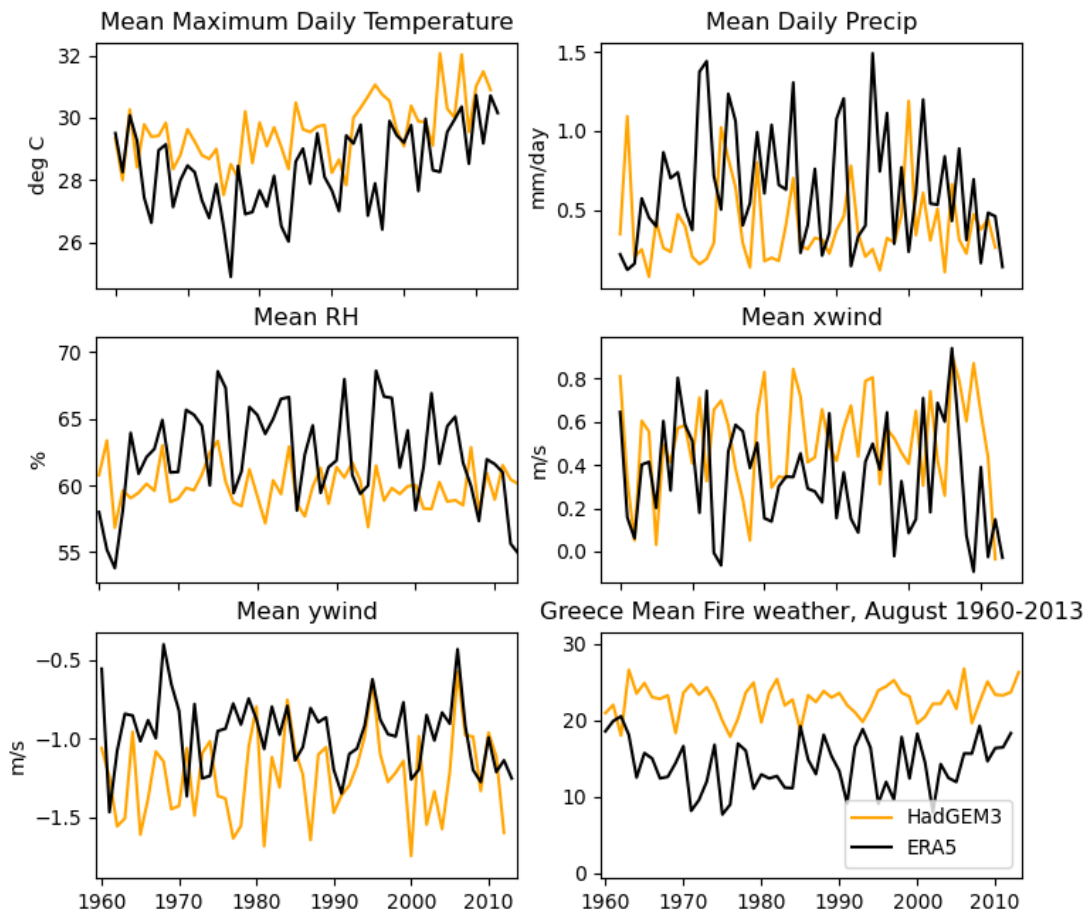
Figure S38: same as **Figure S34** for Western Amazonia.



773
 774 **Figure S39:** same as **Figure S35** for Western Amazonia
 775

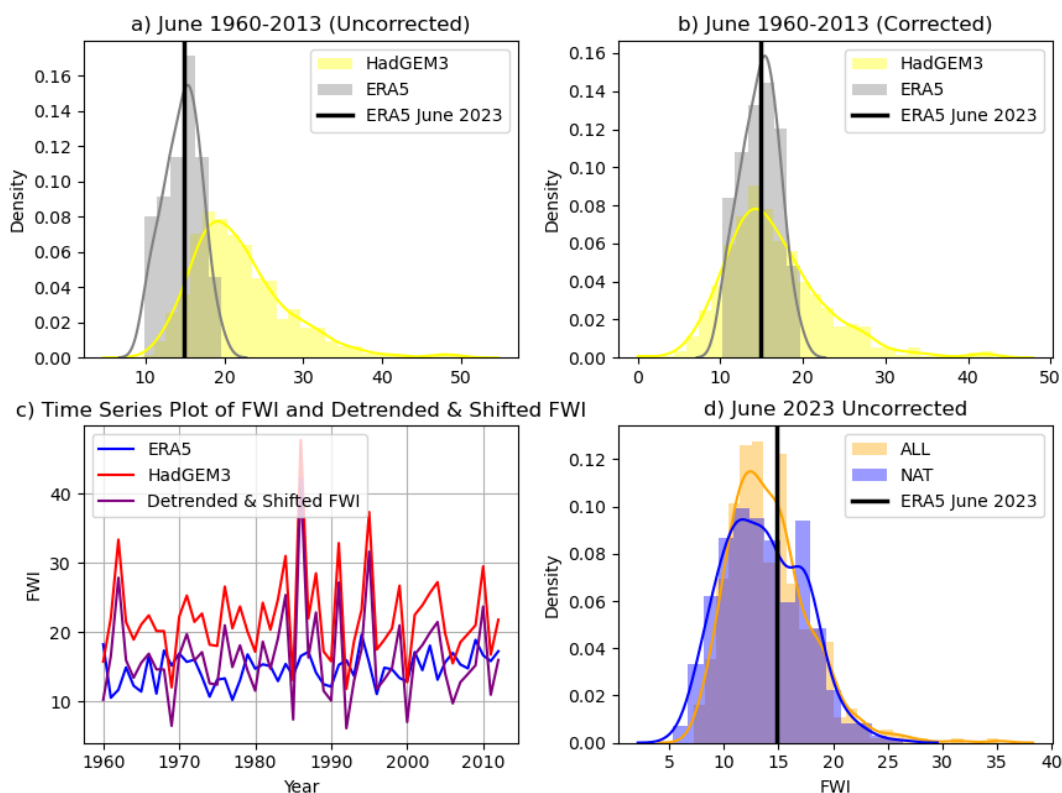
776
 777 **S2.3 Fire Weather attribution**

779 We evaluated each of the component variables used in the FWI against ERA5 reanalysis for
 780 the historical period 1960-2013. In each case, HadGEM3 was slightly biased across the
 781 timeseries, generally simulating conditions that were too hot and dry in Greece (**Figure S40**).
 782 This led to an overall larger bias in the resultant FWI (**Figure S43**). We therefore applied a
 783 linear bias-correction to the HadGEM3 ensemble of FWI (see Data and Data Processing).
 784 Results before and after the bias-correction is applied are shown below for each region.
 785



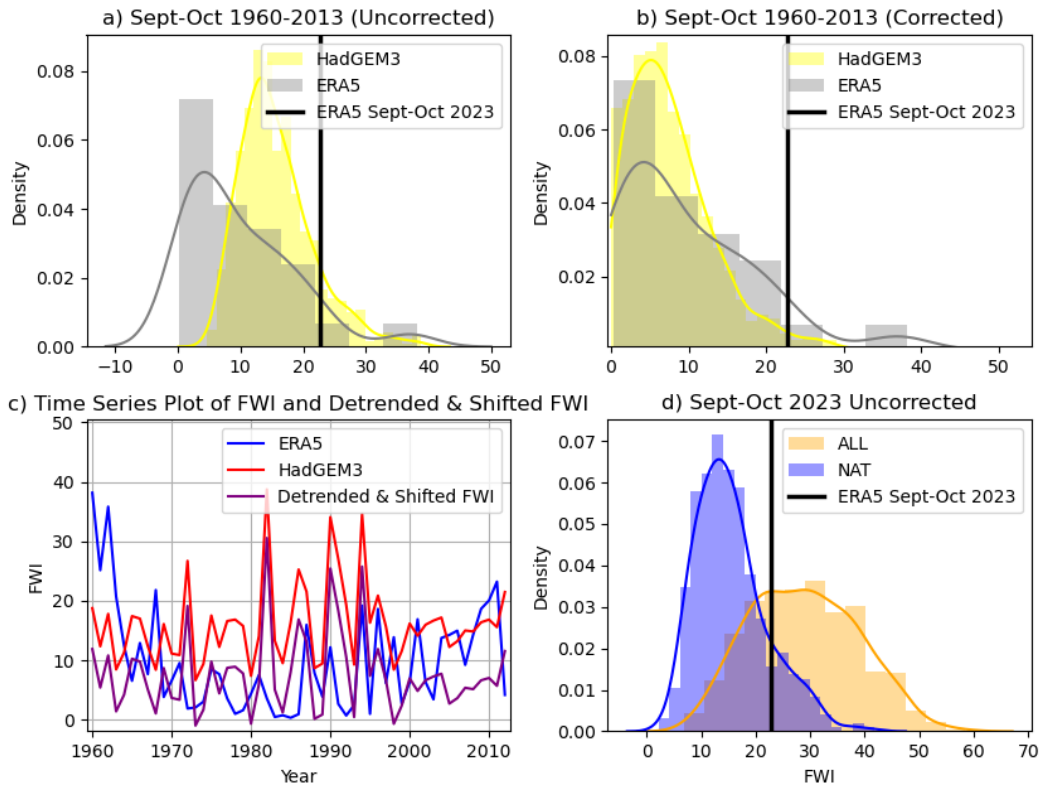
786
 787 **Figure S40:** Individual component variables of the FWI compared to ERA5 reanalysis across
 788 the historical period (1960-2013), and resultant FWI. Here one member from the HadGEM3
 789 historical ensemble is shown (yellow) against ERA5 (black) for one region (Greece), for
 790 illustration
 791
 792

Canada 95th percentile FWI



793
 794 **Figure S41:** Bias correction for Canada. Historical ensemble of HadGEM3 (yellow) compared
 795 to ERA5 (grey) 95th percentile of FWI for the historical period (1960-2013), shown as
 796 probability density before correction (a) and after correction (b), and one member shown as a
 797 timeseries (c, where HadGEM3 is shown in red, ERA5 in blue and corrected HadGEM3 in
 798 purple). HadGEM3 ensemble for 2023 shown before bias-correction (d). ERA5 2023 event
 799 shown as black vertical line on all probability density plots.
 800

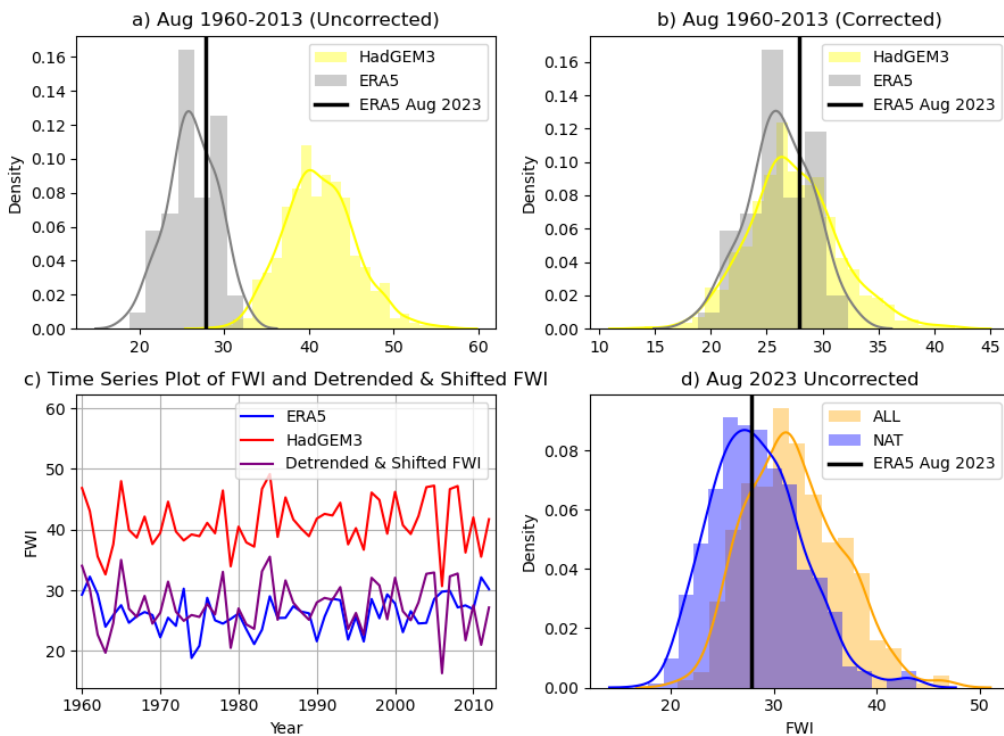
SAM 95th percentile FWI



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Figure S42: As for **Figure S41**, but for Western Amazonia

Greece 90th percentile FWI



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Figure S43: As for **Figure S41**, but for Greece at 90th percentile FWI

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